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Stability of Solutions of Stochastic Differential Equations of Diffusion Type

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Piotr Szlenk

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Major professor

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STABILITY OF SOLUTIONS OF STOCHASTIC DIFFERENTIAL EQUATIONS OF DIFFUSION TYPE

By

Piotr Szlenk

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ABSTRACT

STABILITY OF SOLUTIONS OF STOCHASTIC DIFFERENTIAL EQUATIONS OF DIFFUSION TYPE

By

Piotr Szlenk

The stability and asymptotic stability of solutions of one dimensional stochastic differential equations of diffusion type are studied. Such problems arise frequently in applications to oscillation theory, automatic control and related fields, as it is explained in [9].

In 1971 Khasminskii and Nevelson ([8]), considered the following stochastic differential equation:

$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t)$$

They assumed that σ never vanishes and that every solution of the equation above is recurrent. Under these assumptions they proved that the distance in some suitable metric (the scale metric) between two solutions starting from two different points converges to some random variable which is either zero or is concentrated on two points. This dissertation consists mainly of two parts:

- 1. In the first part (Section 4) the same stochastic differential equation is studied when σ is allowed to vanish. Different cases for which the convergence of the difference of two solutions takes place are analyzed.
- 2. In the second part (Section 5) the case when the drift coefficient b is constant is studied. In this case the limit of the difference of two solutions starting from two different points always exists. First, the case when all solutions are recurrent is discussed. Then the case when solutions are transient is studied. An investigation on the conditions on σ for which the limit of the difference of the solutions is zero is carried out. It is proved that if we assume that $\sigma(x)$ is concave or convex for sufficiently large x, then the limit of the difference of two solutions is zero if and only if $\int_0^{+\infty} (\sigma'(x))^2 dx = +\infty$

The main steps of the proofs of these results hinge on a Comparison Theorem of Skorokhod for the solutions of Stochastic Differential Equations, the convergence theorem for nonnegative supermartingales, representation of continuous local martingales as time-changed Brownian Motions, the exact growth of transient solutions of Stochastic Differential Equations, and the fact that a continuous local martingale is convergent on the set where its quadratic variation converges.

To my parents

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1 INTRODUCTION

In their paper of 1971 ([8]) Khasminskii and Nevelson (for more detailed discussion see also [9], pages 302-309) studied the following problem:

Consider the following stochastic differential equation:

(1)
$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t).$$

Assume the following:

- I) Equation (1) has a unique solution almost surely.
- II) σ never vanishes.
- III) Every solution of (1) is recurrent.

Define the scale function:

$$Q(x) = \int_0^x e^{-2\int_0^y \frac{b(u)}{\sigma^2(u)}du} dy.$$

Then condition III is equivalent to the following:

$$Q(+\infty) = +\infty, Q(-\infty) = -\infty$$

If we put Y(t)=Q(X(t)), then Ito's formula ([4], Theorem II-5.1) implies that:

(2)
$$dY(t) = \sigma_1(Y(t))dW(t)$$
 where $\sigma_1(y) = \sigma(Q^{-1}(y))Q'(Q^{-1}(y))$.

If we take two solutions of (2) starting from two different points (say y_1 , y_2 , $y_2 > y_1$) then

$$Y_2(t) - Y_1(t) = y_2 - y_1 + \int_0^t [\sigma_1(Y_2(s)) - \sigma_1(Y_1(s))] dW(s).$$

Since the solution of (2) is unique, $Y_2(t) - Y_1(t) \ge 0$ for all t a.s. Since $P(\int_0^t [\sigma_1(Y_2(s)) - \sigma_1(Y_1(s))]^2 ds < +\infty) = 1$ for all t, $\int_0^t [\sigma_1(Y_2(s)) - \sigma_1(Y_1(s))] dW(s)$ is a continuous local martingale (see [4]:Chapter II, Section 1). Since every positive local martingale is a supermartingale, $Y_2(t) - Y_1(t)$ is a positive supermartingale, and therefore $Y_2(t) - Y_1(t)$ converges to a positive random variable, say ξ , a.s. as $t \to +\infty$. (see [10], Section 39).

The work in [8] deals with the analysis of the random variable ξ . It can be presented by a Lemma 1.1 and Theorems 1.1 and 1.3 below(the proof of Lemma 1.1 which is a crucial step is contained in the proof of Theorem 1 in [8] and will not be reproduced here). Theorem 1.2 is a modification of Theorem 1.1, where in part (c) the form of the distribution of the limiting random variable ξ is new.

Lemma 1.1 $\sigma_1(y+\xi) = \sigma_1(y)$ for every y a.s.

Theorem 1.1 Under assumptions I, II, III we have: a) if σ_1 is not periodic, then $\xi = 0$.

- b) if σ_1 is periodic with period θ , and $\frac{y_2-y_1}{\theta}=k$ is an integer, then $Y_2(t)-Y_1(t)=y_2-y_1$ a.s. for all t. In particular this is the case if σ_1 is constant.
- c) if σ_1 is periodic with period θ and $\frac{\nu_1-\nu_1}{\theta}=k$ is not an integer, then the distribution of ξ is concentrated on two points $\theta[k]$ and $\theta([k]+1)$, where [k] denotes the integer part of k.

Let us compute the distribution of the random variable ξ in case σ_1 is periodic with period θ . Let $k = \frac{m-m}{\theta}$, $\{k\} = k-[k]$. Repeating the argument of Khasminskii and Nevelson we conclude that

$$\theta[k] \le Y_2(t) - Y_1(t) \le \theta([k] + 1)$$

But

$$Y_2(t) - Y_1(t) = y_2 - y_1 + \int_0^t [\sigma_1(Y_2(s)) - \sigma_1(Y_1(s))] dW(s)$$

and from Lebesgue's dominated convergence theorem we conclude that

$$E\xi = E \lim_{t \to +\infty} (Y_2(t) - Y_1(t)) = y_2 - y_1$$

Therefore we have:

$$\theta[k]P(\xi = \theta[k]) + \theta([k] + 1)(1 - P(\xi = \theta[k])) = y_2 - y_1 = \theta k$$

From this we have

$$P(\xi = \theta[k]) = 1 - \{k\}$$
 , $P(\xi = \theta([k] + 1)) = \{k\}$

Therefore we can restate Theorem 1.1 as follows:

Theorem 1.2 Under assumptions I, II, III we have :

a) if σ_1 is not periodic, then $\xi = 0$.

- b) if σ_1 is constant, then $Y_2(t) Y_1(t) = y_2 y_1$ a.s. for all t.
- c) if σ_1 is periodic with period θ , then

$$P(\xi = \theta[k]) = 1 - \{k\} \text{ and } P(\xi = \theta([k] + 1)) = \{k\} \text{ where } k = \frac{y_2 - y_1}{\theta}.$$

Therefore, if we define a new metric in R by

$$r(x,y) = |Q(x) - Q(y)|$$

then we will have:

Theorem 1.3 Under assumptions I, II, III we have :

For every two solutions of (1) $X_2(t)$ and $X_1(t)$, starting respectively from x_2 and x_1 ($x_2 > x_1$) we have :

$$r(X_2(t), X_1(t)) \rightarrow \xi \ a.s.$$

where $\xi = 0$ if $\sigma_1(y) = \sigma(Q^{-1}(y))Q'(Q^{-1}(y))$ is not periodic,

 $\xi = r(x_2, x_1)$ if σ_1 is constant, and

$$P(\xi = \theta[\frac{r(x_2, x_1)}{\theta}]) = 1 - \{\frac{r(x_2, x_1)}{\theta}\} \quad , \quad P(\xi = \theta([\frac{r(x_2, x_1)}{\theta}] + 1)) = \{\frac{r(x_2, x_1)}{\theta}\}$$

if σ_1 is periodic with period θ .

Our principal contributions in this dissertation are presented in Sections 4 and 5. In Section 4 we will discuss the convergence of the difference of two

solutions of (1), where we allow σ to vanish. In Section 5 we will discuss the same property of solutions of (1), where we assume that the drift coefficient is constant. The novelty of some results in these sections is in dealing directly with the difference of solutions, rather than considering the distance between solutions with respect to a scale metric. Sections 2 and 3 contain preparatory materials for Section 4 and 5.

2 GENERAL THEORY

In this Section we present the known theory which is needed in sections 3, 4, and 5. We refer the reader to [4] for details. We start with the definition of a martingale. Let (Ω, \mathcal{F}, P) be a complete probability space. Let \mathcal{F}_t be a filtration, that is \mathcal{F}_t is a family of σ -algebras for which $\mathcal{F}_s \subseteq \mathcal{F}_t$ if $s \leq t$. Let

$$\mathcal{F}_{t+} = \bigcap_{s>t} \mathcal{F}_s$$

We say that the filtration \mathcal{F}_t satisfies the usual conditions if \mathcal{F}_0 contains all P-null sets and if

$$\mathcal{F}_{t+} = \mathcal{F}_t$$
 for all t .

Definition 2.1 We say that a stochastic process X(t) is a \mathcal{F}_t -martingale (submartingale, supermartingale, respectively) if and only if X(t) is \mathcal{F}_t -measurable, $E(|X(t)|) < +\infty$ for every t, and $E(X(t)|\mathcal{F}_s) = X(s)$ (\geq,\leq , respectively). The martingale is a square-integrable martingale if $EX^2(t) < +\infty$ for every t.

We denote the collection of square-integrable \mathcal{F}_t -martingales by \mathcal{M}_2 . Similarly, we denote the collection of continuous, square-integrable martingales by \mathcal{M}_2^c

For supermartingales we have the following theorem, which is used extensively in the thesis:

Theorem 2.1 ([10], Section 39) Let X(t) be a supermartingale. Let $X^-(t)$ denote the negative part of X(t), that is $X^-(t) = -X(t)\chi_{\{X(t) \leq 0\}}$. Assume

that
$$\sup_{t} EX^{-}(t) < +\infty$$
.
Then $\lim_{t \to +\infty} X(t)$ exists a.s., and $E\lim_{t \to +\infty} X(t) < +\infty$.

In the next two subsections we introduce the stochastic integrals with respect to martingales and local martingales. These integrals are extensions of the classical Ito integrals with respect to Brownian Motion. The need for developing stochastic integrals with respect to Brownian Motion, or more generally with respect to local martingales, arises from the fact, that the sample paths of such processes are of unbounded variation with probability one. Therefore, the usual Lebesgue-Stieltjes integral for the sample paths of such processes cannot be defined.

2.1 STOCHASTIC INTEGRALS WITH RESPECT TO MARTINGALES

In this subsection we define the stochastic integral with respect to a square-integrable martingale M(t). The next Proposition will help us to determine the class of integrands, that is the class of processes for which our integral will be defined.

Proposition 2.1 ([4], Proposition II-2.1) Let $M(t) \in \mathcal{M}_2^c$. Then there is a unique, increasing, integrable process denoted by $\langle M \rangle(t)$, such that $M^2(t) - \langle M \rangle(t)$ is a martingale.

Definition 2.2 The process $\langle M \rangle(t)$ from Proposition 2.1 is called the quadratic variation process of M(t).

Example 2.1 Let B(t) be a Brownian motion. Then $\langle B \rangle(t) = t$.

The next definition specifies the class of processes for which the stochastic integral with respect to the square-integrable martingale M(t) will be defined.

Definition 2.3 Let $M \in \mathcal{M}_2^c$. Let $\mathcal{L}_2(M)$ be the collection of predictable processes $\Phi(t)$ such that

$$(||\Phi||_{2,T}^{M})^{2} = E\{\int_{0}^{T}\Phi^{2}(s,\omega)d\langle M\rangle(s)\} < +\infty \text{ for each } T>0.$$

For $\Phi \in \mathcal{L}_2(M)$, we put

$$||\Phi||_{2}^{M} = \sum_{n=1}^{+\infty} \frac{1}{2^{n}} (||\Phi||_{2,n}^{M} \wedge 1)$$

where $x \wedge y = min(x, y)$.

Proposition 2.2 ([4], Section II-1) $\mathcal{L}_2(M)$ forms a complete metric space with the metric defined by $||\Phi - \Psi||_2^M$.

Let \mathcal{L}_0 be the collection of processes $\Phi(t)$ for which there exists an increasing sequence $\{t_i\}_{i=0}^{+\infty}$ $(0 = t_0 < t_1 < t_2 < ... < t_n <)$ and a sequence of random variables $\{f_i(\omega)\}_{i=0}^{+\infty}$, such that f_i is \mathcal{F}_{t_i} -measurable with $\sup_i ||f_i||_{\infty} < +\infty$, and

$$\Phi(t,\omega) = f_0(\omega)\chi_{\{t=0\}}(t) + \sum_{i=0}^{+\infty} f_i(\omega)\chi_{(t_i,t_{i+1}]}(t)$$

The following Lemma shows that every process from $\mathcal{L}_2(M)$ can be approximated by jump processes from \mathcal{L}_0 .

Lemma 2.1 ([4], Lemma II-2.1) The subspace \mathcal{L}_0 is dense in $\mathcal{L}_2(M)$.

Definition 2.4 Let $X \in \mathcal{M}_2$. Let

$$|X|_T = [EX^2(T)]^{\frac{1}{2}}.$$

Let

$$|X| = \sum_{n=1}^{+\infty} \frac{1}{2^n} (|X|_n \wedge 1).$$

Lemma 2.2 ([4], Lemma 1.2) \mathcal{M}_2 form a complete metric space with the metric defined by |X - Y|. Moreover \mathcal{M}_2^c is a closed subspace of \mathcal{M}_2 .

Now, we will define the stochastic integral with respect to a square integrable martingale M(t). We proceed as follows. For $\Phi \in \mathcal{L}_0$ with the expansion $\Phi(t,\omega) = f_0(\omega)\chi_{\{t=0\}}(t) + \sum_{i=0}^{+\infty} f_i(\omega)\chi_{\{t_i,t_{i+1}\}}(t)$, we define

$$I^{M}(\Phi)(t) = \sum_{i=0}^{+\infty} f_{i}(M(t_{i+1} \wedge t) - M(t_{i} \wedge t)).$$

It can be shown ([4], Section II-2) that $I^{M}(\Phi) \in \mathcal{M}_{2}^{c}$ and

(ISOMETRY)
$$|I^{M}(\Phi)| = ||\Phi||_{2}^{M}$$
.

Therefore we obtain an isometry from \mathcal{L}_0 to \mathcal{M}_2^c . Now using Lemma 2.1 we can extend this definition to the whole space $\mathcal{L}_2(M)$.

From now on, we will write $I^{M}(\Phi)(t) \equiv \int_{0}^{t} \Phi(s) dM(s)$.

Therefore we have defined the stochastic integral of a stochastic process belonging to the class \mathcal{L}_2 with respect to the martingale M. In the next subsection we extend this definition to the more general class of integrators, which are called local martingales.

2.2 STOCHASTIC INTEGRALS WITH RESPECT TO LOCAL MARTINGALES

As before, we follow [4] in the presentation of this section.

Definition 2.5 ([4], Definition II-1.7) Stochastic process M(t) is called a local martingale if there exists an increasing sequence of \mathcal{F}_t -stopping times τ_n , such that $\tau_n \to +\infty$ a.s. and $M(t \wedge \tau_n)$ is a \mathcal{F}_t -martingale for every n. The collection of continuous, square-integrable local martingales will be denoted by $\mathcal{M}_2^{c,loc}$, and \mathcal{M}_2^{loc} will denote the collection of square-integrable, local martingales.

We can define the quadratic variation process of a local martingale M in a same way as in the case of a martingale. Such a process is denoted by $\langle M \rangle(t)$ and $M^2(t) - \langle M \rangle(t)$ is a local martingale.

Definition 2.6 Let $M \in \mathcal{M}_2^{loc}$. Let $\mathcal{L}_2^{loc}(M)$ be the collection of predictable processes $\Phi(t)$ for which there exists a sequence of \mathcal{F}_t -stopping times $\sigma_n \to +\infty$ that

$$E\{\int_0^{T \wedge \sigma_n} \Phi^2(t,\omega) d\langle M \rangle(t)\} < +\infty$$

for every T and n=1,2,3,4,...

Let $M \in \mathcal{M}_2^{loc}$ and $\Phi \in \mathcal{L}_2^{loc}(M)$. Then there exists a sequence of \mathcal{F}_t stopping times $\sigma_n \to +\infty$ such that $M(t \wedge \sigma_n)$ is a square-integrable martingale and $\Phi(t)$ with σ_n satisfies the condition from Definition 2.6. Let $\Phi_n(t) = \chi_{\{\sigma_n \geq t\}} \Phi(t)$. Let $M_n(t) = M(\sigma_n \wedge t)$. Since M_n is square-integrable
martingale we can define the stochastic integral $I^{M_n}(\Phi_n)(t)$. It is known
that $I^{M_m}(\Phi_m)(t) = I^{M_n}(\Phi_n)(t \wedge \sigma_m)$ where $n \geq m$. Therefore there exists a
unique process $I^M(\Phi)(t)$ such that $I^M(\Phi)(t \wedge \sigma_n) = I^{M_n}(\Phi_n)(t)$. The process $I^M(\Phi)(t)$ is called the stochastic integral of Φ with respect to M and is
denoted by $\int_0^t \Phi(s)(\omega) dM(s)$. Such stochastic integral is a continuous local
martingale.

In the next subsection we discuss very interesting and important property of local martingales which will be needed in Section 5.

2.3 LOCAL MARTINGALES AS TIME-CHANGED BROWNIAN MOTIONS

Theorem 2.2 ([4], Theorem II-7.2) Let $M \in \mathcal{M}_2^{c,loc}$. Assume that

$$\lim_{t\to +\infty} \langle M \rangle(t) = +\infty.$$

Let

$$\tau_t = \inf\{u : \langle M \rangle(u) > t\}.$$

Then the stochastic process $B(t) = M(\tau_t)$ is a \mathcal{F}_{τ_t} -Brownian Motion.

The conclusion of the Theorem 2.2 is also true without the assumption that $\lim_{t\to +\infty} \langle M \rangle(t) = +\infty$, but in this case one has to extend the underlying probability space, so that the existence of the appropriate Brownian Motion can be guaranteed.

Definition 2.7 Let $(\Omega, \mathcal{F}, \mathcal{F}_t, P)$ be a probability space with filtration \mathcal{F}_t . Let $(\Omega', \mathcal{F}', \mathcal{F}'_t, P')$ be another probability space. Let $\tilde{\Omega} = \Omega \times \Omega'$, $\tilde{\mathcal{F}} = \mathcal{F} \times \mathcal{F}'$, $\tilde{P} = P \times P'$. Let $\pi(\tilde{\omega}) = \omega$ for $\tilde{\omega} = (\omega, \omega')$. If $\tilde{\mathcal{F}}_t$ is a filtration for $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{P})$ such that $\mathcal{F} \times \mathcal{F}' \supset \tilde{\mathcal{F}}_t \supset \mathcal{F}_t \times \{\Omega', \emptyset\}$, then we call $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{\mathcal{F}}_t, \tilde{P})$ a natural extension of the probability space $(\Omega, \mathcal{F}, \mathcal{F}_t, P)$.

Theorem 2.3 ([4], Theorem II-7.2') Let $M \in \mathcal{M}_2^{c,loc}$. Let

$$\tau_t = \begin{cases} inf\{u : \langle M \rangle(u) > t\} & if \ t < \langle M \rangle(+\infty) \\ +\infty & if \ t \ge \langle M \rangle(+\infty) \end{cases}$$

Let $\hat{\mathcal{F}}_t = \bigvee_{s>0} \mathcal{F}_{\tau_t \wedge s}$. Then there exists a natural extension $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{\mathcal{F}}_t, \tilde{P})$ of the probability space $(\Omega, \mathcal{F}, \hat{\mathcal{F}}_t, P)$ and a $\tilde{\mathcal{F}}_t$ -Brownian Motion B(t), such that $B(t) = M(\tau_t)$ for $t \in [0, \langle M \rangle(+\infty))$. Moreover $M(t) = B(\langle M \rangle(t))$.

We also have the following useful lemma:

Lemma 2.3 ([12], Lemma 34.8) Let $M \in \mathcal{M}_2^{c,loc}$. Then on the set $\{\langle M \rangle (+\infty) < +\infty\}$

$$\lim_{t\to+\infty}M(t)$$

exists and is finite.

Now, we are ready to review some basic facts on stochastic differential equations. These will be used in Sections 3 and 4.

2.4 ONE DIMENSIONAL, TIME-HOMOGENOUS STOCHASTIC DIFFERENTIAL EQUATIONS

Definition 2.8 Let b and σ be real measurable functions on the real line. Let $(\Omega, \mathcal{F}, \mathcal{F}_t, P)$ be a probability space with filtration \mathcal{F}_t . Let W(t) be a \mathcal{F}_t -Brownian Motion. We say that a stochastic process X(t) is a solution of stochastic differential equation

(3)
$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t),$$

with the initial condition $X(0) = x_0$ if and only if

$$X(t) = x_0 + \int_0^t b(X(s))ds + \int_0^t \sigma(X(s))dW(s) \text{ a.s. for every } t.$$

The second integral in the equation above is the stochastic integral with respect to the Brownian Motion W, which was defined in Section 2.1.

In the work that follows we assume that (3) has a unique solution starting at x (for all x), and the functions b and σ are defined on the real line. The most general conditions which guarantee the existence and uniqueness of solutions for such equations are as follows:

Theorem 2.4 ([4], Theorem IV-3.2) Let us assume the following:

- 1. There exists strictly increasing function $\rho(u)$ on $[0,+\infty)$ such that $\rho(0) = 0$ and $\int_{0+\frac{1}{\rho^2(u)}} du = +\infty$ and $|\sigma(x) \sigma(y)| \le \rho(|x-y|)$ for all $x, y \in R$.
- 2. There exists strictly increasing, concave function k(u) on $[0,+\infty)$ with k(0)=0 and $\int_{0+} \frac{1}{k(u)} du = +\infty$, such that $|b(x)-b(y)| \le k(|x-y|)$ for all $x,y \in R$.

Then there exists a unique solution of (3) with the initial condition X(0) = x.

We see that the conditions of Theorem 2.4 are satisfied when functions b and σ are Lipschitz continuous. In this case $\rho(x) = Lx, k(x) = Kx$, where L and K are Lipschitz constants for b and σ , respectively.

In Section 4 we will also need the following comparison theorem due to Skorokhod. It was first proved by Skorokhod (see [13], Section V-3), and extended by other authors.

Theorem 2.5 ([4], Theorem VI-1.1) Let σ be a function satisfying condition 1 of Theorem 2.4. Let b_1 and b_2 be two functions satisfying condition 2 of Theorem 2.4. Consider two stochastic differential equations:

$$dX(t) = b_1(X(t))dt + \sigma(X(t))dW(t).$$

$$dY(t) = b_2(Y(t))dt + \sigma(Y(t))dW(t).$$

Assume that $b_2 \ge b_1$ and $Y(0) = y_0 \ge x_0 = X(0)$. Then for all $t, Y(t) \ge X(t)$ almost surely.

We will now discuss the asymptotic behavior of the solutions of equation (3). Assume that the state space of the solutions of (3) is the interval (l,r), where l and r are finite or infinite. Assume also, that $\sigma(x) \neq 0$ for all $x \in (l,r)$. Let X(t) be the solution of (3) and X(0) = x.

Let $\eta = \inf\{t : X(t) = l \text{ or } X(t) = r\}$ be the explosion time. Define the scale function :

$$Q(x) = \int_{x_0}^{x} e^{-2\int_{x_0}^{y} \frac{b(u)}{\sigma^2(u)} du} dy \text{ where } x_0 \in (l, r)$$

The following Theorem describes the behavior of solutions of (3) as t approaches the explosion time η :

Theorem 2.6 ([4], Theorem VI-3.1) We have the following:

1) If
$$Q(l+) = -\infty$$
 and $Q(r-) = +\infty$, then

$$P_x(\eta = +\infty) = P_x(\liminf_{t \to \eta} X(t) = l) = P_x(\limsup_{t \to \eta} X(t) = r) = 1.$$

2) If
$$Q(l+) > -\infty$$
 and $Q(r-) = +\infty$, then

$$P_x(\lim_{t\to\eta}X(t)=l)=1.$$

3) If
$$Q(l+) = -\infty$$
 and $Q(r-) < +\infty$, then

$$P_x(\lim_{t\to n}X(t)=r)=1.$$

4) If
$$Q(l+) > -\infty$$
 and $Q(r-) < +\infty$, then

$$P_x(\lim_{t \to \eta} X(t) = l) = 1 - P_x(\lim_{t \to \eta} X(t) = r) = \frac{Q(r-) - Q(x)}{Q(r-) - Q(l+)}.$$

Definition 2.9 ([9], Section III-7 and III-8) We say that the solution of (3) with initial condition $X(0) = x_0$ is recurrent in the interval (l,r) if for every $y \in (l,r)$, P(X(t) = y for infinitely many t) = 1.

From Theorem 2.6, we see, that when $P(\eta = +\infty) = 1$, then every solution of (3) is recurrent if and only if $|Q(l+)| = |Q(r-)| = +\infty$.

Let us now define:

$$k(x) = \int_{x_0}^x Q'(y) \int_{x_0}^y \frac{1}{\sigma^2(z)Q'(z)} dz dy \text{ where } x_0 \in (l,r).$$

We have:

Lemma 2.4 ([4], Corollary VI-3.1) If $k(r-) < +\infty$, then $Q(r-) < +\infty$. If $k(l+) < +\infty$, then $Q(l+) > -\infty$.

The theorem below gives the conditions under which the explosion occurs (Feller test of explosion).

Theorem 2.7 ([4], Theorem VI-3.2)

a) If
$$k(r-) = k(l+) = +\infty$$
, then

$$P_x(\eta = +\infty) = 1$$
 for all $x \in (l, r)$

b) If
$$k(r-) < +\infty$$
 or $k(r+) < +\infty$, then

$$P_x(\eta < +\infty) > 0$$
 for all $x \in (l,r)$

- c) $P_x(\eta < +\infty) = 1$ for all $x \in (l,r)$ if and only if one of the following holds :
 - $(i) \quad k(r-)<+\infty \ and \ k(l+)<+\infty.$
 - (ii) $k(r-) < +\infty$ and $Q(l+) = -\infty$.
 - (iii) $k(l+) < +\infty$ and $Q(r-) = +\infty$.

3 DIFFUSIONS ON THE FINITE INTERVAL

As in the introduction let us consider the stochastic differential equation :

(4)
$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t), X(0) = x.$$

The following results are preliminary to later work and they are mostly known. However, we include derivations of them for completeness of the presentation, and some proofs may be of independent interest.

Throughout this Section we will assume that the coefficients b and σ are Lipschitz continuous.

Let us consider the case where for some numbers c and d (c < d) we have :

- (i) $\sigma(c) = \sigma(d) = 0$.
- (ii) for all $p \in (c, d)$, $\sigma(p) \neq 0$.
- (iii) $b(c) \ge 0$ and $b(d) \le 0$.

We consider the solutions of (4) starting from the interior of (c,d), that is $x \in (c,d)$. Let Z(t) and Y(t) satisfy the equation (4) with initial conditions Z(0) = c and Y(0) = d respectively. Let \underline{b} and \overline{b} be the Lipschitz continuous functions, for which we have $\underline{b} \leq b \leq \overline{b}$ and $\underline{b}(c) = \overline{b}(d) = 0$. We see, that Z(t) = c satisfies the stochastic differential equation:

$$d\underline{Z}(t) = \underline{b}(\underline{Z}(t))dt + \sigma(\underline{Z}(t))dW(t) , \ \underline{Z}(0) = c$$

Similarly, we see, that $\bar{Y}(t) = d$ satisfies the stochastic differential equation :

$$d\bar{Y}(t) = \bar{b}(\bar{Y}(t))dt + \sigma(\bar{Y}(t))dW(t)$$
, $\bar{Y}(0) = d$

From Comparison Theorem (Theorem 2.5) we conclude, that if X(t) is the solution of (4) with the initial condition $X(0) = x \in (c, d)$, then $c = \underline{Z}(t) \le Z(t) \le X(t) \le Y(t) \le \overline{Y}(t) = d$. Therefore, under assumptions (i)-(iii), the solution starting from interior of (c,d) will never leave [c,d]. Let $X_1(t)$ and $X_2(t)$ be two solutions of (4) starting from x_1 and x_2 , respectively, $x_1 < x_2$. Let Q(x) be the scale function, that is

$$Q(x) = \int_{x_0}^x e^{-2\int_{x_0}^y \frac{b(u)}{e^2(u)} du} dy \text{ for some } x_0 \in (c, d).$$

We will use the following lemmas:

Lemma 3.1 a) If b(c) > 0, then $Q(c) = -\infty$.

b) If
$$b(d) < 0$$
, then $Q(d) = +\infty$.

Proof.

a) Since σ is Lipschitz continuous,

$$\frac{1}{\sigma(x)^2} \ge \frac{1}{L^2(x-c)^2} \text{ for all } x \in (c,d) \text{ and some constant } L.$$

Since b(c) > 0, $b(x) > \epsilon$ for all $x \in (c, c + \delta)$ for some δ and ϵ . Without loss of generality we can assume that $x_0 < c + \delta$. We have:

$$Q(c) = \int_{x_0}^{c} e^{-2\int_{x_0}^{y} \frac{b(u)}{\sigma^2(u)} du} dy < -\int_{c}^{x_0} e^{2\int_{y}^{x_0} \frac{\epsilon}{L^2(u-c)^2} du} dy$$
$$= -\int_{c}^{x_0} e^{\frac{2\epsilon}{L^2} (\frac{1}{y-c} - \frac{1}{x_0-c})} dy = -\infty.$$

The proof of b) is entirely similar to the proof of a).

Lemma 3.2 (see [9], Lemma V-2.1) Let X(t) be the solution of (4) with initial condition $X(0) = x_0$, where $x_0 \in (c, d)$. Then if functions b and σ are Lipschitz continuous, then

$$P(X(t) = c \text{ for some } t) = 0 \text{ and } P(X(t) = d \text{ for some } t) = 0.$$

From Lemma 3.2 we conclude that the explosion time

 $\eta = \inf\{t : X(t) = c \text{ or } d\} = +\infty \text{ a.s. From Comparison Theorem (Theorem 2.5), } X_1(t) \leq X_2(t) \text{ for all t a.s. We have the following three cases :}$

CASE 1.
$$Q(c) > -\infty$$
, $Q(d) < +\infty$

In this case by Lemma 3.1 we have b(c)=b(d)=0. Since $X_1(t) \leq X_2(t)$, then

$$\{\omega: X_1(t) \to d\} \subseteq \{\omega: X_2(t) \to d\} \text{ and } \{\omega: X_2(t) \to c\}$$

$$\subseteq \{\omega: X_1(t) \to c\}$$

From Theorem 2.6 and Lemma 3.2 we conclude

$$P(\lim_{t \to +\infty} X_2(t) - X_1(t) = 0) = P(\lim_{t \to +\infty} X_1(t) = d) + P(\lim_{t \to +\infty} X_2(t) = c)$$

$$= 1 - \frac{Q(x_2) - Q(x_1)}{Q(d) - Q(c)}$$

and

$$P(\lim_{t \to +\infty} X_2(t) - X_1(t) = d - c) = P(\lim_{t \to +\infty} X_1(t) = c, \lim_{t \to +\infty} X_2(t) = d)$$

$$= \frac{Q(x_2) - Q(x_1)}{Q(d) - Q(c)}.$$

CASE 2. $Q(c) = -\infty$, $Q(d) < +\infty$ or $Q(c) > -\infty$, $Q(d) = +\infty$ If $Q(c) = -\infty$ and $Q(d) < +\infty$, then by Theorem 2.6 and Lemma 3.2 $X_1(t) \to d$ and $X_2(t) \to d$ as $t \to +\infty$. If $Q(c) > -\infty$ and $Q(d) = +\infty$, then by the same argument $X_2(t) \to c$ and $X_1(t) \to c$ as $t \to +\infty$. Therefore in these cases $X_2(t) \to X_1(t) \to 0$ a.s.

 $X_1(t) \to 0$ a.s.

CASE 3. $Q(c) = -\infty$, $Q(d) = +\infty$ Let $Y_2(t) = Q(X_2(t))$ and $Y_1(t) = Q(X_1(t))$. Then as it was explained in Section 1. $Y_1(t)$ satisfies the stochastic differen-

explained in Section 1, $Y_i(t)$ satisfies the stochastic differential equation : $dY_i(t) = \sigma_1(Y_i(t))dW(t)$ with initial condition $Y_i(0) = Q(x_i)$, i=1,2. Therefore we conclude from Section 1 that $Y_2(t) - Y_1(t) \to \xi$. $\xi = 0$ when $\sigma_1(y)$ is not periodic and ξ is concentrated on two points if $\sigma_1(y)$ is periodic.

We have the following:

Lemma 3.3 If b(c) > 0 or b(d) < 0, then the function $\sigma_1(y) = \sigma(Q^{-1}(y))Q'(Q^{-1}(y))$ cannot be periodic.

Proof.

Without loss of generality we can assume that c=0. Assume that b(0)>0. Since b is continuous, there exists $\epsilon>0$ such that $b(x)>\epsilon$ for all $x\in[0,\delta]$ for some δ . Let z_0 be such that $\frac{2b(z)}{\sigma(z)}>L$ for all $z\leq z_0$, where L is the Lipschitz constant for σ . We will show that $f(z)=\sigma(z)e^{2\int_z^\delta \frac{b(u)}{\sigma^2(u)}du}$ is decreasing in the

interval $[0, z_0]$. We have :

$$\frac{f(z+h)-f(z)}{h} = \frac{\sigma(z+h)e^{2\int_{s+h}^{\delta} \frac{b(u)}{\sigma^2(u)}du} - \sigma(z)e^{2\int_{s}^{\delta} \frac{b(u)}{\sigma^2(u)}du}}{h}$$

$$= \sigma(z+h)\frac{e^{2\int_{s+h}^{\delta} \frac{b(u)}{\sigma^2(u)}du} - e^{2\int_{s}^{\delta} \frac{b(u)}{\sigma^2(u)}du}}{h} + \frac{\sigma(z+h)-\sigma(z)}{h}e^{2\int_{s}^{\delta} \frac{b(u)}{\sigma^2(u)}du}$$

$$\leq \sigma(z+h)\frac{e^{2\int_{s+h}^{\delta} \frac{b(u)}{\sigma^2(u)}du} - e^{2\int_{s}^{\delta} \frac{b(u)}{\sigma^2(u)}du}}{h} + Le^{2\int_{s}^{\delta} \frac{b(u)}{\sigma^2(u)}du}$$

The right hand side of the inequality above tends to $(L-2\frac{b(z)}{\sigma(z)})e^{2\int_{z}^{\delta}\frac{b(u)}{\sigma^{2}(u)}du} \leq 0$ as $h \to 0$. This means that for all $z \leq z_{0}$, there exists h such that f(z) is decreasing in [z, z + h]. Therefore, we can conclude that f(z) is decreasing on $[0, z_{0}]$. Since f(z) is positive and decreasing for z small enough, $\lim_{z \to 0} f(z)$ exists (finite or infinite). It is easy to see that $\lim_{y \to -\infty} Q^{-1}(y) = 0$, and therefore $\lim_{y \to -\infty} \sigma_{1}(y)$ exists. Therefore $\sigma_{1}(y)$ cannot be periodic.

The same proof holds for the case b(d) < 0. This completes the proof.

Therefore in CASE 3 if b(c) > 0 or b(d) < 0, then σ_1 cannot be periodic, so $Y_2(t) - Y_1(t) \to 0$. Since Q^{-1} is uniformly continuous, $X_2(t) - X_1(t) \to 0$ too. Therefore, if b(c) > 0 or b(d) < 0, then $X_2(t) - X_1(t) \to 0$. Unfortunately, σ_1 can be periodic (or constant) when $b(c) = \sigma(c) = 0$, and $b(d) = \sigma(d) = 0$ in case d is finite. In such cases the convergence takes place in the scale metric and not in the Euclidean metric as the following examples show.

Example 3.1 Consider the following stochastic differential equation on $(0,+\infty)$:

$$dX(t) = \frac{1}{2}X(t)dt + X(t)dW(t)$$

In this case $Q(x) = \ln(x)$ and $\sigma_1(y) = 1$. The solution of this equation is $X(t) = xe^{W(t)}$, which is called the geometric Brownian Motion.

Example 3.2 Consider the following stochastic differential equation on $\left(-\frac{\pi}{2}, +\frac{\pi}{2}\right)$:

$$dX(t) = -\frac{1}{2}sin(2(X(t)))cos^{2}(X(t))dt + cos^{2}(X(t))dW(t).$$

In this case, $Q(x) = \tan x$ and $\sigma_1(y) = 1$. The solution of this equation is $X(t) = \arctan(\tan x + W(t))$.

We can easily see that in these two examples $X_2(t) - X_1(t)$ does not have a limit as $t \to +\infty$, but the limit exists in the sense of convergence in the "scale metric" (r(x,y) = |Q(y) - Q(x)|), since $\ln y e^{W(t)} - \ln x e^{W(t)} = \ln \frac{y}{x}$ and $\tan y + W(t) - \tan x - W(t) = \tan y - \tan x$, respectively.

4 SINGULAR DIFFUSIONS

In this Section we consider again the stochastic differential equation:

(5)
$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t), X(0) = x$$

We investigate convergence of solutions starting from two different points, but we allow $\sigma = 0$. Let us start with an example that shows that in some cases it seems to be impossible to find any sense of convergence for $X_2(t) - X_1(t)$. Therefore some classification of points on the real line seems to be necessary.

Example 4.1 Let coefficients b and σ of equation (5) be such that :

- (i) for some numbers $c, p, d, c , we have <math display="block">\sigma(c) = \sigma(p) = \sigma(d) = 0.$
- (ii) for all v in (c, p), $\sigma(v) \neq 0$.
- (iii) for all v in (p, d), $\sigma(v) \neq 0$.

Let us also assume that b(c) = 0, b(p) > 0, b(d) < 0, and for some $x_0 \in (c, p)$ we have

$$\int_{c}^{x_0} e^{2\int_{y}^{x_0} \frac{b(u)}{\sigma^2(u)} du} dy < +\infty.$$

Consider two solutions of our equation $X_1(t)$ and $X_2(t)$ with initial conditions x_1 and x_2 $(x_1 < x_2)$, respectively, both in (c,p). Then

$$P(\lim_{t \to +\infty} X_1(t) = c, \lim_{t \to +\infty} X_2(t) = c) = P(\lim_{t \to +\infty} X_2(t) = c) = \frac{Q(p) - Q(x_2)}{Q(p) - Q(c)}$$

$$P(for some \ t_1, t_2 : X_1(t_1) = p \ and \ X_2(t_2) = p) = P(X_1(t_1) = p \ for \ some \ t_1)$$

$$= \frac{Q(x_1) - Q(c)}{Q(p) - Q(c)}$$

$$P(\lim_{t \to +\infty} X_1(t) = c, \text{ for some } t_2 : X_2(t_2) = p)$$

$$= P(\lim_{t \to +\infty} X_1(t) = c, \liminf_{t \to +\infty} X_2(t) = p, \limsup_{t \to +\infty} X_2(t) = d)$$

$$= 1 - \frac{Q(p) - Q(x_2)}{Q(p) - Q(c)} - \frac{Q(x_1) - Q(c)}{Q(p) - Q(c)} = \frac{Q(x_2) - Q(x_1)}{Q(p) - Q(c)}$$

These calculations follow from work in Section 3, Comparison Theorem (Theorem 2.5) and Theorem 2.7. Therefore, there is a positive probability (namely $\frac{Q(x_2)-Q(x_1)}{Q(d)-Q(c)}$) that $X_1(t) \to c$, $\liminf_{t \to +\infty} X_2(t) = p$ and $\limsup_{t \to +\infty} X_2(t) = d$, so with positive probability the limit of $X_2(t) - X_1(t)$ does not exist.

Let $A = \{x: \sigma(x) = 0\}$. From now on, $X^x(t)$ will denote the solution of (5) with the initial condition $X^x(0) = x$. We make the following assumptions:

- I. For every real number x one of the following holds:
 - (i) There exists z > x, such that $0 = \sigma(z) = b(z)$.
 - (ii) There exist y > x and z > x, such that $\sigma(y) = \sigma(z) = 0$ and $b(y)b(z) \le 0$.
 - (iii) $sup\{x: x \in A\} < +\infty$, that is A is bounded from above.
- II. For every real number x one of the following holds:
 - (i) There exists z < x, such that $0 = \sigma(z) = b(z)$.
 - (ii) There exist y < x and z < x, such that $\sigma(y) = \sigma(z) = 0$ and $b(y)b(z) \le 0$.
 - (iii) $inf\{x: x \in A\} > -\infty$, that is A is bounded from below.

We will use the following notation:

$$\tau_x = \inf\{t: X(t) = x\}, P_y(\tau_x < +\infty) = P(X^y(t) = x \text{ for some } t).$$

 au_x is the time when process X(t) hits a point x for the first time.

Note, that τ_x is the whole collection of random variables, since it depends on X(0).

 $P_y(\tau_x < +\infty)$ is the probability, that a process starting at y reaches x in finite time.

Similarly, for the subset B of the real line, we introduce

$$\tau_B = \inf\{t : X(t) \in B\}, P_v(\tau_B < +\infty) = P(X^v(t) \in B \text{ for some } t).$$

Definition 4.1 We say that x and y communicate $(x \sim y)$ if and only if $P_x(\tau_y < +\infty) > 0$ and $P_y(\tau_x < +\infty) > 0$.

Relation "~" defines the equivalence classes in R.

Definition 4.2 We say that x is strictly inessential if and only if for some y < x and z > x

$$P(\limsup_{t\to+\infty}X^x(t)\leq y)>0 \ and \ P(\liminf_{t\to+\infty}X^y(t)\geq z)>0.$$

CASE 1 from Section 3 illustrates the set of the strictly inessential points.

Proposition 4.1 Let x be strictly inessential, and let I_x be its equivalence class under " \sim ". Then

- a) I_x consists of at least two points.
- b) $\sigma(z) \neq 0$ for all z in I_x .
- c) I_x is an open set.
- d) I_x is a connected set.
- e) If $y \sim x$, then y is strictly inessential.
- f) If I_x is bounded, i.e., $I_x = (c,d)$, then $\sigma(c) = \sigma(d) = 0$ and $b(c) \le 0 \le b(d)$.

Proof.

- a) We have $\sigma(x) \neq 0$ because otherwise x would not be strictly inessential. So there exists a neighborhood U of x such that $\sigma(y) \neq 0$ for all $y \in U$. So $x \sim y$ for all $y \in U$.
- b) Suppose that there exists z in I_x (z > x), such that $\sigma(z) = 0$. Then if $b(z) \ge 0$, then $P_z(\tau_x < +\infty) = 0$ so x and z do not communicate. If b(z) < 0, then $P_x(\tau_z < +\infty) = 0$ so x and z do not communicate.
- c) follows from a) and b).

- d) If $y \in I_x$ and $z \in I_x$, then for all u such that $y \le u \le z$ we have $P_y(\tau_z < +\infty) = P_y(\tau_u < +\infty)P_u(\tau_z < +\infty) > 0 \text{ and}$ $P_z(\tau_y < +\infty) = P_z(\tau_u < +\infty)P_u(\tau_y < +\infty) > 0, \text{ so } u \in I_x. \text{ Therefore } I_x \text{ is connected and it must be an open interval.}$
- e) We know that I_x is an interval, say (c,d) and $\sigma(z) \neq 0$ for all z in I_x . Therefore we can define the scale function:

$$Q(z) = \int_{z}^{z} e^{-2\int_{z}^{y} \frac{b(u)}{\sigma^{2}(u)} du} dy.$$

Since x is strictly inessential, then Q(z) must be bounded and then for all v in I_x (see Theorem 2.6)

$$P(\limsup_{t\to +\infty} X^v(t) \leq c) = \frac{Q(d)-Q(v)}{Q(d)-Q(c)} > 0$$

$$P(\liminf_{t\to +\infty} X^v(t)\geq d)=1-\frac{Q(d)-Q(v)}{Q(d)-Q(c)}>0$$

So v is strictly inessential for all v in I_x .

f) Suppose that $\sigma(c) \neq 0$. Then from d) there exists y < c such that $y \sim x$ so $y \in I_x$ which is the contradiction. So $\sigma(c) = 0$. The same way we can show that $\sigma(d) = 0$. Now suppose that b(c) > 0. Then we conclude from Section 3 that $P(\limsup_{t \to +\infty} X^x(t) \geq d) = 1$, so x cannot be strictly inessential. Therefore we must have $b(c) \leq 0$. The same way we show that $b(d) \geq 0$.

Definition 4.3 We say that x is right inessential if

i) x is not strictly inessential.

and

ii) There exists z > x such that $P_x(\tau_x < +\infty) > 0$ and $P_z(\tau_x < +\infty) = 0$.

Definition 4.4 We say that x is left inessential if

i) x is not strictly inessential.

and

ii) There exists z < x such that $P_x(\tau_x < +\infty) > 0$ and $P_z(\tau_x < +\infty) = 0$.

Proposition 4.2 Let x be right inessential and let I_x be its equivalence class under " \sim ". Then

- a) $I_x = \{x\}$ if and only if $\sigma(x) = 0$ and b(x) > 0.
- b) If I_x consists of at least two points, then I_x is an open and connected set, and $\sigma(z) \neq 0$ for all z in I_x .
- c) If I_x is bounded, i.e., $I_x = (c, d)$, then $\sigma(c) = \sigma(d) = 0$, $b(c) \ge 0$ and b(d) > 0.
- d) y is right inessential for all y in I_x .

Proof.

- a) Suppose that $I_x = \{x\}$. Then $\sigma(x) = 0$ because otherwise x will communicate with y for every $y \in U$, where U is some neighborhood of x. If $b(x) \leq 0$, then for every y > x $P_x(\tau_y < +\infty) = 0$ and then x is not right inessential. So b(x) > 0. If $\sigma(x) = 0$ and b(x) > 0, then for every y > x $P_y(\tau_x < +\infty) = 0$ and for every z < x $P_x(\tau_z < +\infty) = 0$ so x cannot communicate with any other point. So $I_x = \{x\}$.
- b) If there exists $z \in I_x$ such that $\sigma(z) = 0$, then x and z cannot communicate by the same argument as in Proposition 4.1(b). Therefore $\sigma(z) \neq 0$ for all z in I_x . From that we conclude, that for all $z \in I_x$ there exists a neighborhood U of z, such that $\sigma(u) \neq 0$ for all $u \in U$. Then, $u \sim z$ for all $u \in U$, so since $z \in I_x$ and $U \subseteq I_x$, then I_x is open. The fact that I_x is connected follows by the same argument as in Proposition 4.1(d).
- c) $\sigma(c) = \sigma(d) = 0$ for the same reason as in Proposition 4.1(e). If $b(d) \leq 0$, then for every z > d $P_x(\tau_z < +\infty) = 0$ and for every z such that $x \leq z \leq d$ we have $x \sim z$. So x cannot be right inessential. Therefore b(d) > 0. Suppose now that b(c) < 0. Since $\sigma(z) \neq 0$ for all $z \in (c, d)$, then we can define

$$Q(z) = \int_{z}^{z} e^{-2\int_{z}^{y} \frac{b(u)}{\sigma^{2}(u)} du} dy.$$

Since b(d) > 0 and b(c) < 0, then Q is bounded in (c,d), $P(\limsup_{t \to +\infty} X^x(t) \le c) > 0$ and $P(\liminf_{t \to +\infty} X^x(t) \ge d) > 0$, so x is strictly inessential which is the contradiction.

d) Follows from (c).

Similar facts are true when x is left inessential.

Corollary 4.1 a) If x is right inessential, then I_x must be bounded from above.

b) x is right inessential if and only if there exists z > x such that $P_x(\tau_z < +\infty) = 1$ and $P_z(\tau_x < +\infty) = 0$.

Similar facts are true when x is left inessential.

Definition 4.5 We say that x is essential if the following two conditions hold:

- (i) For all y such that $P_x(\tau_y < +\infty) > 0$ we have $P_y(\tau_x < +\infty) > 0$
- (ii) For all y < x and for all z > x, $P(\limsup_{t \to +\infty} X^x(t) \le y) = 0$ or $P(\liminf_{t \to +\infty} X^x(t) \ge z) = 0$.

CASES 2 and 3 from Section 3 are examples of essential states.

Proposition 4.3 Let x be essential, and let I_x be its equivalence class under \sim .

a) $I_x = \{x\}$ if and only if $\sigma(x) = b(x) = 0$.

- b) If I_x consists of at least two points, then I_x is an open and connected set, and $\sigma(z) \neq 0$ for all z in I_x .
- c) If I_x is bounded, i.e., $I_x = (c, d)$, then $\sigma(c) = \sigma(d)$, $b(c) \ge 0$, and $b(d) \le 0$.
- d) If $y \sim x$, then y is essential.

Proof.

- a) Assume that I_x = {x}. If σ(x) ≠ 0, then x communicate with points of some neighborhood of x, so I_x ≠ {x}. If b(x) > 0, then x is right inessential. If b(x) < 0, then x is left inessential. Therefore b(x) = 0.
 If σ(x) = b(x) = 0, then obviously I_x = {x}. Such point is called a trap (see [5], Section 3.4).
- b) The proof of this statement is entirely similar to the proof of Proposition 4.2(b).
- c) As before let

$$Q(z) = \int_x^z e^{-2\int_z^y \frac{b(u)}{\sigma^2(u)} du} dy.$$

Because x is essential Q cannot be bounded in (c,d) (if it is bounded, then x is strictly inessential). If b(c) < 0, then Q is bounded from below so $Q(z) \to +\infty$ as $z \to d$ and then x is left inessential. If b(d) > 0, then Q is bounded from above so $Q(z) \to -\infty$ as $z \to c$ and then x is right inessential. So $b(c) \ge 0$ and $b(d) \le 0$.

d) If $I_x = \{x\}$, then there is nothing to prove. If $I_x \neq \{x\}$, then I_x is an interval, say (c,d) and Q from (c) is not bounded. If $Q(y) \to +\infty$ as $y \to d$, then for every $z \in (c,d)$ $P(\liminf_{t \to +\infty} X^z(t) = c) = 1$ so for all y > c $P(\liminf_{t \to +\infty} X^z(t) \ge y) = 0$. If $Q(y) \to -\infty$ as $y \to c$, then for all $z \in (c,d)$ $P(\limsup_{t \to +\infty} X^z(t) = d) = 1$ so for every y < d $P(\limsup_{t \to +\infty} X^z(t) \le y) = 0$ Therefore (ii) from Definition 4 is satisfied. (i) from Definition 4 is satisfied because for all $z \notin I_x$ $P_x(\tau_z < +\infty) = 0$ from (c).

Note that strictly inessential, right inessential, left inessential and essential points cover all points in R^1 . Distinguishing between them requires the knowledge of the scale function Q in each equivalence class with respect to \sim .

Proposition 4.4 Suppose that x is not strictly inessential. Then one of the following holds:

- a) There exists an essential point z such that $P_x(\tau_{I_x} < +\infty) = 1$, and if z_1 and z_2 are two such points, then $z_1 \sim z_2$.
- b) There exists z, such that $P(\lim_{t\to +\infty} X^x(t) = z) = 1$.

Proof.: We will use the assumptions I and II from the beginning of this Section. If x is essential, then there is nothing to prove because (a) holds. Suppose that x is right inessential. We cannot have $x > \sup\{z : z \in A\}$, because then x would be essential or strictly inessential. Let us assume that

either the condition (i) or (ii) from Assumption I holds. Then there exists z > x such that $\sigma(z) = 0$ and $b(z) \le 0$.

Let $d = \inf\{z \in A : b(z) \le 0 \text{ and } z > x\}$. Then $\sigma(d) = 0$ and $b(d) \le 0$ from continuity of b and σ .

Since x is right inessential it follows from the Corollary 4.1, that there exists y > x such that $P_x(\tau_y < +\infty) = 1$ and $P_y(\tau_x < +\infty) = 0$.

Let $B = \{z \geq x : P_x(\tau_x < +\infty) = 1 \text{ and } P_x(\tau_x < +\infty) = 0\}$. We will show that B consists of at least two points. Clearly one of them is y, such that y is the upper bound for I_x . Since b(y) > 0 and b is continuous, then there exists \underline{b} and δ , such that $b(u) > \underline{b}$ for all $u \in (y - \delta, y + \delta)$. Let $b_1(u) = b(u) \wedge \underline{b}$. Let Y(t) be the solution of the equation:

$$dY(t) = b_1(Y(t))dt + \sigma(Y(t))dW(t) Y(0) = u \in (y - \delta, y + \delta)$$

From comparison theorem (Theorem 2.5) we have that for all $t \ Y(t) \le X^u(t)$ a.s. Let $\tau = \inf\{t : Y(t) \notin (y - \delta, y + \delta)\}$. We will show that $E_u \tau < +\infty$. Let $f(u) = \frac{y + \delta - u}{b}$. Then f(u) > 0 in $(y - \delta, y + \delta)$ and

$$\frac{1}{2}\sigma^2(u)f''(u) + \underline{b}f'(u) = -1$$

So $E_u \tau < +\infty$ for all $u \in (y - \delta, y + \delta)$. It is obvious that $\tau \geq \tau_{y+\delta}$ P_u -a.s. for all $u \in [y, y + \delta]$. Therefore $P_y(\tau_u < +\infty) = 1$ for all $u \in [y, y + \delta)$ and we have $P_x(\tau_u < +\infty) = P_x(\tau_y < +\infty)P_y(\tau_u < +\infty) = 1$ and $P_u(\tau_x < +\infty) = P_u(\tau_y < +\infty)P_y(\tau_x < +\infty) = 0$. So $u \in B$ for all $u \in [y, y + \delta)$. Therefore B consists of at least two points. It is obvious that B is connected because if we have $P_x(\tau_y < +\infty) = 1$, $P_y(\tau_x < +\infty) = 0$, $P_x(\tau_x < +\infty) = 1$, $P_x(\tau_x < +\infty) = 0$ for some y and z y < z, then

 $P_x(\tau_u < +\infty) \ge P_x(\tau_s < +\infty) = 1$ and $P_u(\tau_x < +\infty) \le P_y(\tau_x < +\infty) = 0$ for every u such that $y \le u \le z$.

So B must be an interval. Let $c = \sup\{z : z \in B\}$. We will show that $\sigma(c) = 0$ and $b(c) \leq 0$. If $\sigma(c) \neq 0$, then $\sigma(u) \neq 0$ for all u such that $u \in (c-\delta, c+\delta)$. Let $c_1 = \sup\{x \leq c : \sigma(x) = 0\}$. We have $c_1 < c$, $b(c_1) > 0$, so Lemma 3.1 and Theorem 2.6 imply that

$$\limsup_{t\to+\infty}X^x(t)>c.$$

Therefore, we conclude that for $u \in [c, c+\delta)$ we have that $P_x(\tau_u < +\infty) = 1$ and $P_u(\tau_x < +\infty) = 0$ which contradicts the definition of c. So $\sigma(c) = 0$. Assume now that b(c) > 0. If $\sigma(u) \neq 0$ for $u \in (c - \delta, c)$ for some $\delta > 0$, then $c \in B$ and by the same argument as above we conclude that there is a $\delta' > 0$ such that $u \in B$ for all $u \in (c, c + \delta')$, what is again contradiction with definition of c. So assume that there exists z_n such that $\sigma(z_n) = 0$ for all $n, z_n < z_{n+1}$, and $\lim_{n \to +\infty} z_n = c$. Since b(c) > 0, then there exists n_0 , such that for all $n \geq n_0$ b(z) is bounded away from zero for every $z \in [z_{n_0}, c]$. Let $b \geq 0$ be such that $b(z) \geq b$ for every $z \in [z_{n_0}, c]$. Let Y(t) be a solution of the equation:

$$dY(t) = \underline{b}dt + \sigma(Y(t))dW(t) , Y(0) = z_{n_0+1}$$

From comparison theorem we have that $Y(t) \leq X^{z_{n_0+1}}(t)$ a.s. for all t. Let $\eta = \inf\{t: Y(t) \notin (z_{n_0}, c)\}$. We will show that $E_{z_{n_0+1}}\eta < +\infty$. Let $f(x) = \frac{c-x}{b}$. Then f(x) > 0 in (z_{n_0}, c) and

$$\frac{1}{2}\sigma^2(x)f''(x) + \underline{b}f'(x) = -1$$

So $E_{z_{n_0+1}}\eta < +\infty$.

Since for all $t Y(t) \leq X^{z_{n_0+1}}(t)$, then $E_{z_{n_0+1}}\tau_c < +\infty$ so

 $P_{z_{n_0+1}}(\tau_c < +\infty) = 1$. Therefore, $P_x(\tau_c < +\infty) = 1$ and since b(c) > 0, then $P_x(\tau_u < +\infty) = 1$ for all $u \in (c, c+\delta)$ for some δ . This again is the contradiction with the definition of c. So we must have $\sigma(c) = 0$, $b(c) \le 0$ and therefore c = d. Now if there exists c < d such that $\sigma(c) = 0$ and $\sigma(c) \ne 0$ for all $c \in (c, d)$, then we must have that $c \in (c, d)$ is an interval of essential points and $P_x(\tau_{(c,d)} < +\infty) = 1$. If there exists an increasing sequence $c \in (c, d)$ such that $c \in (c, d)$ for all $c \in (c, d)$, then since $c \in (c, d)$ and $c \in (c, d)$ for every $c \in (c, d)$ for all $c \in (c, d)$ and $c \in (c, d)$ for all $c \in (c, d)$ and $c \in (c, d)$ for all $c \in (c, d)$ for all $c \in (c, d)$ and $c \in (c, d)$ for every $c \in (c, d)$ for all $c \in (c, d)$ for all c

This completes the proof of the Proposition in case (i) or (ii) from Assumption I holds. The proof of the Proposition in case (iii) from Assumption I holds is entirely similar and will not be reproduced. The proof for the case when x is left inessential can be carried out similarly.

Definition 4.6 Let x and y be two points which are not strictly inessential. Let us define the following relation:

 $x \approx y$ if and only if one of the following holds:

i) There exists an essential point z such that $P_x(\tau_{I_x} < +\infty) = 1$ and $P_y(\tau_{I_x} < +\infty) = 1$.

or

ii) There exists z such that $P(\lim_{t\to +\infty} X^x(t) = z) = P(\lim_{t\to +\infty} X^y(t) = z) = 1$.

Let us show that " \approx " defines the equivalence classes among all points which are not strictly inessential. It follows from Proposition 4.4 that $x \approx x$. It is obvious that if $x \approx y$, then $y \approx x$. So assume that $x \approx y$ and $y \approx z$. In this case, either

a) There exists an essential point r such that $P_x(\tau_{I_r} < +\infty) = 1$ and $P_y(\tau_{I_r} < +\infty) = 1$

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b) There exists a point u such that $P(\lim_{t\to +\infty} X^x(t) = u) = P(\lim_{t\to +\infty} X^y(t) = u) = 1$.

And similarly for y and z: either

c) There exists an essential point v such that $P_y(\tau_{I_v}<+\infty)=P_z(\tau_{I_v}<+\infty)=1$

or

- d) There exists a point p such that $P(\lim_{t\to +\infty} X^y(t) = p) = P(\lim_{t\to +\infty} X^z(t) = p) = 1$.
- If (a) and (c) occur, then $I_v = I_r$ and $x \approx z$.
- If (a) and (d) occur, then we must show that

$$P(\lim_{t\to+\infty}X^x(t)=p)=1.$$

Since

$$P_{y}(\tau_{I_{r}} < +\infty) = 1, \ P(\lim_{t \to +\infty} X^{y}(t) = p) = 1,$$

then Theorem 2.6 applied to the interval I_r implies that for all $u \in I_r$ $X^u(t) \to p$ a.s. Since $P_x(\tau_{I_r} < +\infty) = 1$, then from Markov property we get $P(\lim_{t \to +\infty} X^x(t) = p) = 1$ and $x \approx z$.

Similarly if we have (b) and (c), then $x \approx z$.

If we have (b) and (d), then u = p and $x \approx z$.

Thus, "≈" defines equivalence classes among all points which are not strictly inessential.

Now, we are ready to state and prove the main result of this section. It should be noted, that Theorem 4.1 below deals with the Euclidean distance between two solutions of (5). Results of Remark 4.1 following Theorem 4.1 are concerned with the convergence of the distance between the two solutions of (5) with respect to the scale metric which is weaker than the Euclidean metric.

Theorem 4.1 Let x and y be two points which are not strictly inessential.

- 1. If $x \approx y$, then one of the following holds:
 - a) There exists an essential point z, such that I_z is a right ray and $P_x(\tau_{I_z} < +\infty) = P_y(\tau_{I_z} < +\infty) = 1$.
 - b) There exists an essential point z, such that I_z is a left ray and $P_x(\tau_{I_z} < +\infty) = P_y(\tau_{I_z} < +\infty) = 1$.
 - c) $I_x = I_y = (c, d)$ for some finite numbers c and d, and $b(c) = b(d) = \sigma(c) = \sigma(d) = 0$.

d)
$$P(X^{y}(t) - X^{x}(t) \to 0) = 1$$
.

2. If
$$P(X^y(t) - X^x(t) \rightarrow 0) = 1$$
, then $x \approx y$.

Proof.

1. Assume that $x \approx y$. If there exists z such that

$$P(\lim_{t\to+\infty}X^x(t)=z)=P(\lim_{t\to+\infty}X^y(t)=z)=1,$$

then there is nothing to prove. Suppose that there exists an essential point z such that

$$P_x(\tau_{I_*} < +\infty) = P_y(\tau_{I_*} < +\infty) = 1.$$

We need to show, that d) occurs when a), b), c) do not occur. Let $\tau = \inf\{t : X^x(t) \in I_z\}$, $\eta = \inf\{t : X^y(t) \in I_z\}$. Let $\mu = \tau \vee \eta$ $(x \vee y)$ denote the maximum of x and y). Consider two dimensional Markov process $(X^x(t), X^y(t))$. Assume x < y. From CASE 3 from Section 3 and from Lemma 3.3 we conclude that for $u \in I_z$ and $v \in I_z$

$$P(\inf_{s>0}\sup_{t>s}(X^v(t)-X^u(t))\leq \delta)=1$$

From Markov property for $(X^x(t), X^y(t))$, we have for every δ

$$P(\inf_{s>\mu}\sup_{t>s}(X^{y}(t)-X^{x}(t))\leq\delta)$$

$$= E_{x,y}(P[\inf_{s\geq 0}\sup_{t\geq s}(X^{X^x(\mu)}(t)-X^{X^y(\mu)}(t))\leq \delta])=1.$$

Therefore

$$P(\inf_{s>0}\sup_{t>s}(X^y(t)-X^x(t))\leq \delta)=1$$

for every $\delta > 0$. So

$$P(\lim_{t\to +\infty}(X^x(t)-X^y(t))=0)=1.$$

2. Let us assume that

$$P(\lim_{t\to +\infty}(X^y(t)-X^x(t))=0)=1.$$

x and y are not strictly inessential. For x we have : either

a) There exists an essential point z such that $P_x(\tau_{I_x} < +\infty) = 1$

or

b) There exists a point u such that $P(\lim_{t\to +\infty} X^x(t) = u) = 1$

Similarly for y: either

c) There exists an essential point v such that $P_{v}(\tau_{I_{v}} < +\infty) = 1$

or

d) There exists p such that

$$P(\lim_{t\to +\infty} X^{y}(t)=p)=1$$

If (b) and (d) occur, then u = p and $x \approx y$. Suppose that (a) and (c) occur. If $I_x = I_v$, then $x \approx y$. Suppose that $I_x \neq I_v$. Then $I_z \cap I_v = \phi$. Let \bar{I}_z denote the closure of I_z . Since

$$P(\lim_{t\to+\infty}(X^{y}(t)-X^{x}(t))=0)=1,$$

then $\bar{I}_{v} \cap \bar{I}_{v} = \{r\}$ and

$$P(\lim_{t\to+\infty}X^x(t)=r)=P(\lim_{t\to+\infty}X^y(t)=r)=1,$$

so $x \approx y$.

If (a) and (d) occur, then $p \in \bar{I}_z$ and we must have

$$P(\lim_{t\to+\infty}X^x(t)=p)=1,$$

so $x \approx y$. Similarly for the case when (b) and (c) occur. This completes the proof.

Remark 4.1 The case when a), b) or c) from part 1 of Theorem 4.1 occur must be treated separately.

Let us consider the case when a) occurs first. Assume $I_z=(c,+\infty)$. If $x \notin I_z$ or $y \notin I_z$, then we conclude from the proof of Proposition 4.4 that b(c)>0. Assume that $Q(+\infty)=+\infty$. By Lemma 3.3, $\sigma_1(y)=\sigma(Q^{-1}(y))Q'(Q^{-1}(y))$ cannot be periodic and Theorem 1.3 (together with Markov property for the two dimensional process $(X^x(t),X^y(t))$) implies that $r(X^y(t),X^x(t))\to 0$ a.s. If $Q(+\infty)<+\infty$, then $X^x(t)\to +\infty$ and $X^y(t)\to +\infty$ a.s. Then $r(X^y(t),X^x(t))\to 0$ a.s.

If both $x \in I_z$ and $y \in I_x$, then b(c) can be zero. It follows from the discussion of Section 3, that in this case $r(X^y(t), X^x(t)) \to \xi$ a.s. for some

nonnegative random variable ξ . $\xi=0$ in case $\sigma_1(y)=\sigma(Q^{-1}(y))Q'(Q^{-1}(y))$ is not periodic, and ξ is concentrated on two points in case $\sigma_1(y)$ is periodic.

The case when b) occurs can be treated similarly.

The case when c) occurs was discussed in Section 3.

5 DIFFUSIONS WITH CONSTANT DRIFT

In this Section we again consider the equation:

(6)
$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t), X(0) = x$$

We investigate the situation when we drop the assumption III from Section 1 on the recurrence nature of the solutions of (6).

Let $X_1(t)$ and $X_2(t)$ be two solutions of this equation starting from x_1 and x_2 respectively. Deterministic examples $(\sigma = 0)$ show that in general $X_2(t) - X_1(t)$ does not have to have a limit as $t \to +\infty$. But if the drift b is nonincreasing, then we have $X_2(t) - X_1(t) \to \xi \ge 0$.

We will consider the special case where b(u) = c, where c is a positive constant; that is, we consider the stochastic differential equation:

(7)
$$dX(t) = cdt + \sigma(X(t))dW(t)$$

where c>0 is a constant and σ satisfies the conditions which guarantee the existence and uniqueness of the solutions of (7). We would like to investigate the limit of the difference of two solutions starting at two different points. Let $X_1(t)$ and $X_2(t)$ be two different solutions of (7) starting at x_1 and x_2 $(x_2 > x_1)$ respectively:

$$X_2(t) = x_2 + ct + \int_0^t \sigma(X_2(s))dW(s).$$

$$X_1(t) = x_1 + ct + \int_0^t \sigma(X_1(s))dW(s).$$

Then $Y(t) = X_2(t) - X_1(t) = x_2 - x_1 + \int_0^t [\sigma(X_2(s)) - \sigma(X_1(s))] dW(s)$ is a positive supermartingale, and therefore there is a random variable $\xi \geq 0$ such that $X_2(t) - X_1(t) \longrightarrow \xi$ a.s. Our problem is to determine for which σ the limit $\xi = 0$.

We consider first the case when the solutions of (7) are recurrent in some interval $(d,+\infty)$, where $d = \sup\{x : \sigma(x) = 0\}$. Therefore we make the assumption that $\sigma(x)$ does not vanish for x large enough.

Recall that

$$Q(x) = \int_{x_0}^x e^{-2c \int_{x_0}^y \frac{1}{e^2(u)} du} dy , \ x_0 \in (d, +\infty).$$

By Theorem 2.6, we have $Q(+\infty) = +\infty$, and therefore σ must be unbounded. We may now reproduce the proof given in [8] to conclude that if $X_2(t) - X_1(t)$ tends to the nonzero limit, then σ must be periodic. But if σ is periodic, then it must be bounded. Therefore if all solutions of (7) are recurrent in the interval $(d,+\infty)$, then $\lim_{t\to +\infty} (X_2(t) - X_1(t)) = 0$.

We now proceed to the case when all solution of (7) are transient. We will make the following assumptions:

- I. σ is differentiable, and there exists a_1 , such that for all $x > a_1$, σ' is monotone (that is σ is either convex or concave).
- II. All solutions of (7) tend to $+\infty$ as $t \to +\infty$.

Assumption II is equivalent to the assumption that all solutions of (7) are transient. From assumption I it follows that there exists b such that $|\sigma(x)| > 0$ for x > b. Without loss of generality we may assume that b = -1

and $\sigma(x) > 0$ for x > -1, so we can define a scale function:

$$Q(x) = \int_0^x e^{-2c\int_0^y \frac{1}{\sigma^2(u)}du} dy.$$

From assumption II and Theorem 2.6 it follows that Q(x) has a finite limit as $x \longrightarrow +\infty$. Let us denote this limit by d. That is $Q(x) \longrightarrow d$ as $x \longrightarrow +\infty$.

We will need the following Lemma:

Lemma 5.1 If σ is locally Lipschitz continuous, $X_1(t) \neq X_2(t)$ for all t a.s..

Proof.

Let $\tau_n = \inf\{t: X_2(t) - X_1(t) = \frac{1}{n}\}$ and $\tau = \inf\{t: X_2(t) - X_1(t) = 0\}$.

Assume $P(\tau < +\infty) > 0$.

Following ([3]) we see that:

$$X_2(t\wedge\tau_n)-X_1(t\wedge\tau_n)=$$

$$=(x_2-x_1)\exp\{\int_0^{t\wedge\tau_n}\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)}dW(s)-\frac{1}{2}\int_0^{t\wedge\tau_n}(\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)})^2ds\}$$

Letting $n \to +\infty$ we have for $t < \tau$ (see also [6]):

$$X_2(t) - X_1(t) =$$

$$= (x_2 - x_1) \exp \left\{ \int_0^t \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s) - \frac{1}{2} \int_0^t (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds \right\}.$$

On the set $\{\tau < +\infty\}$

$$\lim_{t\to\infty}\int_0^{t\wedge\tau} \left(\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)}\right)^2 ds$$

exists and is finite, because

$$\left(\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)}\right)^2 \leq L^2$$
 where L is the Lipschitz constant for σ .

Therefore (Lemma 2.1)

$$\lim_{t\to+\infty}\int_0^{t\wedge\tau}\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)}dW(s)$$

exists and is finite. This is of course a contradiction because

$$\int_0^t \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s) - \frac{1}{2} \int_0^t (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds$$

should tend to $-\infty$ since $X_2(t)-X_1(t)\longrightarrow 0$ as $t\to \tau$ on the set $\{\tau<+\infty\}$.

The main result in this Section is the following theorem:

Theorem 5.1 Under assumptions I and II, $X_2(t) - X_1(t) \to 0$ if and only if $\int_{-\infty}^{+\infty} [\sigma'(u)]^2 du = +\infty$.

Proof of Theorem 5.1

Because of Lemma 5.1 and since $X_2(t) - X_1(t)$ is a positive local martingale we have ([6], see also [3]):

$$X_2(t) - X_1(t)$$

$$=(x_2-x_1)\exp\big\{\int_0^t\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)}dW(s)-\frac{1}{2}\int_0^t(\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)})^2ds\big\}$$

From assumption I and Lemma 5.1 we conclude that $X_2(t) - X_1(t) > 0$ a.s.

for all t. Therefore $X_2(t) - X_1(t) \to 0$ a.s. if and only if

$$\int_0^t \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s) - \frac{1}{2} \int_0^t (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds \ a.s.$$

tends to $-\infty$, which will happen if and only if

$$\int_0^t (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds = +\infty.$$

Indeed, on the set $\{\omega: \int_0^{+\infty} \left(\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)}\right)^2 ds < +\infty\}$

$$\lim_{t\to+\infty}\int_0^t\frac{\sigma(X_2(s))-\sigma(X_1(s))}{X_2(s)-X_1(s)}dW(s)$$

exists and is finite (Lemma 2.1), therefore on the set

$$\{\omega: \int_0^{+\infty} (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds < +\infty\} \ X_2(t) - X_1(t) \text{ will not tend to } 0.$$

If

$$\int_0^{+\infty} \left(\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)}\right)^2 ds = +\infty \ a.s.,$$

then let us define $\tau_t = \inf\{u: \int_0^u \left(\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)}\right)^2 ds > t\}$.

Then from Theorem 2.2 we conclude that $M(t) = \int_0^{\tau_t} \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s)$ is a new Brownian motion and therefore

$$M(t) - \frac{1}{2} \int_0^{\tau_t} (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds = M(t) - \frac{1}{2}t \to -\infty \quad a.s.$$

Since $\lim_{t\to +\infty} (X_2(t) - X_1(t))$ exists, then

$$\lim_{t\to+\infty} \left\{ \int_0^t \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s) - \frac{1}{2} \int_0^t \left(\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} \right)^2 ds \right\}$$

exists too. Since

$$\int_0^{\tau_t} \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s) - \frac{1}{2}t \to -\infty$$

therefore almost surely

$$\int_0^t \frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)} dW(s) - \frac{1}{2} \int_0^t (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds \to -\infty$$

Therefore $X_2(t) - X_1(t) \to 0$ a.s. if and only if

$$\int_0^t (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds \to +\infty$$

Now we need to show that under assumptions I and II

$$\int_0^t \left(\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)}\right)^2 ds = \begin{cases} +\infty & \text{if } \int^{+\infty} [\sigma'(u)]^2 du = +\infty \\ < +\infty & \text{if } \int^{+\infty} [\sigma'(u)]^2 du < +\infty \end{cases}$$

Let us first assume that σ is bounded. Then from the assumptions I and II it follows that $|\sigma'(x)|$ is decreasing for x large enough because otherwise σ would grow faster than a linear function, and then Q(x) would tend to $+\infty$ as $t \to +\infty$. It is known ([2]) that $\frac{X(t)}{t} \to c$ a.s. So for $t > T_{\epsilon}(\omega)$ we have $(c - \epsilon)t \le X_1(t) \le X_2(t) \le (c + \epsilon)t$.

From the assumption II it also follows that there is $T_{a_1}(\omega)$ such that for all $t > T_{a_1}(\omega)$ we have $X_2(t) > a_1$ and $X_1(t) > a_1$. Therefore we have from the assumption I and from the mean value theorem:

$$\int_{0}^{T_{a_{1}} \vee T_{\epsilon}} (\frac{\sigma(X_{2}(s)) - \sigma(X_{1}(s))}{X_{2}(s) - X_{1}(s)})^{2} ds + \int_{T_{a_{1}} \vee T_{\epsilon}}^{+\infty} [\sigma'((c + \epsilon)t)]^{2} dt$$

$$\leq \int_{0}^{+\infty} (\frac{\sigma(X_{2}(s)) - \sigma(X_{1}(s))}{X_{2}(s) - X_{1}(s)})^{2} ds$$

$$\leq \int_{0}^{T_{a} \vee T_{\epsilon}} (\frac{\sigma(X_{2}(s)) - \sigma(X_{1}(s))}{X_{2}(s) - X_{1}(s)})^{2} ds + \int_{T_{a} \vee T_{\epsilon}}^{+\infty} [\sigma'((c - \epsilon)t)]^{2} dt$$

Therefore if $\int_{0}^{+\infty} [\sigma'(x)]^2 dx = +\infty$, then $\int_{0}^{+\infty} (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds = +\infty$ a.s. and if $\int_{0}^{+\infty} [\sigma'(u)]^2 du < +\infty$, then $\int_{0}^{+\infty} (\frac{\sigma(X_2(s)) - \sigma(X_1(s))}{X_2(s) - X_1(s)})^2 ds < +\infty$ a.s. The theorem is proved for bounded σ . It is easy to see that if $\sigma(x)$ is bounded, then under the assumptions I and II $\int_{0}^{+\infty} [\sigma'(x)]^2 dx < +\infty$, so in this case ξ cannot be zero.

Now assume that σ is unbounded. It follows from the assumptions I and II that there exists x_0 such that for $x > x_0$, $\sigma'(x) > 0$ and $\sigma'(x)$ is decreasing. Indeed. Since $\sigma(x)$ is convex or concave for large x's, then $\sigma(x)$ is decreasing or increasing for large x's. Since it is unbounded it must be increasing. Next, if it is convex, then it grows at least as fast as a linear function and then $Q(x) \to +\infty$, which is contradiction with assumption II. Therefore $\sigma'(x) \geq 0$ and σ' is decreasing for x large enough, so σ is concave.

Let $T_{x_0} = \inf\{t : X_2(t) \ge x_0, X_1(t) \ge x_0\}$. Since

$$\int_{0}^{T_{x_{0}}} \left(\frac{\sigma(X_{2}(s)) - \sigma(X_{1}(s))}{X_{2}(s) - X_{1}(s)}\right)^{2} ds + \int_{T_{x_{0}}}^{+\infty} \left[\sigma'(X_{2}(s))\right]^{2} ds$$

$$\leq \int_{0}^{+\infty} \left(\frac{\sigma(X_{2}(s)) - \sigma(X_{1}(s))}{X_{2}(s) - X_{1}(s)}\right)^{2} ds$$

$$\leq \int_{0}^{T_{x_{0}}} \left(\frac{\sigma(X_{2}(s)) - \sigma(X_{1}(s))}{X_{2}(s) - X_{1}(s)}\right)^{2} ds + \int_{T_{x_{0}}}^{+\infty} \left[\sigma'(X_{1}(s))\right]^{2} ds$$

then we need to investigate the convergence of $\int_0^{+\infty} [\sigma'(X(s))]^2 ds$ where X(s) is a solution of (7) starting at x > -1. Let Y(t) = Q(X(t)), so that Y(t) is a process on the natural scale and (see Section 1) $dY(t) = \sigma_1(Y(t))dW(t)$ where $\sigma_1(y) = \sigma(Q^{-1}(y))Q'(Q^{-1}(y))$ and $Y(0) = y_0 = Q(x)$.

From Theorem 2.3 we conclude that there exists a Brownian Motion B(t) starting at

 y_0 on some natural extension (see Definition 2.7) of our basic probability space, such that $Y(t) = B(\langle Y \rangle(t))$. Let $M(t) = Y(t) - y_0$. Let $\gamma_t = \langle M \rangle(t)$. It is known, that

$$\gamma_t = \langle M \rangle(t) = \int_0^t \sigma_1^2(Y(s)) ds$$

Let A_t be the inverse function of γ_t . Then we have :

$$t = \int_0^t \frac{1}{\sigma_1^2(Y(s))} d\gamma_s$$

So

$$A_{t} = \int_{0}^{A_{t}} \frac{1}{\sigma_{1}^{2}(Y(s))} d\gamma_{s} = \int_{0}^{t} \frac{1}{\sigma_{1}^{2}(Y(A_{u}))} du$$

but from Theorem 2.3 we see, that $Y(A_u) = B(u)$ and therefore

$$A_t = \int_0^t \frac{1}{\sigma_1^2(B(u))} du$$

So we have shown, that Y(t) can be represented as $B(\gamma_t)$ where γ_t is the inverse function to

 $A_t = \int_0^t \frac{1}{\sigma_1^2(B(u))} du$ and B(s) is a Brownian motion starting at $y_0 = Q(x)$.

Let
$$\tau_d = \inf\{s : B(s) = d\}$$

We have:

$$\int_0^{+\infty} [\sigma'(X(s))]^2 ds = \int_0^{+\infty} [\sigma'(Q^{-1}(B(\gamma_s)))]^2 ds = \int_0^{\tau_d} \frac{[\sigma'(Q^{-1}(B(u)))]^2}{\sigma_1^2(B(u))} du$$

$$= \int_0^{\tau_d} \frac{[\sigma'(Q^{-1}(B(u)))]^2}{\sigma^2(Q^{-1}(B(u)))[Q'(Q^{-1}(B(u)))]^2} du.$$

Let

$$F(x) = \frac{[\sigma'(Q^{-1}(x))]^2}{\sigma^2(Q^{-1}(x))[Q'(Q^{-1}(x))]^2}$$

We will need the following lemma which follows from ([1]):

Lemma 5.2 Let d > 0 and let $F : (-\infty, d) \longrightarrow R$ be nonnegative and continuous.

Let $\tau_d = \inf\{t : B(t) = d\}$, where B(s) is a Brownian motion starting at x < d.

Then

 $\int_0^{\tau_d} F(B(u))du < +\infty$ a.s. if and only if $\int_0^d F(y)(d-y)dy < +\infty$ and

if
$$\int_0^d F(y)(d-y)dy = +\infty$$
, then $\int_0^{\tau_d} F(B(u))du = +\infty$ a.s.

Let us show the proof of this lemma:

Proof. Let B(u) be a Brownian motion starting at 0. It is known that :

$$\int_0^{\tau_{d-s}} F(x+B(u)) du = \int_{-\infty}^{d-x} F(x+y) l_{\tau_{d-s}}^{y} dy$$

where l_t^y is the local time for Brownian motion.

We have:

$$\int_{-\infty}^{d-x} F(x+y) l_{\tau_{d-s}}^{y} dy = \int_{-\infty}^{-x} F(x+y) l_{\tau_{d-s}}^{y} dy + \int_{-x}^{d-x} F(x+y) l_{\tau_{d-s}}^{y} dy =$$

$$= \int_{-\infty}^{0} F(z) l_{\tau_{d-s}}^{z-x} dz + \int_{0}^{d} F(z) l_{\tau_{d-s}}^{z-x} dz$$

Since $l_{\tau_{d-x}}^{z-x}$ is continuous and almost surely

$$\inf_{u \le \tau_{d-s}} B(u) > -\infty,$$

then $\int_{-\infty}^{0} F(z) l_{\tau_{d-x}}^{z-x} dz < +\infty$ a.s., so we have to investigate the convergence of $\int_{0}^{d} F(z) l_{\tau_{d-x}}^{z-x} dz$. It is known that ([11])

$$l_{\tau_{d-x}}^{z-x} = \frac{1}{2}(W_1^2(d-x-z+x) + W_2^2(d-x-z+x)) = \frac{1}{2}(W_1^2(d-z) + W_2^2(d-z))$$

where W_1 and W_2 are two independent Brownian motions starting at 0. Therefore our problem is to determine for which functions F

$$\int_0^d F(z)W(d-z)^2 dz = \int_0^d F(d-t)W^2(t)dt$$

is finite a.s., where W is a Brownian motion. First it follows from Bluhmentahl's 0-1 law that

$$P(\int_0^d F(d-t)W^2(t)dt < +\infty) = 0 \text{ or } 1$$

Now, assume that $\int_0^d F(d-t)tdt < +\infty$. Then

$$E\int_{\epsilon}^{d}F(d-t)W^{2}(t)dt \to \int_{0}^{d}F(d-t)tdt < +\infty$$

so $E \int_0^d F(d-t)W^2(t)dt < +\infty$ which implies

$$P(\int_0^d F(d-t)W^2(t)dt < +\infty) = 1$$

Assume now that

$$P(\int_0^d F(d-t)W^2(t)dt < +\infty) = 1$$

Let $H = \{g : [0,1] \to R : \int_0^d F(d-t)g^2(t)dt < +\infty\}$. Then H is a Hilbert space with the scalar innerproduct given by $: (g,h) = \int_0^d F(d-t)g(t)h(t)dt$ If $\int_0^d F(d-t)W^2(t)dt < +\infty$, then W(t) defines a gaussian random element W with values in H, and therefore

$$E \mid\mid W \mid\mid^2 = E \int_0^d F(d-t)W^2(t)dt = \int_0^d F(d-t)tdt < +\infty$$

Which completes the proof of the lemma.

From this lemma it follows that we need to check whether the integral

$$\int_0^d \frac{[\sigma'(Q^{-1}(x))]^2}{\sigma^2(Q^{-1}(x))[Q'(Q^{-1}(x))]^2} (d-x) dx$$

is finite.

We have:

$$\int_{0}^{d} \frac{[\sigma'(Q^{-1}(x))]^{2}}{\sigma^{2}(Q^{-1}(x))[Q'(Q^{-1}(x))]^{2}} (d-x) dx = \int_{0}^{+\infty} \frac{[\sigma'(y)]^{2}}{\sigma^{2}(y)[Q'(y)]} (d-Q(y)) dy$$

$$= \int_{0}^{+\infty} \frac{[\sigma'(y)]^{2}}{\sigma^{2}(y)} e^{2c \int_{0}^{y} \frac{1}{\sigma^{2}(u)} du} \int_{y}^{+\infty} e^{-2c \int_{0}^{x} \frac{1}{\sigma^{2}(u)} du} dz dy$$

Assume first that $\int_0^{+\infty} (\sigma'(y))^2 dy < +\infty$. Then $\sigma(y)\sigma'(y) \to 0$, as $y \to +\infty$. Indeed, first we show that

$$\liminf_{y\to+\infty}\sigma(y)\sigma'(y)=0.$$

Suppose that it is not the case. Then there is $\delta > 0$ and y_0 such that for all $y > y_0$, $\sigma(y)\sigma'(y) \ge \delta$. Then $\sigma'(y)^2 \ge \frac{\delta}{\sigma^2(y)}$ for $y \ge y_0$. But $\int_{y_0}^{+\infty} \frac{\delta}{\sigma^2(y)} = +\infty \text{ because of assumption II, so } \int_{y_0}^{+\infty} (\sigma'(y))^2 dy = +\infty \text{ which is a contradiction to } \int_0^{+\infty} (\sigma'(y))^2 dy < +\infty. \text{ Hence}$

$$\liminf_{y\to+\infty}\sigma(y)\sigma'(y)=0.$$

Next we have : for every $\epsilon > 0$ there is y_0 such that for all $z \ge y_0$ and $y \ge y_0$, $\int_y^z (\sigma'(u))^2 du \le \epsilon$. Then since σ' is decreasing, we have :

$$\epsilon \geq \int_{u}^{z} (\sigma'(u))^{2} du = \sigma(z)\sigma'(z) - \sigma(y)\sigma'(y) - \int_{u}^{z} \sigma(u) d\sigma'(u) \geq \sigma(z)\sigma'(z) - \sigma(y)\sigma'(y)$$

Choose y such that $\sigma(y)\sigma'(y) \leq \epsilon$. Then for all $z \geq y$ we have $2\epsilon \geq \sigma(z)\sigma'(z)$ so $\sigma(z)\sigma'(z) \to 0$ as $z \to +\infty$. It follows now from d'Hospital's rule that

$$\lim_{y\to+\infty}\frac{1}{\sigma^2(y)}e^{2c\int_0^y\frac{1}{\sigma^2(u)}du}\int_y^{+\infty}e^{-2c\int_0^z\frac{1}{\sigma^2(u)}du}dz=$$

$$= \lim_{y \to +\infty} \frac{-e^{-2c \int_0^y \frac{1}{\sigma^2(u)} du}}{2\sigma(y)\sigma'(y)e^{-2c \int_0^y \frac{1}{\sigma^2(u)} du} - \sigma^2(y) \frac{2c}{\sigma^2(y)} e^{-2c \int_0^y \frac{1}{\sigma^2(u)} du}} = \lim_{y \to +\infty} \frac{1}{2c - 2\sigma(y)\sigma'(y)} = \frac{1}{2c}.$$

Therefore if $\int_0^{+\infty} (\sigma'(y))^2 dy < +\infty$, then

$$\int_0^{+\infty} \frac{[\sigma'(y)]^2}{\sigma^2(y)} e^{2c \int_0^y \frac{1}{\sigma^2(u)} du} \int_y^{+\infty} e^{-2c \int_0^x \frac{1}{\sigma^2(u)} du} dz dy < +\infty.$$

Assume now that $\int_0^{+\infty} (\sigma'(y))^2 dy = +\infty$. It is easy to see that :

$$\frac{d}{dy} \left[e^{2c \int_0^y \frac{1}{\sigma^2(u)} du} \int_y^{+\infty} e^{-2c \int_0^z \frac{1}{\sigma^2(u)} du} dz \right] = \frac{2c}{\sigma^2(y)} e^{2c \int_0^y \frac{1}{\sigma^2(u)} du} \int_y^{+\infty} e^{-2c \int_0^z \frac{1}{\sigma^2(u)} du} dz - 1$$

Therefore we have by integration by parts:

$$2c \int_{0}^{a} \frac{(\sigma'(y))^{2}}{\sigma^{2}(y)} e^{2c \int_{0}^{y} \frac{1}{\sigma^{2}(u)} du} \int_{y}^{+\infty} e^{-2c \int_{0}^{z} \frac{1}{\sigma^{2}(u)} du} dz dy$$

$$= \int_{0}^{a} (\sigma'(y))^{2} \left[\frac{2c}{\sigma^{2}(y)} e^{2c \int_{0}^{y} \frac{1}{\sigma^{2}(u)} du} \int_{y}^{+\infty} e^{-2c \int_{0}^{z} \frac{1}{\sigma^{2}(u)} du} dz dy$$

$$= \int_{0}^{a} (\sigma'(y))^{2} \left[\frac{2c}{\sigma^{2}(y)} e^{2c \int_{0}^{y} \frac{1}{\sigma^{2}(u)} du} \int_{y}^{+\infty} e^{-2c \int_{0}^{z} \frac{1}{\sigma^{2}(u)} du} dz - 1 \right] dy + \int_{0}^{a} (\sigma'(y))^{2} dy$$

$$= (\sigma'(a))^{2} \left[e^{2c \int_{0}^{a} \frac{1}{\sigma^{2}(u)} du} \int_{a}^{+\infty} e^{-2c \int_{0}^{z} \frac{1}{\sigma^{2}(u)} du} dz \right] - (\sigma'(0))^{2} d$$

$$-2 \int_{0}^{a} \sigma'(y) e^{2c \int_{0}^{y} \frac{1}{\sigma^{2}(u)} du} \int_{y}^{+\infty} e^{-2c \int_{0}^{z} \frac{1}{\sigma^{2}(u)} du} dz d\sigma'(y) + \int_{0}^{a} (\sigma'(y))^{2} dy$$

$$\geq (\sigma'(a))^{2} e^{2c \int_{0}^{a} \frac{1}{\sigma^{2}(u)} du} \int_{a}^{+\infty} e^{-2c \int_{0}^{z} \frac{1}{\sigma^{2}(u)} du} dz - (\sigma'(0))^{2} d + \int_{0}^{a} (\sigma'(y))^{2} dy$$

$$\geq \int_{0}^{a} (\sigma'(y))^{2} dy - (\sigma'(0))^{2} d \longrightarrow +\infty \quad as \quad y \to +\infty$$

Therefore if $\int_0^{+\infty} (\sigma'(y))^2 dy = +\infty$, then

$$\int_0^{+\infty} \frac{[\sigma'(y)]^2}{\sigma^2(y)} e^{2c \int_0^y \frac{1}{\sigma^2(u)} du} \int_y^{+\infty} e^{-2c \int_0^z \frac{1}{\sigma^2(u)} du} dz dy = +\infty$$

which completes the proof of the Theorem.

The following example shows that, two solutions of (7) can hit each other with positive probability in case σ is not Lipschitz continuous.

Example 5.1 Let W(s) denote the Brownian Motion starting at 0. Let

$$B(t) = \int_0^t sgn(W(s))dW(s)$$

Consider the following stochastic differential equation:

(8)
$$dX(t) = dt + 2\sqrt{X(t)}dB(t) \quad X(0) = x$$

Let $\eta = \inf\{t : X^x(t) = 0\}$. Theorem 2.7 implies that $P_x(\eta < +\infty) = 1$ for all $x \geq 0$. Consider two solutions $X_1(t)$ and $X_2(t)$ starting at x_1 and x_2 respectively $(x_2 > x_1)$. Comparison theorem (Theorem 2.5) implies that $X_2(t) \geq X_1(t)$ for all t a.s. Since $P_{x_2}(\eta < +\infty) = 1$ and $X_1(t) \geq 0$ a.s., then $\tau = \inf\{t : X_2(t) = X_1(t)\} < +\infty$ a.s.

In fact we can solve the equation (8) (see [7], exercise 5.35). The solution of (8) with initial condition X(0) = x is given by:

$$X(t) = \begin{cases} (\sqrt{x} + B(t))^2 & \text{if } t < \tau_{\sqrt{x}} \\ W^2(t - \tau_{\sqrt{x}}) & \text{if } t \ge \tau_{\sqrt{x}} \end{cases}$$

where $\tau_x = \inf\{t : B(t) = -x\}.$

6 SUMMARY

In this dissertation we present some results concerning the stability and asymptotic stability of the solutions of stochastic differential equations. These are in most cases extensions of work of Khasminskii and Nevelson ([8]).

Unlike the case of equilibrium points, the stability properties of arbitrary solutions of stochastic differential equations are not thoroughly studied. For the case of equilibrium points there is a well developed theory created by Khasminskii and others (see [9]).

In Section 3 we analyzed the behavior of solutions on a finite interval. Most facts presented there are well known, and we treated them as an introduction to sections 4 and 5.

In Section 4 we investigated the stability properties of the solutions of stochastic differential equations where we allow $\sigma = 0$, which was not allowed by Khasminskii and Nevelson (see [8]). We showed that some classification of the points of the real line is necessary, and we established stability results for various cases. It should be noted, that most of these results are formulated in terms of convergence in the Euclidean metric, which is stronger than the scale metric considered by Khasminskii and Nevelson.

In Section 5 we treated the case when the drift coefficient b is constant. In this case the recurrence property of solutions from [8] may be violated, but the limit of the difference of two solutions still exists. Under some additional assumptions, we showed that differences converge to zero almost surely. We believe that these additional assumptions may be relaxed in the future, and a similar analysis may be carried out without them.

There is a variety of additional problems connected with the discussion of Section 5. One interesting problem is to give conditions on the drift and diffusion coefficients b and σ , under which two solutions starting from two different points never hit each other (with probability one). More specifically consider the stochastic differential equation:

$$dX(t) = b(X(t))dt + \sigma(X(t))dW(t)$$

Let $X_1(t)$ and $X_2(t)$ denote two solutions starting from two different points. The problem is to give conditions on b and σ under which $\tau = inf\{t : X_1(t) = X_2(t)\}$ is almost surely finite. Another problem is to

 $\tau = inf\{t: X_1(t) = X_2(t)\}$ is almost surely finite. Another problem is to determine when $X_2(t) - X_1(t) \to 0$ (a.s.) when the drift coefficient is decreasing and non-constant, without further restrictions on σ .

The discussion of this dissertation does not cover the case when the coefficients of the stochastic differential equation are time-dependent. We point out, that in this case questions similar to those considered in this dissertation may be posed. The main difficulty in analyzing those problems is to find an analogue of the scale function Q. When the coefficients of stochastic differential equation do not depend on time, then the scale function "removes" the drift, so the martingale theory can be applied. Therefore, there is a hope that once an analogue of the scale function for the time dependent case is introduced, then similar results to these obtained in this dissertation can be established.

The methods developed and used in this dissertation do not apply directly to higher dimensional systems since they rely on linear ordering of R^1 . Other methods are needed to study higher dimensional problems.

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