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EFFECTS OF ACCOUNTING TECHNIQUES ON THE STUDY OF MARKET POWER

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by

Kiran Verma

A DISSERTATION

Submitted to Michigan State University in partial fulfillment of the requirements for the degree of

DOCTOR OF PHILOSPHY

Department of Accounting

ABSTRACT

EFFECTS OF ACCOUNTING TECHNIQUES ON THE STUDY OF MARKET POWER

by

Kiran Verma

The purpose of this study is to examine the effects of alternative accounting techniques on inferences drawn from the relationship between firm profitability and its market power. Accounting data have been widely used as sample evidence in such analyses. As accounting rules allow firms considerable leeway in their chioce of accounting techniques, it has been alleged that inferences drawn from such analyses are potentially misleading. This is particularly true for time series analyses at the firm level where interfirm differences due to accounting techniques are not aggregated out.

This study is conducted in two parts. The first part deals with the development of a model linking a firm's accounting profit to its market share, capital stock, and raw materials used in production. The second part examines the robustness of this model to estimation using accounting data that are generated by alternative accounting techniques (e.g. LIFO vs. FIFO, Straight line vs. Double Declining Balance, and Historical cost vs. Replacement cost).

Unlike model developed in previous research, the model in this study specifically incorporates the presence of inventories. This is done in order to describe more closely the economic phenomenon of measuring the accounting cost of goods sold and hence the accounting profit.

Robustness of the model developed above is tested with respect to two parameters, (1) its explanatory power, and (2) inferences drawn from it about the firm's market power. The results show that both of these parameters are sensitive to the choice of certain accounting techniques. This effect is most pronounced for the change in inventory techniques and almost nonexistant for the change in depreciation techniques. The results are mixed for the change from Historical cost to Replacement cost. To Sumi, Charu, and Priya

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				Page
L	ST OF	TABLE	S	V
I.	INTRO	DUCTIO	N	1
II.	BACKGROUND		3	
	2.1 2.2 2.3	Effect Purpos Profit	s of Alternative Accounting Techniques e of the Dissertation ability, and Market Structure	3 6 7
III.	THE MODEL		11	
	3.1 3.2 3.3	A Simp Factor Effect	le Model of a Profit Maximizing Firm s of Production and Accounting Values s of Accounting Techniques on the	11 13
		Coeffi	cients of Interest	28
IV.	HYPOI	HESES		23
۷.	METHO	DOLOGY		27
	5.1 5.2 5.3 5.4 5.5 5.6	Retate Invent Restat Restat Data Method	ment procedures ory Restatement ement for Depreciation ement for Replacement Cost s of Analysis	27 28 30 31 28 29
		5.6æ 5.6b 5.6c	Analysis of Variance Kendall's Coefficient of Concordance Spearman's Rank Correlation Coefficien	35 39 t 41
VI	EMPIJ	RICAL R	ESULTS AND ANALYSIS	43
	6.1 6.2 6.3 6.4	Genera Parame Analys Adjust 6.4b 6.5b	l Results tric t test is of Variance ed R ² Spearman's Rank Correlation Coefficien	43 43 45 46 46
	6.5	Market	power test criterion	47
		6.5 a 6.5b	Kendall's Coefficient of Concordance Spearman's Rank Correlation Coefficien	48 t 48

Page

VII	Conclusions	97
VIII	APPENDIX	101
IX	FOOTNOTES	109
x	BIBLIOGRAPHY	113

•

LIST OF TABLES

Table		Page
1.	Combinations of Accounting Techniques.	28
2.	Pairvise comparison of alternative accounting techniques.	22
3.	Firms, Industry SIC codes, and reported accounting methods.	34
4.	Results of the time series OLS regression with 36 observations for the Armco Inc.	49
5.	Results of the time series OLS regression with 36 obsevations for the USX Corporation.	58
6.	Results of the time series OLS regression with 36 obsevations for the Lukens Corporation.	51
7.	Results of the time series OLS regression with 36 obsevations for the LTV Corporation.	52
8.	Results of the time series OLS regression with 36 observations for the General Motor Corporation.	53
9.	Results of the time series OLS regression with 36 observations for the Ford Motor Company.	54
18.	Results of the time series OLS regression with 36 observations for the American Motors Corporation.	55
11.	Results of the time series OLS regression with 36 observations for the Chrysler Corporation.	56
12.	Adjusted R ² from estimation of equation (15) using combinations of accounting techniques.	57
13.	Test Criterion, a _g +a _l MS from estimation of equation (15).	57
14.	Significance (Z-statistic) of individual coefficients.	58
15.	't' test for difference in means of individual coefficients.	59
16.	Adjusted R^2 , analysis of variance for the 'Iron and Steel' industry.	60
17.	Adjusted R ² , analysis of variance for the 'Automobile' industry.	60

18.

- 19. Adjusted R², firm rankings, and Kendall's 61 coefficient of concordance for the combined sample of the 'Iron and Steel and 'Automobile' industries.
- 20. For the comparison LIFO vs.FIFO:SLHC, Adjusted R², 62 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile' industries.
- 21. For the comparison LIFO vs.FIFO:DDBHC, Adjusted R², 62 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile' industries.
- 22. For the comparison SL vs. DDB:LIFOHC, Adjusted R², 63 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile'industries.
- 23. For the comparison SL vs. DDB:FIFOHC, Adjusted R², 63 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile' industries.
- 24. For the comparison RC vs. HC:LIFOSL, Adjusted R², 64 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile' industries.
- 25. For the comparison RC vs. HC:FIFOSL, Adjusted R², 64 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile' industries.
- 26. For the comparison RC vs. HCO:LIFODDB, Adjusted R², 65 firm rankings, and Spearman's rank correlation coefficient for the 'Iron and Steel' and 'Automobile' industries.
- 27. For the comparison RC vs. HC [FIFODDB, Adjusted R², 65 firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

28.	Adjusted R ² , Spearman's rank correlation coefficient for the combined sample of the 'Iron and Steel' and 'Automobile' industries.	66
29.	Adjusted R ² , firm rankings, and Kendall's coefficient of concordance for the Iron and Steel industry.	67
30.	For the comparison LIFO vs. FIFO:SLHC, Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry	68
31.	For the comparison LIFO vs. FIFO:DDBHC, Adjusted R [±] , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	68
32.	For the comparison SL vs. DDB:LIFOHC,Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	69
33.	For the comparisonSL vs. DDB:FIFOHC, Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	69
34.	For the comparison RC vs. HCLIFOSL, Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	70
35.	For the comparison RC vs. HCFIFOSL, adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	70
36.	For the comparison RC vs. HCLIFODDB, Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	71
37.	For the comparison RC vs. HCFIFODDB, Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.	71
38.	Adjusted R ² , Spearman's rank correlation coefficient for the Automobile industry.	72
39.	Adjusted R ² , firm rankings, and kendall's coefficient of concordance for the Automobile industry.	73
40.	For the comparison LIFO vs. FIFO:SLHC, Adjusted R ² , firm rankings, and Spearman's rank correlation coefficient for the Automobile industry.	74

- 63. For the comparison SL vs. DDB:FIFOHC, Test 87 criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.
- 64. For the comparison RC vs. HCLIFOSL, Test criterion, 88 firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.
- 65. For the comparison RC vs. HCFIFOSL, Test criterion, 88 firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.
- 66. For the comparison RC vs. HCLIFODDB, Test criterion, 89 firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.
- 67. For the comparison RC vs. HCFIFODDB, Test criterion, 89 firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.
- 68. Market power test criterion, firm rankings 98 and Spearman's rank correlation coefficient for the Iron and Steel industry.
- 69. Market power Test criterion, firm rankings 91 and Kendall's coefficient of concordancefor the Automobile industry.
- 70. For the comparison LIFO vs.FIFO: SLHC, Test 92 criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 71. For the comparison LIFO vs. FIFO: DDBHC, Test 92 criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 72. For the comparison SL vs. DDB: LIFOHC, Test 93 criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 73. For the comparison SL vs. DDB:FIFOHC, Test 93 criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 74. For the comparison RC vs. HCLIFOSL, Test criterion, 94 firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

- 75. For the comparison RC vs. HCFIFOSL, Test criterion, 94 firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 76. For the comparison RC vs. HCLIFODDB, Test criterion, 95 firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 77. For the comparison RC vs. HCFIFODDB, Test criterion, 95 firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
- 78. Market Power Test criterion, firm rankings 96 and Spearman's rank correlation coefficient for the Automobile industry.

I INTRODUCTION

Establishing methods for assessing the desirability and effects of alternative accounting techniques has long been a goal of accounting researchers and accounting policy makers. Despite an abundance of empirical and theoretical research in the area, several of the issues are still unresolved. The purpose of this study is to examine the effects of accounting techniques in the context of a specific economic analysis and with respect to a specific class of users. The economic analysis involves the study of market performance in the area of industrial organization by researchers and antitrust policy makers.

Two important reasons for analyzing the impact of different accounting methods on the relationship between profitability and market structure are the following. First, as mentioned before, the issue of accounting techniques is of considerable importance for accounting policy makers. Second, the study of determinants of market structure has relevance not only for economists, but also for regulators, legislators, and lawyers dealing with antitrust litigation (Adams [1979]). In this context, accounting rates of return are frequently used as a measure of profitability and treated as an index of monopoly power.¹

Accounting numbers are among the most frequently used data in empirical work on industrial organization. The systematic properties of these numbers are used as a basis for making inferences about market structure and market performance factors (such as the presence of monopoly power and above normal profits). Thus, potential effects of accounting techniques on such inferences are an important factor to be considered by researchers in this area. According to Gonedes and Dopuch [1979], this role of accounting numbers and the large cost of making avoidable, unvarranted inferences are sufficient grounds for being interested in the effects of accounting techniques.

1I BACKGROUND

2.1 Effects of Alternative Accounting Techniques

Research examining the effects of alternative accounting techniques has taken several approaches. Prior to the mid 1960's most attempts to deal with the question used 'a priori' models which purported to establish criteria for identifying the 'best' or 'optimal' accounting procedures (e.g. Paton and Littleton [1940], Chambers [1966], and Edwards and Bell [1965]). Subsequently the 'a priori' approach came under heavy criticism because (1) it lacked theoretical background, and (2) using this approach one could declare just about any set of accounting procedures 'superior'. These failings have led to an increased emphasis on the use of empirical analyses to assess the effects and desirability of alternative accounting techniques.

Other approaches taken by researchers in this area include the 'materiality' studies and the studies based on the connection between market equilibrium and accounting techniques. The 'materiality' research examined whether the process generating an accounting number according to one accounting technique is significantly different (say by a percent rule) from the process generating the same number according to a different technique. This test of 'significance' usually involved point by point comparison

of observed values conditional on alternative accounting techniques (see Rosenfeld [1969], and Peterson [1973]). These studies, however, ignored the time series issues involved, and as such provided no basis for inferring that the two stochastic processes corresponding to alternative accounting techniques were different just because contemporaneous drawings from each process were different.

In other studies (e.g. Dopuch and Watts [1972], Dopuch and Drake [1966], and Dopuch and Pincus [1984]) it has been argued that significance (materiality) could be determined by assessing the effect of a change in accounting technique on parameters of the time series process generating the accounting number of interest. These studies do not answer the question of whether predictive ability, or variability, is a sufficient basis for assessing the effects or the desirability of accounting techniques. Similar deficiencies characterize the simulation studies (examples of the simulation studies are Simmons and Gray [1969], and Greenball [1968 and 1969]).

On the other hand, market equilibrium studies are based on explicit theoretical models of resource allocation under uncertainty (see Ball and Foster [1982] for a review of such studies). These studies suffer from two common problems. (1) When do capital market agents learn about a firm's choice of accounting methods (date of

event), and (2) how does one control for confounding information which is generally released along with the "main event"? These models provide some underpinnings for empirical work on techniques (see Gonedes and Dopuch [1974] for a discussion), solid underpinnings are also available in other areas, such as the area of industrial organization.

The approach used by a number of researchers working in the area of industrial organization can be categorized as the "True" or "Ideal" number approach (see e.g. Ayanian [1975], Bloch [1976], Clarkson [1977], and Weiss [1969]). These studies are based on the premise that accounting techniques lead to 'biased' income numbers because they are not computed according to procedures asserted to provide values of 'Ideal' or 'True' income.

In some of these studies it was found that inferences based on restated values differed from those based on reported values. These studies, nonethless, did not specifically assess the potential effects of accounting techniques on inferences that were based on values of accounting numbers. More importantly, in several of these studies it was asserted that the restatement techniques used were somehow superior to conventional techniques in that they led to 'more informative' numbers. The restatement techniques used in these studies, however, are just as arbitrary as conventional techniques since the

models they are based on are not descriptive of a world with uncertainty, imperfect markets, and multiperiod horizons.

2.2 Purpose of the dissertation

There are two basic problems of using accounting based variables in economic analyses which can be conceptualized as follows (see Nair [1977]),

A ← B

^B1 B₂

here A is the unobservable economic variable such as economic profit, B is the accounting surrogate such as accounting income, and B_1 and B_2 are accounting income numbers computed under alternative accounting techniques such as LIFO vs. FIFO. The first problem deals with whether the accounting variable is a good surrogate for the economic variable, i.e., whether the same modeled relationship holds for the economic and the accounting variable. The second problem is that even after picking the surrogate accounting variable, the problems continue because of the several different ways of producing the accounting number. Most of the previous studies have only examined the second problem.² In the proposed study, both of these problems are dealt with. The first problem is overcome by developing a structural model which specifically links accounting numbers with parameters of market structure such as the elasticity of demand, product differentiation, and market share. This model also attempts to deal with the issue of multiperiod horizons, which according to Gonedes and Dopuch [1979] (footnote 18) is one of the basic concerns.

Based on this model, the second problem is examined by forming and testing hypotheses about the potential effects of alternative inventory (LIFO vs. FIFO), depreciation (st. line (SL) vs. double declining balance (DDB)), and asset valuation (historical cost (HC) vs. replacement cost (RC)) techniques on inferences about market power. Instead of presuming that the same relationship exists between independent and dependent variables when estimated using accounting instead of economic values, the present study will model the relationship between accounting values and the parameters of interest directly, and based on it will hypothesize the effects of alternative accounting techniques.

2.3 Profitability and Market Structure

Collusion in oligopolistic industries that is potentially damaging to consumers has not only been the cause of legal restrictions and increased regulation, but

has also given rise to an enormous amount of research attempting to relate market concentration to the exercise of monopoly power.³ Nost such studies have measured market structure by an index of concentration⁴ (e.g. the four firm concentration ratio as a surrogate for market power), and performance by accounting profit (such as gross profit/sales as a surrogate for the Lerner Index).⁵ In general, these studies found a positive relationship between concentration and accounting profitability measures.

A major criticism of these studies has been the potential introduction of biases by the use of Accounting Rates of Return (ARR). Benston [1985] discussed several studies that have used the ARR as a surrogate for the Internal Rate of Return (IRR). He pointed out that if the differences between accounting measures and economic measures were randomly distributed with respect to variables of interest, the result would not seriously alter inferences about the parameters of interest. Opinion is divided, however, as to whether this error is actually randomly distributed. Hagerman and Senbet [1976], and Long and Ravenscraft [1984] have claimed that there is no correlation between choices of accounting techniques and concentration, while others (e.g., Watts and Zimmerman [1978], Dhalival, Salamon and Smith [1981], and Hagerman and Zmijevski [1979]) have found systematic

differences in the accounting methods adopted by firms of different sizes. This latter result implies that the relationship between concentration and profits may be present because accounting mismeasurement is positively correlated with size (which in this case is presumed to be a measure of concentration).

Several other researchers (see Harcourt [1965], Livingstone and Salamon [1978], and Fisher and McGowan [1983]) developed theoretical models of firms' investment patterns and showed that the ARR is a poor surrogate for the firms' IRR, and therefore according to Fisher and McGowan,

Accounting Rates of Return even if properly measured provide almost no information about economic rates of return. (See Fisher and McGowan [1983], pp. 82)

Salamon [1985] came to a similar conclusion using conditional IRR estimates and said that it was the nature and the strength of the measurement error in the ARR, rather than the correctness of the underlying economic argument, that accounted for the observed positive association between firm size and firm profitability.

In his discussion of the Fisher and McGowan article, Martin [1984] pointed out that in economics literature there are different definitions of economic profit, only one of which is the discount rate that equates the present

value of its expected net revenue stream to its initial outlay (the definition of the IRR used in the theoretical studies mentioned before). Another definition that arises normally in the formal models of profit maximizing firms is the Lerner Index of monopoly power (or the pricemarginal cost margin that was defined in footnote 5). Martin further goes on to state that,

It is doubtful whether any measure of profitability can be unambiguously identified as "correct" to the exclusion of all others, for the purposes of economic analysis. Fisher and McGowan's discussion of what they call the "economic rate of return" does not establish that measures of profitability based on the Lerner index are inappropriate for economic analyses. (Martin [1985], pp. 582)

The model developed in the proposed study is based on the Lerner Index of monopoly power. Unlike earlier models, however, where accounting values are treated as surrogates of economic values, this model relates market power directly to accounting values of gross profit, plant and equipment, and the value of raw materials used in the production of the goods sold.

III THE MODEL

The model relating accounting profit rate to market power is a variation on the model developed by Martin [1985], and Clarke, Davies and Waterson [1976]. Model development is divided into three sections. The first section gives the background to a simple model of a profit maximizing firm in an oligopolistic environment. In the second section, the model is developed further to encompass factors of production, and relate the firm's profit rate (specifically its accounting profit rate) to its market power. In the third section, effects of accounting techniques on the key parameters of the model are analyzed.

3.1 A simple model of a profit maximizing firm

Market power can be thought of as the ability of a firm to influence the price of its own product. This notion of market power can be captured in the context of a profit maximizing firm in a monopolistic environment (i.e., the case of a single seller economy) as,

$$T = P(Q)Q-C(Q)$$
 (1)

where T is the profit, P is the price that the firm charges for its output, Q is the quantity sold, and C(Q)is a cost function. Note that here the price P is a function of the quantity sold, Q. In this simple case, it is assumed that the quantity of goods sold is equal to the quantity of goods manufactured. Assuming the usual conditions for a production function that are neccessary for profit maximization, the first order conditions give,⁶

$$\frac{P - MC}{P} = \frac{1}{\epsilon_{OP}}$$
(2)

where MC is the Marginal cost, and $\epsilon_{\rm QP}$ is the price elasticity of demand for the firm. By adding and subtracting average cost (AC) from both sides of the equation,

$$\frac{P - AC}{P} = \frac{1}{\epsilon_{OP}} + (\frac{1}{FC} - 1)\frac{AC}{P}$$
(3)

where $FC = \lambda C/MC$ is the function coefficient is a measure of the economies of scale of a firm. The function coefficient will be large when the cost of production of a marginal unit is small as compared to the average cost of production. A function coefficient of one implies constant returns to scale.

Extending this case of a single seller to an oligopolistic environment means that now the price will depend on not only the output of the firm but also on the output of all rival firms. Defining α as a measure of collusion among firms as,

$$\alpha = \frac{Q-i}{Qi} \frac{dQi}{dQ-i}$$
(4)

where gi = firm's own output, and $g-i = the output of the rival firms. Making the assumption that <math>\alpha$ is a constant,

and maximizing the profits of the firm in an oligopolistic environment, eqn. (3) becomes,⁷

$$\frac{P - AC}{P} = \frac{\alpha + Si(1 - \alpha)}{\epsilon_i} + \left[\frac{1}{FC} - 1\right] \frac{AC}{P}$$
(5)

here Si = $\frac{9i}{9-i+9i}$ is the firm's market share, and (i is the price elasticity of demand for firm i. For the simple model given by eqn. (5), the test for market power involves the intercept term $[\alpha+Si(1-\alpha)]/(\epsilon)$. When the firm is in pure competition, the price elasticity of demand (i + α , and so the intercept should + 8. Therefore, the presence of market power would imply an intercept that is not equal to 'zero'.

In order to estimate eqn. (5) for a profit maximizing firm, an evaluation of the optimal factors of production or AC is needed. This is done in the next section by maximizing the one period economic profit (change in value) of a firm.

3.2 Factors of Production and Accounting Values

To evaluate the cost of the factors of production, AC in eqn. (5), the present discounted value of firm is maximized over one period. It is assumed that the firm invests in capital at the beginning of the period, and produces and sells at the end of the period. Then the present discounted value (PDV) of a firm can be written as the discounted value of the firm at the end of the period less the value at the beginning of the period, i.e.

$$PDV = - P_{\theta}^{R}(R_{E} + I) - b_{\theta}^{R}(Sr_{t} + \overline{R}_{t}) + (\underline{1}_{1+r^{*}})(P(S_{t+1}, \overline{Q})\overline{Q})$$
$$- WL - C^{F}(S_{t+1}) - C^{R}(Sr_{t} + \overline{R}_{t}) + P_{1}^{K}(1 - \gamma)(R_{E} + I)$$
$$+ b_{1}^{R}(1 - \Delta)(Sr_{t} + \overline{R}_{t}) + P(S_{t+1}, \overline{Q})S_{t+1})$$
(6)

here,

- P^K₈ = Per unit price of capital stock other than rav materials (plant and equipment) at the beginning of the period.
- K = Endowment of Plant and Equipment inherited from the last period.
- I = Investment in Plant and Equipment during the time period.
- b_{θ}^{R} = Price per unit of capital consisting of raw materials.
- Srt = Stocks of raw materials at the beginning of the period(beginning inventory of raw materials).

R₁ = Rav materials purchases.

S₁ = Stocks of finished goods inventory at the beginning of the period.

r* = Discount rate for the period.

P = Price of goods sold.

- \overline{Q} = Quantity of goods sold.
- W = Per unit labor cost.
- L = Total units of Labor used during the period.
- C^{F} = Per unit holding cost for finished goods inventory.
- C^{R} = Per unit holding cost for raw materials inventory.
- P_1^K = Per unit price of Plant and Equipment at the end of the period.
- γ = Depreciation rate Plant and Equipment.

- b_1^R = Per unit price of raw materials at the end of the period.
- Δ = Rate of usage of raw materials. Note: Δ = 1 implies no ending raw materials inventory.

It is important to note the following points about equation (6) for the PDV.

- 1. This equation differs from the PDV formulation in earlier models, in that, both raw materials and finished goods inventories have been incorporated in the model. It is no longer assumed that all of the goods manufactured are sold. This inclusion of inventories in the model has been done in order to analyze specifically the effects of alternative accounting inventory techniques(LIFO vs. FIFO) and to separate their effects from the effects of alternative long lived asset valuation techniques (e.g. HC vs. RC, and SL vs. DDB)
- 2. In this model price P is a function of the quantity of finished goods sold Q, and the ending finished goods inventory S_{t+1} . This formulation is based on the argument that the presence of finished goods inventory helps meet consumer demands more efficiently, and therefore, the firm should be able to charge a higher price for its product which includes the provision of this service.
- 3. Last period's finished goods inventory, S₊, is not

explicitly included in the model because its value to the firm or its opportunity cost is not determinable. This is because it is assumed that the firm can only produce the good and is not able to buy it in the outside market.⁸

4. The per unit inventory holding costs, C^F, and C^R for finished goods and raw materials respectively, are assumed to be constants. Also it is assumed that all costs to hold last period's finished goods inventory were paid out in the previous period and hence do not feature in this period's decision problem.

Therefore, the decision problem is to maximize PDV or equivalently $(1+r^*)$ PDV.

Equation (6) can be simplified by substituting $\mu \tau$ and μ_{Δ} in the equation where,

$$\mu_{\gamma} = \mathbf{r} \star + \gamma + (1 - \gamma) \left(\frac{\mathbf{P}_{1}^{K} - \mathbf{P}_{0}^{K}}{\mathbf{P}_{0}^{K}} \right)$$

and
$$\mu_{\Delta} = \mathbf{r} \star + \Delta + (1 - \Delta) \left(\frac{\mathbf{b}_{1}^{R} - \mathbf{b}_{0}^{R}}{\mathbf{b}_{0}^{R}} \right)$$

Substituting μ_{γ} and μ_{Δ} in eqn. (6), the decision problem becomes,

$$\begin{array}{rcl} & \text{Max} & P(S_{t+1}, \overline{Q})\overline{Q} - WL - C^{P}(S_{t+1}) \\ & & (L, I, R_{t}, S_{t+1}, Q) \\ & & - C^{R}(Sr_{t} + \overline{R}_{t}) - \mu_{\gamma} P_{\theta}^{K}(R_{E} + I) \\ & & - \mu_{\Delta} b_{\theta}^{R}(Sr_{t} + R_{t}) \\ & & + (1/1 + r^{*}) P(s_{t+1}, Q)S_{t+1} \\ & & \text{s.t.} S_{t} + Q(R_{E} + I, L, Sr_{t} + \overline{R}_{t}) = \overline{Q} + S_{t+1} \end{array}$$

here, $Q(K_{\rm E} + I, L, Sr_{\rm t} + R_{\rm t}) = Quantity of finished goods$ $manufactured.⁹ Note that in this simplified form, <math>\mu_{\gamma}$ can be thought of as the implicit rental for services of a unit of capital consisting of Plant and Equipment that costs a dollar, and μ_{Δ} is likewise the implicit rental for capital that consists of raw materials.

A detailed mathematical analysis of the following section is given in the appendix. The first order conditions (eqns. (λ .1), (λ .2), (λ .3), (λ .4), and (λ .5) in the appendix) are,

$$\delta P / \delta S_{t+1} (\overline{Q} + (1/1+r^*) \mathbf{s}_{t+1}) = C^F + \mu = MC$$
 (9)

$$\mu_{\gamma} P_{\theta}^{K} / MP_{K} = \mu = MC \qquad (10)$$

 $(C^{R} + \mu_{\Delta} b_{\theta}^{R}) / MP_{R} = \mu = MC \qquad (11)$

$$W/MP_{L} = \mu = MC \qquad (12)$$

here MP_{K} , MP_{R} , and MP_{L} are the marginal products with respect to capital (K and R) and labor respectively.

Using eqns. (10), (11), and (12) to evaluate the factors of production we get eqn. (A.9) from the appendix,

$$\frac{PQ - WL - b\overline{R}_{t} - C^{R}R}{PQ} = (1 - FC + \frac{\omega FC}{\epsilon_{i}}) + (\frac{1 - \omega}{\epsilon_{i}})FC.Si + K + \frac{\mu_{\gamma}P_{0}K}{PQ} + \frac{bR_{t}}{PQ}(h^{\mu} - 1)$$
(13)

here h is a parameter that characterizes an inventory policy under which the ending inventory of raw materials available for use are kept at a constant proportion to the optimal quantity of raw materials to be used in production (see the appendix for a complete discussion). Equation (13) in terms of accounting Sales Revenue, and accounting Cost of Goods Sold is (from eqn. (A.12) from the appendix),

Accounting Gross Profit =
$$(1 - FC + \frac{\alpha FC}{\epsilon_i}) + (\frac{1 - \alpha}{\epsilon_i})^{FC.Si}$$

Sales $+ \frac{\mu_{\gamma}}{\alpha I} \frac{1}{\frac{\Delta I}{CGS} + 1} \frac{Plant and Equipment}{Sales}$

Change in the values rav
of raw material material
+
$$\binom{h\mu}{\Delta} = 1$$
 [1] [inventories + purchases] (14)
 $\frac{\Delta I}{CGS} + 1$ Sales

So the equation to be estimated is of the form,

 $\frac{\text{Sales} - CGS}{\text{Sales}} = a_0 + a_1S_i + a_2[\frac{1}{\Delta I}] + a_2[\frac{1}{\Delta I}]$ Sales $\frac{\Delta I}{CGS} + 1$ Sales

here ΔI is the change in finished goods inventory and e_i is a stochastic error term. For the evaluation of market power, the coefficients of interest are,

$$\mathbf{a}_{\mathbf{R}} = 1 - \mathbf{FC} + (\alpha \mathbf{FC}/\epsilon_{i}), \qquad (16)$$

and,
$$a_1 = [(1 - \alpha)/\epsilon_i]FC$$
 (17)

From eqns. (16) and (17), $[\alpha+S_i(1-\alpha)]/\epsilon_i$ can be shown to be,

$$\frac{\left[\alpha+S_{i}(1-\alpha)\right]}{\epsilon_{i}} = 1 - \frac{1-\frac{1-S_{i}}{FC}}{FC}$$
(18)

Now for the pure competition case, as $\epsilon_i \rightarrow \infty$, 1 - [1 - $(a_0 + a_1S_i)$]/FC should $\rightarrow 0$. Therefore, the test criterion

for the presence of market power depends on, 1 - [1 - $(a_g + a_1S_i)$]/FC being + 0. For the case when $\alpha > 0$, the test criterion for the presence of market power is of the form, ¹⁰

 $1 - FC < a_{g} + a_{1}S_{i}$ (19) Hence, the critical coefficients for the test of market power are a_{g} , and a_{1} .

3.3 Effects of Accounting Techniques on the coefficients of interest

In the model given by eqn. (15) it is assumed that sales S, and market share S_i , can be measured without error ¹², while the CGS, Plant and Equipment, and inventories are susceptible to errors introduced by accounting techniques. The possible estimations of the model are given in Table 1.

Table 1. Combinations of Accounting Techniques.

RC	(LIFO) ¹³	HC(LIFO)	HC(FIFO)	
SL	1	2	4	
DDB	-	3	5	

By comparison of the parameter estimates in each of the

above cases, it is possible to test the effect of the alternative combinations of techniques on the test criterion given by eqn. (19), the magnitudes of the individual coefficients (a_g, a_1, a_2, a_3) , and on the fitted R² of the estimated equations.

From econometric theory it can be hypothesized that even though the coefficients of interest, a_g and a_1 , do not involve variables with systematic errors, the errors in other variables, however, can impact the coefficients of interest. The magnitude and direction of this bias will depend on the relative magnitudes of the covariance matrices of errors with respect to the covariance matrices of the true variables.¹⁴ In eqn. (15), it can be assumed that the error in the term (1/((Δ I/CGS) + 1)) is very small.¹⁵ Therefore, the errors in the last two variables on the right hand side of eqn. (15) are primarily due to errors in plant and equipment, and raw material inventory measurements respectively. The effects of these errors can be highlighted by comparing different estimations as shown in the following table.
-						
	Comparison of	Highlights the effect of	ł	conditional on		
λ	(1) with (2)	HC vs. RC	ł	SL		
B	(2) with (4)	LIFO vs. FIFO	ł	HC and SL		
С	(3) with (5)	LIFO vs. FIFO	ł	HC and DDB		
D	(2) with (3)	SL vs. DDB	ł	HC and LIFO		
E	(4) with (5)	SL vs. DDB	ł	HC and FIFO		

Table 2. Pairwise comparison of alternative accounting accounting techniques.

IV HYPOTHESES

The following hypotheses are formulated in order to test the robustness of the estimated model with respect to alternative accounting techniques. The hypotheses are divided into three parts.

As a preliminary test, the first set of hypotheses can be formulated in order to test the robutness of the induvidual estimated coefficients. The hypotheses can be stated as,

(la) H_gla: a_c = 0 H_Ala: a_c = 0

here superscript $c = 0, \ldots, 3$ corresponds to the coefficients a_0, \ldots, a_3 .

(1b) H_0 lb: $a_c^{j;m} = a_c^{'j;m}$ for all j, j, 'm, and c $j \neq j'$

$$H_{\lambda}^{lb}: a_{c}^{jim} \neq a_{c}^{jim}$$

here (j,j') denote a pair of alternative accounting techniques(e.g. LIFO, FIFO), and m denotes a particular combination of techniques that j and j'are conditional on.

The null hypothesis la states that the individual estimated coefficients are insignificantly different from zero. The rejection of this hypothesis and proper signs on the coefficients would indicate that the model is robust with respect to estimation.

The null hypothesis (1b) states that there is no effect of accounting techniques on the estimated coefficients. This would imply that the model is robust with respect to accounting techniques.

Hypotheses (1a) are tested by examining the 't' statistics of the estimated coefficients (see Christie [1986]). Hypotheses (1b) are tested by using the 't' test for difference in means with unknown variances.

A second set of hypotheses can be formulated based on the fitted R^2 of the estimated equations.

(2)
$$H_0^{1}$$
: $R^2_j = R^2_j^{*}$
 H_{λ}^{1} : $R^2_j \neq R^2_{j^{*}}$

here j, and j' refer to different sets of accounting techniques.

These hypotheses are tested in two ways. First, the form of analysis outlined in section 3.3 can be viewed as a single factor experiment with 8 subjects or firms being measured under combinations of 5 levels of factors (accounting methods, combinations 1,...,5 in Table 1). The proposed technique is the repeated measures technique described by Myers [1972 (pp. 186-187)]. Second, the

above hypotheses are tested using nonparametric tests such as the Kendall's coefficient of concordance and the Spearman's correlation coefficient.

The third set of hypotheses are formulated in order to test the senstivity of the test criterion given by eqn. (19). Under the assumption of constant returns to scale, 16 the test criterion(T) for market power is,

$$\mathbf{T} = \mathbf{a}_0 + \mathbf{a}_1 \mathbf{S}_1 > \mathbf{0}$$

Using this test criterion, the firms can be ranked by market power. The stability of this ranking system with respect to estimation techniques is an important policy issue. The hypotheses based on the test criterion can be stated as,

(3a)
$$H_{g}$$
3a: $T^{j} = T^{k}$ for all j,k
j + k

 H_A^{3a} : $T^{j} \neq T^k$

where $j,k = 1, \ldots, 7$ denote the combinations of techniques given in Table 1.

(3b) H_g3b: $T^{j+m} = T^{j+m}$ for all j, j', and m j+j' H_A3b: $T^{j+m} + T^{j+m}$ The null hypotheses state that the test for market power is stable with respect to alternative accounting techniques, whereas the alternative hypotheses state that it is not. Hypotheses (3a, and 3b) are tested by examining the shifts in firm rankings due to alternative accounting techniques. This shift in ranks is tested by using a non parametric test such as the Kendall's coefficient of concordance and the Spearman's correlation coefficient.

A rejection of the alternative hypotheses (3a, and 3b) would indicate that the estimation of the model is robust with respect to accounting methods.

V METHODOLOGY

The methodology consists of estimating equation(15) with time series data at the firm level. This method is preferable to using industry level crossectional data for the following two reasons. <u>First</u>, as the risks faced by different firms in an industry are different, it is unlikely that the implicit rate of rental for capital stock (μ) is the same for every firm in the crossectional sample. Therefore, the use of crossectional data is not appropriate for the estimation of this model. The <u>second</u> reason is that the issue of firm specific market power is aquestion with direct policy implications, and therefore its stability with respect to estimation techniques is worth investigating.

In, summary, the research method is to restate reported inventory and capital stock figures for a sample of eight firms. This restatement yields a set of five financial statements for each firm, i.e. 40 in all. The model is first estimated with the reported data and then with the restated data. The results are then compared in order to examine the effects of alternative accounting methods on the relevant estimated coefficients.

5.1 <u>Restatement Procedure</u>

One of the ways of examining the effects of accounting techniques on economic analyses proposed by Gonedes and Dopuch [1979] was to restate reported accounting numbers of the firm and repeat the economic analyses with the restated numbers.

The restatement routines used in this study for inventory and depreciation amounts are the methods proposed by Nair [1977] which basically make use of the Dollar Value LIFO method for inventory and asset layering techniques for depreciation restatements respectively. The restatement to Replacement Cost is done using the method proposed by Falkenstien and Wiel [1977].

5.2 <u>Inventory Restatement</u>

The restatement method used for inventories is the Dollar-Value LIFO method. This method essentially removes from the inventory effects that are solely due to price changes. These methods are illustrated in detail by Nair [1977]. As Nair pointed out, in order to apply these restatement procedures the following simplifying assumptions have to be made. Though these assumptions are fairly restrictive, they can be relaxed if more detailed data were available. The assumptions are,

1. For firms which use combintions of techniques for inventory, it is assumed that the primary method of inventory valuation is used for all inventory. This simplifying assumption is neccessary because not every firm in the sample discloses the precise breakdown of the different inventory methods in use for that firm. If more precise data were available, it would be better to restate only the relevant portion. For example if a firm has 88 percent of its inventory on FIFO and 28 percent on LIFO, restatement to LIFO should be done only on 28 percent of the inventory.

2. In restating LIFO inventory to FIFO it is assumed that all inventory is aquired 5 years before the test period. Similarly for FIFO to LIFO restatement, the ending FIFO inventory 5 years before the test period is assumed to be the beginning LIFO inventory. This is done in order to approximate what the inventory might have been had the firm always been using LIFO.

3. The indices used in this study are the wholesale price indices for various commodity groups that are published on a monthly basis by the <u>Survey of Current</u> <u>Business</u>.

4. For firms that were divesified into several lines of business, the price index used is the one for its primary line of business.

5.3 <u>Restatement for Depreciation</u>

The method outlined by Nair for restating depreciation figures is based on layering assets by using capital expenditure data. These layers are depreciated according to the alternative depreciation techniques and the resultant depreciation amounts aggregated to yield the total depreciation expense for the period. Since in this study it is the restated capital stock series that is used as one of the independent variables, this series is reconstructed using restated depreciation figures beginning from a period three years before the test period. For example, the restated capital stock figure for the year 19x is computed as shown below,

Restated Capital Stock 19x=Capitol Stock 19(x-3) +Eaquisitions- Eretirements -Edepreciation+Eadjustments to accumulated depreciation account

Where the aquisitions etc. are summed over the three periods prior to the test period. The following basic assumptions are made in the restatement of depreciation amounts.

 The first-in-first-out flow is assumed with respect to property acquisitions and retirements.

2. The salvage value is assumed to be zero.

 Multiindustry firms are restated using the useful life most appropriate for their primary line of business.

5.4 Restatatement to Replacement Cost

The method used for restatement to Replacement cost is proposed by Falkenstien and Wiel [1977]. The price indices used for restatement are published by the Bureau of Labor Statistics's Wholesale Sale Price Index (WPI) for various commodities.

According to Falkenstien and Wiel, FIFO inventory method is a good approximation for replacement cost of inventories, and LIFO cost of goods sold with adjustments for depleted inventory layers is a good approximation of the replacement cost of goods sold.

The procedure is somewhat more complicated for restating the Capital Stock series. The method requires breakdown of the Plant and Equipment account into separate building and equipment accounts and determining the average age of depreciable assets. This is estimated by dividing the accumulated depreciation balance by the repective depreciation charges for that period and then averaging over the test period. This estimated average is a reasonable proxy for the base year from which to measure the replacement cost of fixed assets. Finally appropriate price indices are applied to the net building and equipment accounts to yield the replacement cost of Plant and Equipment.

The basic assumptions and caveats for restatement to replacement cost are,

1. A major approximation in this restatement procedure is the overestimation of the average age of Plant and Equipment which results in overestimates for the increments used to adjust historical cost to replacement cost. As mentioned above, the formula used for estimating the average age is,

Average age of = Accumulated Depreciation depreciable assets depreciation charge for the year

However, if the firm is using fully depreciated assets, then the balance sheet shows the historical cost of these assets while deducting accumulated depreciation equal to that amount. Fully depreciated assets add nothing to the balance sheet totals nor to the depreciation charges. Yet, because they remain on the balance sheet, they increase the numerator of the above equation without adding to the denominator, thus overestating the average age of depreciable assets. This overestimation is impossible to remove without detailed knowledge of the fully depreciated assets profile for each firm.

2. Another approximation where replacement cost can be different from that computed by the firm is due to the fact that the firm's internally developed price indices could be different from the price indices that are available from the Bureau of Labor Statistics.

Due to the above reasons the restated values of Plant and Equipment are at best an approximation of the actual replacemnt costs. The above mentioned overstatement could bias the results towards significant effects of the change to replacement cost. This potential bias is discussed more fully in the results and analysis section.

5.5 <u>Deta</u>

The time series data for 40 quarters was obtained from the quarterly SEC reports. The sample consists of eight firms in two durable goods industries that have traditionally been interesting to researchers and policy makers as to their levels of concentration. The industries, firms, and their reported methods of accounting techniques are given in Table 3.

Firm	Industry SIC code	Firms Reported accounting methods
Armco Corp.	3310	LIFO SL HC
U.S. Steel Corp.		LIFO SL HC
Lukens Corp.		LIFO SL HC
LTV Corp.		LIFO SL HC
General Motors Corp	. 3711	LIFO DDB HC
Ford Motor Co.		LIFO DDB HC
American Motors Cor	P.	FIFO DDB HC
Chrysler Corp.	-	FIFO DDB HC

Table	3.	Firms,	Industry	SIC	codes,	and	reported
		account	ing metho	ds.			

There was only one firm, U.S. Steel, which had a significant merger (over 48 percent of total assets) with Marathon Oil, a firm that is in a different Indsustry. This neccessitated netting out the effects of this merger for the years 1982-1984. In order to compute market shares, industry data for the Iron and Steel was obtained from the publication 'Iron and Steel Institute' which publishes monthly statistics on shipments and imports of Iron and Steel. The source of industry data for the automobile industry is the 'Ward's Automotive Book'.

5.6 <u>Methods of Analysis</u>

As mentioned before, the main purposes of this study are twofold, (1) to examine the performance of the analytical model developed in section 3.1, and (2) to evaluate the effects of alternative accounting techniques on the performance of this analytical model. For the first purpose, the results of estimation are examined with respect to significance and meaningfulness of the estimated coefficients. For the second purpose, the results due to estimation with alternative accounting techniques are compared with respect to the overall goodness of fit (R^2) , the significance of the individual coefficients, and the magnitude of the test criterion $(a_0+a_1S_i)$ for market power. For this set of analyses the null hypothesis is that for a particular firm, change in accounting methods will cause no difference in either the goodness of fit of the equation, or in the estimated coefficients that are relevant for the purposes of estimating market pover.

In order to test the first set of hypotheses, the parametric 't' test is used to test the significance of individual coefficients.

5.6a Analysis of Variance

The second set of hypotheses formulated in section IV are tested by using a specific version of the Analysis of

Variance technique called the Repeated Measures technique [Meyers 1972 pp. 164-86]. This technique is particularly appropriate for an experimental design where different treatments are applied repeatedly to the same subjects. This technique reduces error variances thereby improving the power of the test, but results in observations that are dependent. The form of analysis outlines in section 3.3 can be viewed as a single factor experiment with 8(or four per industry) subjects (firms) being measured under combinations of five levels of factors(combinations of accounting techniques). The dependent variable in this analysis is the adjusted R² of the estimated equations which is an accepted measure of the explanatory power of a model.¹⁹ The proposed model underlying the analysis of the dependent variable R² is as follows,

 $Y_{jk} = \mu + \alpha_{j} + \beta_{k} + \epsilon_{jk} \text{ with } j=1,\dots,5, \text{ and}$ $k=1,\dots,8$

here $\alpha_j = \mu_j - \mu$ is the main effect of the accounting method factor.

 $\beta_{\mu} = \mu_{\mu} - \mu$ is the main effect of the firm factor.

The null hypotheses of interest can be stated as,

$$H_{02a}: \qquad \alpha_j = 0 \text{ for all } j$$
$$H_{02b}: \qquad P_k = 0 \text{ for all } k$$

The assumptions underlying the Repeated Measures technique are that,

- 1. ϵ_{ik} are i.i.d normal.
- 2. $E(\epsilon_{jk}) = 0$ and $E(\epsilon_{jk})^2 = \sigma^2$ i.e. that the variance of all ϵ 's is the same for all subjets.

Though these assumptions are restrictive and hardly met for most populations, Box [1953] points out that in the case where group sizes are equal, the analysis of variance is not very sensitive to nonnormality or inequality of variances. It is important to point out, however, that the condition for randomness of sample selection is not completely met. The sample selection could be considered random to the point that the sample of 8 firms (4 in each industry) were chosen out of the 101 firms for which requests had been sent for data. The two industries chosen for the sample were the ones for which atleast four firms complied with the request for data. Therefore, the selection of a particular firm is not expected to be correlated with the changes in the dependent variable due accounting techniques.

The third set of hypotheses formulated in section IV can be tested by using nonparametric statistics. Under the assumption of constant returns to scale, the test criterion for market power given by equation(19) under the assumption of constant returns to scale is $a_g + a_1 S_i$. Therefore, the null can be stated as,

$$H_{03}$$
: $T_j = T_k$ for all j,k, j=k

where T denotes the test criterion, and j and k denote different combinations of accounting techniques.

The significance of changes in estimated market power due to different accounting techniques is tested by examining the significance of the change in the rankings of firms by market power. The ranking of firms by market power within a particular industry or within the larger set of all firms, is a relevant question when considering policy issues such as merger guidlines or antitrust litigation. A statistically significant shift in the ranks due to estimation with a different combination of accounting techniques would indicate a significant effect of accounting techniques on the estimation of market power. The test used to measure this shift in ranks is the Kendall's coefficient of concordance W (Siegel [1956], pp. 229-38] and the Spearman's rank Correlation Coefficient [Siegel, 1956, pp. 205-13]. These methods are particularly appropriate for small samples. The same methods are used to test the stability of the rankings of firms by the best goodness of fit measure for the estimated model. Also, Kendall's Coefficient of Concordance is used to test the rankings of accounting methods for the best fit (i.e. which method leads to the results with the most expanatory power for a particular firm). These tests are performed in two parts, first with the partial sample at the industry level(i.e. for the four firms in each industry), and <u>second</u>, with the combined sample of both industries. This two part testing of results is done in order to highlight the inter as well the intra differences in results between the two industries incorporated in the sample.

5.6b Kendall's coefficient of concordance W

This statistic is used to measure the strength of the agreement among different rankings of certain entities by differnt ranking mechanisms. In this study the entities are the firm's Test Criterion and Adjusted R^2 and the

ranking mechanisms are the accounting techniques. The statistic W can be defined as

$$W = \frac{S}{\frac{1}{12} k^2 (N^3 - N)}$$

Where S = Sum of the squares of deviations from the mean
of the sum of ranks assigned to each entity
being ranked.
k = Number of sets of ranks = number of ranking

mechanisms N = Number of entities ranked.

This method is illustrated by using the data in Table 59

Firms Acct. Methods	Àrm co	USS	Lukens	LTV
LIFOSLHC	2	1	3	4
LIFODDBHC	2	1	3	4
FSH	2	4	3	1
FDH	2	4	3	1
R	2	1	4	3
Sum of the ranks Mean of the	10 e ranks = 50	11 8/4 =12.5	16	13
S = (10-12 + (13	.5) ² + (11 - 12.5) ² = 2	- 12.5)* + (16 21	5 - 12.5) ²	
Here $k = 5$, N = 4			
Therefore	$W = \frac{1}{(1/12)}$	$\frac{21}{(5)^2}$	- 0. 168	

The Coefficient of Concordance W, expresses the degree of agreement in rankings among the five methods of estimation. It is important to note that a W equal to 1 indicates perfect agreement among the five methods of estimation and therefore implies that accounting techniques have no effect on the estimated results. The Null hypothesis for this statistic is that of complete independence of rankings, l.e. W = 8, therefore, by the nature of this statistic, in order to show significant effects of accounting techniques we need to accept the Null²⁸. In order to be statistically significant, the value of estimated S has to be >= the critical value of S that is available from the tables for this technique. Critical value of S from table P (Siegel pp. 286) is 62.6. Since the estimated S = 21 is not >= 62.6, the null hypothesis of independence of rankings is not rejected. Therefore in the case of Iron and Steel Industry, Accounting techniques have significant effects on the ranking of firms by market power.

5.6c Spearman's Rank Correlation Coefficient r_

Tests using this coefficient are appropriate when the agreement between two sets of rankings is being analyzed (e.g. for the pairwise comparisons like LIFO vs. FIFO: SLHC. This method of analysis is illustrated by using the data for the Test Criterion in table 60.

Ranks

Firms	a Armco	USX	Lukens	LTV	
Combination of Techniques					
LIFOSLHC FIFOSLHC	2 2	1 4	3 3	4 1	
Difference di	8	-3	8	3	
5444 - 1					

Then the Spearman's Rank Correlation Coefficient is given by,

 $r_{s} = 1 - \frac{6\Sigma di^{2}}{N^{3} - N}$

Since N = 4, therefore r^{S} = -0.80. Here r_{s} = 1 would indicate perfect agreement between different systems of rankings or in other words no effects of accounting techniques. From the table of critical values of r_{s} (Siegel pp 284) critial value of r_{s} for N = 4 is 1.00. Therefore -0.8 <u>not</u> being >= 1.00, indicates that the null hypothesis of independence of rankings cannot be rejected. Once again here, accepting the null implies that accounting techniques have significant effects. Therefore in the comparison for the results of LIFO vs. FIFO: SLHC, the results show that the effects of the change from LIFO to FIFO on the rankings by market power are statistically significant.

VI EMPIRICAL RESULTS AND ANALYSIS

The results in this section are divided into five parts. Part (1) consists of general results about the performance of the model. Part (2) gives the results of a parametric 't' test which is done in order to test the effects of accounting techniques on the individual estimated coefficients. Part (3) gives the preliminary analysis of variance results about the effects of accounting techniques on the goodness of fit of the estimated equations. Part (4) gives the results of nonparametric tests such as the Kendall's Coefficient of Concordance and the Spearman's Correlation Coefficient on the rankings of adjusted R^2 of the estimated equations. . Finally part (5) gives the results of the nonparametric tests for the robustness of firm rankings by the market power test criterion. These nonparametric tests are conducted in order to analyze the effects of accounting techniques on the estimated results jointly (using Kendall's coefficient of concordance) as well as on the basis of pairwise comparison (e.g. LIFO vs. FIFO or SL vs.DDB), (using Spearman's correlation coefficient).

6.1 <u>General Results</u>

The results of regressions with reported and restated variables are presented in Tables 4 to 11.

The F statistic significant for every estimation of the model. A summary of these tables for the adjusted R² and the test criterion T are given in Tables 12 and 13. The adjusted R² of the fitted equations ranges from 0.31 for Lukens (LIFODDBHC) to 0.98 for General Notors (LIFOSLHC and LIFODDBHC). The Test Criterion (T = $a_0 + a_1 MS$) ranges from 0.25 for LTV (LIFOSLDDB) to 1.298 for Ford Motor Co. (FIFODDBHC).

Table 14 gives the results for hypotheses (1a). The Z-statistic (see Christie [1985]) for individual coefficients is significant for the constant term, a_g, the coefficient for plant and equipment, a_2 , and the coefficient for raw materials, a_. The market power test criterion, T, is also highly significant. Contrary to model predictions, the coefficients for market share and plant and equipment frequently have negative signs. The negative sign for the market share variable , however, is insignificant. The reason for the insignificance of the market share coefficient may be due the fact that since market share for a firm changes very slowly, it is probably highly correlated with the constant term. As for the negative coefficient for the plant and equipment variable, the negative sign persists only for the Iron and Steel industry. This may be due to the fact that for the time period under analysis, this industry has been documented as making large capital expenditures for

inefficient and outdated plant and equipment (see Adams [1986] pp. 182), which could be a possible explanation for the observed negative returns to plant and equipment increases.

6.2 Parametiric t test

This test is done in order to test Hypotheses (1b). The results are given in Table 15. As can be seen each of the comparisons is highly significant.

6.3 Analysis of variance

As mentioned in section 5.5, the dependent variable for the analysis of variance is the adjusted R^2 of the estimated equations. The results for the analysis of variance are given in Tables 16, 17, and 18. From the significant F statistics it can be seen that there is an insignificant effect of accounting techniques for either the subsamples of the two industries or for the combined sample. It can be noted that even though statistically insignificant, the effect of accounting techniques is greater for the Iron and Steel industry than for the Automobile industry. On the other hand the firm effect is highly significant in all three cases thereby indicating that the expalanatory power of the model varies accross firms or that there are additional firm specific characteristics that are not captured by this model. This last result is in line with what is to be expected for a simple model such as the one developed in this study.

6.4 Adjusted R²

The adjusted R² of the fitted equation is a measure of the explanatory power of the model. This variable has been studied by researchers in order to rank models (Jorgenson []) and its robustness with respect to accounting techniques has implications for the kinds of inferences drawn from such analyses.

6.4a Kendall's coefficient of concordance

The Kendall's coefficients of concordance for the rankings by the adjusted R^2 are given in Tables 18, 28, and 38. the results indicate that the rankings of firms by the explanatory power of the model are not affected by estimation with alternative accounting techniques. This result is true for the combined sample (Table 18) as well as for the industry subsamples for the Iron and Steel (Table 28) and Automobile (Table 38) industries.

6.4b Spearman's rank correlation coefficient

The consolidated results for the Spearman's correlation coefficient are given in Tables 27, 37, and 47. For this pairwise analysis the result of no effects due to accounting techniques only holds consistently for

the SL vs. DDB. All other comparisons show significant effects.

6.5 Market pover test criterion

The market power test criterion, which as mentioned earlier is a measure of market power of the firm. A statistically significant shift in the rankings of firms by their market power will indicate a significant effect of accounting techniques. This in turn would indicate the possible introduction of a bias in the estimated coefficients due to alternative accounting techniques. The test statistics used to examine the significance of shifts in these ranks are the Kendall's coefficient of concodance, and the Spearman's correlation coefficient.

6.5a <u>Kendall's coefficient of Concordance</u>

The Kendall's coefficient of concordance for rankings by market power are given in tables 49, 59, and 71. the reasults in Tables 49, and 71 indicate that the ranking of firms by market power is not affected by alternative accounting techniques for the combined sample as well as for the subsample for the automobile industry. However for the subsample of the iron and steel industry, there is a significant effect. This result would seem to indicate that the estimation procedure is more robust towards accounting techniques for the automobile industry than for the iron and steel industry.

6.5b Spearman's rank correlation coefficient

The results of the Spearman's Correlation Coefficient test are given in tables 58 - 78 for the test criterion. It can be seen that in the case of ranks based on the Test Criterion, the change from SL to DDB has and insignificant effect, and the change form RC to HCFIFO has a significant effect across all samples. All other changes are statistically significant at the individual industry level but are insignificant at the combined level. It is interesting to note that by rank ordering the r_g values, the effect is the most pronounced for RC vs. HCLIFO (smallest r_g), and the least for LIFO vs. FIFOSL.

Cc	onstant ^a g	Market share ^a 1	Ra PLE Nat ^a 2	terial [©] 3	Adj. R ² :	F Crit Stat. a	rest terion T + a ₁ MS ¹	Durbin Watson
LIFOSLHC	8.634* (6.98)	-0.555 (-1.15)	-0.067* (-3.23)	-0.516 (-4.21)	* 0.69	21.2	8.592* (7.88)	2.22
LIFODDBHC	8.64 3* (6.96)	8.630 (-1.37)	-0.062* (-2.94)	-0.535 (-4.19)	* 0.67)	19.7	0. 596 (6.86)	2.13
FIFOSLHC	0.999* (13.8)	-0.855* (-2.03)	-0.058* (-3.22)	-1.029 (-8.31)	* 0.88	38.1	8.936* (12.1)	2.55
FIFODDBHC	C 1.001* (7.34)	-8.888 * (-2.65)	-0.059* (-6.60)	-1.031 (-3.11)	* 8.79)	36.7*	0.936* (7.54)	1.83
RC	0.568* (7.34)	-0.720* (-2.65)	-0.366* (-6.60)	-0.375 (-3.11)	* 0.75)	28.4*	8.514* (7.54)	1.88
() t ¹ Averaç	t values ge Marke	in para t Share (nthesis 0.074					
* signif	Eicant a	t 0.01						

Table 4. Results of the time series OLS regression with 36 observations for Armco Inc.

Co	nstant	Market share	R: P&E Ma	ev terial	Adj. R ²	F Crit Stat.	terion T,	Durbin Watson
	8	•1	•2	a 3		•	+a,MS*	
LIFOSLHC	8.875*	-0.189	-0.009	-8.948	.88	62 . 8*	0.829*	2.82
	(13.5)	(-1.52)	(-1.49)	(-12.9)			(4.13)	
LIFODDBHC	0. 883×	-0.187	0.011	-0.958	e. 87	62. 8 *	0.838	2.03
	(13.00	(-1.49)	(-1.35)	(-13.0))		(3.99)	
FIFOSLHC	8. 819*	-0.197	-0.013	-0.850	.71	22.4*	0.771	1.85
	(7.16)	(-0.68)	(-0.77)	(-6.40)			(10.2)	
F I FODDBHC	0.823*	-8.192	-0.018	-0.846	. 0.74	22.4*	8.774	1.85
	(7.15)	(-0.66)	(-0.83)	(-6.50)			(18.4)	
RC	0.805*	-0.085	-0.010	-0.865		30.3*	8.784	2.88
	(8.34)	(-0.48)	(-1.47)	(-7.56))		(10.35))
() t	values	in para	nthesis					
¹ Averag	e Marke	t Share	0.243					
* signif	icant a	t 0.01						

Table 5. Results of the time series OLS regression with 36 observations for USX Corporation.

Table 6.	Results of the	r time se	ri es OLS	regression with 36	
	observations :	for Luken	s Steel	Corporation.	

							rest	December 2 -
Co	nstant	market share	PSE Ma	nv terial	R ²	F Cris	T,	Watson
	8	•1	•2	a 3		a	9 ⁺ a ₁ MS ¹	
LIFOSLHC	0.426* (5.40)	3.160 (0.60)	-8.833* (-2.12)	-0.348* (-3.55)	0.32	5.2*	8.442 * (6.85)	2.182
L I FODDBHC	8.418* (5.38)	3.433 (0.65)	-0.830* (-2.82)	-0.344* (-3.49)	8.31	5.0*	0.4 35* (5.86)	2.197
FIFOSLHC	0.864* (9.46)	-6.578 (-1.18)	-0.021 (-1.32)	-0.930* (-7.34)	8.63	21.0*	8.840 * (8.83)	2.21
FIFODDBHC	8.869* (9.79)	5.850 (0.94)	-0.150 (-1.04)	8 .955* (-7.39)	8.66	15.5	0.840* (9.02)	1.98
RC	0.432* (5.63)	-0.655 -0.130	-0.001* (-2.06)	-0.350* (-3.79)	8.36	6.0*	0.428 (6.43)	2.17
() t	values	in pere	nthesis					
1 Averag	e Marke	t Share	8.005					
* signif	icant a	t 8.81						

						Test			
		Market	R	BV	λdj.	F Cri	terion	Durbin	
Co	nstant	share	P&E Ma	terial	R²	Stat.	T	Watson	
	20	•1	•2	a 3		•	9 ^{+a} 1 ^{MS}		
LIFOSLHC	0.295	0.467	-0.092	-0.200	8.67	18.1*	0.251	2.85	
	(1.72)	(0.95)	(-4.38)	(-8.87)			(1.68)		
LIFODDBHC	0.331*	-0.055	0.100×	0.145	0.73	24.2*	8. 326*	2.05	
	(2.12)	(-0.12)	(-5.60)	(-0.73)			(2.38)		
FIFOSLHC	1.069*	-0.121	-0.059*	-1.289	0.82	38.7*	1.058	2.05	
	(8.78)	(-0.16)	(-2.25)	(-10.0)	I		(12.49)	
FIFODDBHC	1.231*	-1.861	0.010×	-1.217	0.85	45.9*	1.228*	2.42	
	(9.96)	(-1.44)	(-4.6)	(-13.82	:)		(15.6)		
RC	0.400×	8.462	0.065*	0.037	0.64	16.7*	0.443*	2.07	
	(2.18)	(0.87)	(-3.52)	(-1.46)	I		(2.56)		
() t	values	in para	nthesis						
1 Average	e Marke	t Share	0.0927						
* signif	icant a	t 0.01							
-									

Table 7. Results of the time series OLS regression with 36 observations for LTV Corporation.

					Test				
0-		Market	RI DE R Mar		Adj.	F Cri	terion	Durbin	
CO		snere B.	ral me	a_ a_	R4	Stat.	т .+а. MS	watson	
	-0	-1	-2	-3			01		
LIFOSLHC	1.878*	-0.178	-0.885	-1.898	8.98	168*	0. 997*	2.13	
	(22.5)	(-0.17)	(-0.72)	(-69.9))		(46.1)		
LIFODDBHC	1.858*	-0.168	8.001	-1.878*	. 0.98	128*	8.984*	1.74	
	(22.5)	(-1.63)	(0.89)	(-78.28			(74.86))	
FIFOSLHC	1.158*	-0.218	8.083	-1.188*	. 0.92	130*	10.60*	2.15	
	(12.8)	(-1.46)	(0.36)	(-17.0))		(20.8)		
FIFODDBHC	1.150*	-8.218	8.884	-1.182	8.91	138*	1.864*	2.15	
	(12.9)	(-1.46)	(0.36)	(-16.95	;)			(20.4)	
RC	1.079*	-1.126	-0.000	1.060*	8.93	185*	1.827*	2.07	
	(17.1)	(-0.93)	(-0.46)	(145.0)			(55.8)		
() t	values	in para	nthesis						
1 Average	e Marke	t Share	0.410						
* signif	icant a	t 8.81							

Table 8. Results of the time series OLS regression with 36 observations for General Motor Corporation.

.

							lest	
		Market	R	8V	Adj.	F Crit	terion	Durbin
Co	nstant	share	PLE Me	terial	R² S	St a t.	T,	Watson
	a .	•1	•2	•3		•	9 +a ₁ MS.	
LIFOSLHC	1.028*	-0.133	-0.011	-1.879*	. 0.94	184*	0.993*	1.96
	(19.7)	(-0.93)	(-0.57)	(-18.1))		(24.2)	
LIFODDBHC	1.020*	(-0.248	(-0.033	-1.006*	8.95	180*	0. 956*	2.19
	(20.9)	(-1.85)	(-1.87)	(-14.8))		(21.8)	
FIFOSLHC	1.328*	-8.778*	0.073	-1.186*	.82	53 .8 *	1.128*	2.65
	(11.2)	(-2.12)	(-1.48)	(-8.31))		(7.69)	
FIFODDBHC	1.298*	-0.675	-0.056	-1.17*	6.81	51.7*	1.123*	2.56
	(11.1)	(-1.91)	(-1.21)	(-7.22)			(7.78)	
RC	8.946*	-0.223	-0.038*	8.983*	8.88	88.5×	0.888*	2.58
	(13.7)	(-1.15)	(-2.69)	(-10.7)			(15.2)	
() t	values	in para	nthesis					
1 Average	e Marke	t Share	8.268					
* signif	icant a	t 0.01						

Table 9. Results of the time series OLS regression with 36 observations for the Ford Notor Company.

Table 18. Results of the time series OLS regression with 36 observations for American Motors Corporation.

Test	
terion	Durbin
T,	Watson
0 ^{+a} 1 ^{M*}	
8.828*	2.68
(9.63)	
0.834*	2.68
(6.98)	
8.717*	2.72
(9.03)	
8.715*	2.71
(9.73)	
8.907*	2.93
(20.5)	
	0.715* (9.73) 0.907* (20.5)

			•			Test	Durah i a
Co	nstant	Market Share	PLE Ma	BV terial	Adj. Rž	F Criterion Stat. T	Vatson
	a ₀	•1	•2	•3		a ₀ +a ₁ MS	
LIFOSLHC	0. 597* (8.47)	8.925 (1.73)	0.03 1 (1.12)	-8.744* (-8.58)	8.87	68.9* 8. 679* (8.83)	2.42
LIFODDBHC	0. 598* (8.95)	0.615 (1.32)	0.012 (0.36)	-0.681* (-8.20)	0.89	53.1* 0.653 (9.02)	2.48
FIFOSLHC	0.533* (7.88)	8.812 (8.821)	-0.032 (-1.21)	-0.481* (-5.21)	.87	58.1* 8.534 (7.82)	2.70
FIFODDBHC	0.525* (8.13)	8.014 (0.03)	-0/050 (-1.53)	-0.468* (-5.23)	0.87	60.1* 0.526* (8.86)	2.70
RC	0.760* (11.2)	0.728* (2.00)	0.014* (2.18)	-0.93* (-14.0)	0.89	69.7* 0 .825* (9.50)	2.06
() t	values	in p ara	nthesis				
¹ Averag	e Marke	t Share	8.889				
* signif	icant a	t 0.01					

Table 11. Results of the time series OLS regression with 36 observations for Chrysler Corporation.
Combination of Techniques			of					
	ARMCO	USX	LTV	lukens	GM	FORD	АМС	CHRYSLER
LIFOSLHC	0.69	0.88	0.67	0.32	0.98	0.94	8.72	8.87
LIFODDBSL	0.69	0.87	0.73	0.31	<u>0.98</u>	<u>8.95</u>	0.72	0.89
FIFOSLHC	0.80	8.71	8.82	8.63	0.92	8.82	8.73	8.87
FIFODDBHC	8.79	8.74	8.85	0.66	0.91	0.81	<u>0.73</u>	9.87
RC	0.75	0.75	8.66	0.36	0.93	0.88	0.92	0.89
den	otes re	ported	comb	ination	of te	chniqu	es .	

Table 12. Adjusted R^2 from the OLS estimation of equation (15).

Table 13. Test Criterion, $a_0 + a_1 MS$ from the OLS estimation of equation (15).

0									
Technique	S ARMCO	USX	LTV	LUKENS	GM	FORD	AMC	CHRYSLER	_or
LIFOSLHC	0.592	0.829	0.251	0.442	0. 997	0.993	0.828	0.679	
LIFODDBSL	0.596	0.838	8.326	0.435	0.984	0.956	8.834	0.653	
FIFOSLHC	0.936	0.771	1.058	8.848	1.060	1.128	8.717	8.539	
FIFODDBHC	0.936	8.784	1.220	0.840	1.864	1.298	0.715	0.526	
RC	0.514	8.784	8.443	0.428	1.027	8.888	0.987	0.825	
	_ denote	es repo	orted (combinat	ion o	f techi	niques	•	

Combination of	Constant	Market Share	P&E Mi	Rav aterial] cri	Cest terion
Techniques	a 0	•1	•2	a 3	T =	a ₆ +a ₁
LIFOSLHC	29.27**	-0.26	-3.38*	-43.3	8××	36.84**
LIFODDBHC	29.61**	-0.20	-4.08*	-41.9	7**	44.80**
FIFOSLHC	26.78**	-1.16	-3.49*	-24.0	4**	30.45**
FIFODDBHC	27.21**	-1.79	-3.38*	-22.9	7**	33.44**
RC	27.69**	-8.51	-4.87*	-33.4	8 * *	43.76**
Reported Combination	29.45**	-1.13	-5.43*	-40.9	8××	45.64**
* Signif ** Signif	icant at .0 icant at .0	1. 01.				

Table 14. Significance (Z-statistic) of individual coefficients.

Coefficients

Comparison	a ₀	•1	a 2	a 3					
LIFO vs. FIFO	**	**	**	**					
SL vs. DDB	**	**	**	**					
RC vs.HC	**	**	**	**					
** significant at	0.01								

Table 15.	't' test for difference	in means	of	individual
	coefficients			

Table 16.	Adjusted R ² , analysis of variance for the 'Iron and Steel' industry.								
Effe ct	Estimated F statistic	Critical value of F	Level of significance						
Firm	11.5	$F_{3,12} = 5.95$	8.81						
Accountin Method	ng 1.52	$F_{4,12} = 3.25$	not significant						

Table 17. Adjusted R², analysis of variance for the 'Automobile' industry.

Effect	Estimated F statistic	Critical value of F	Level of significance
Firm	6.22	F _{3,12} = 5.95	0.01
Accountin Method	g 1.21	$F_{4,12} = 3.25$	not significant

Table 18. Adjusted R², analysis of variance for the 'Iron and Steel' and the 'Automobile' industries.

Effect	Estimated F statistic	Critical value of F	Level of significance
Firm	23.97	$F_{7,28} = 3.36$	8.01
Accountin Method	g 0.54	F _{4,28} = 4.07	not significant

Table 19. Adjusted R², firm rankings, and Kendall's coefficient of concordance for the combined sample of the 'Iron and Steel and 'Automobile' industries.

Combination	Firms								
of Techniques	ARMCO	USX	LTV	LUKENS	GM	FORD	AMC	CHRYSLER	
				Adusted	R ²				
LIFOSLHC	<u>8.69</u>	0.88	0. 67	0.32	8.98	8.94	8.72	8.87	
LIFODDBSL	8.69	8.87	8.73	8.31	8.98	0.95	8.72	0.89	
FIFOSLHC	0.88	8.71	8.82	8.63	8.92	8.82	8.73	8.87	
FIFODDBHC	8.79	8.74	8.85	8.66	8.91	8.81	0.73	<u>0.87</u>	
RC	0.75	0.75	0.66	0.36	8.93	0.88	8.92	8.89	
				Rank	5				
LIFOSLHC	6	4	8	7	1	2.5	5	2.5	
LIFODDBHC	7	4	8	5	1	2	6	3	
FIFOSLHC	5	7	8	3.5	1	3.5	6	2	
FIFODDBHC	5	6	8	3	1	4	7	2	
RC	5.5	5.5	7	8	1	4	2	3	
Kendall's c Value of S Value of	oeffic = 736.9 = 24.53	ient o 5 *	f conc	ordance	W = (.7865			
* significa	nt at (9.91							

Table 28. For the comparison LIFO vs.FIFO:SLHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

	Firms							
Combination of Techniques	Arnco	USX	Lukens	LTV	GM	FM	λМ	Chrysler
			Adjus	ited R ²	1			
LIFOSLHC FIFOSLHC	0.69 0.80	8.88 8.7 1	0.32 0.63	8.67 8.82	0.98 0.92	0.94 0.82	8.72 8.73	0.88 0.87
				Ranks	6			
LIFOSLHC FIFOSLHC	5 5	3.5 7	7 8	6 3.5	1	2 3.5	4 6	3.5
Spearman's : * significa	Rank Co nt at 6	rrela .05	tion Coe	fficie	nt r _s	= 0.7	' 0 5×	

Table 21. For the comparison LIFO vs.FIFO:DDBHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination	r 1 CRS							
of Techniques	Armco	USX	Lukens	LTV	GM	FM	АМ	Chrysler
			λ	djuste	rd R ²			
L I FODDBHC F I FODDBHC	8. 67 8.79	0.87 0.74	0.31 0.66	0.73 0.85	0.98 0.91	0.95 0.81	0. 72 0.73	0.88 0.87
				Ra	inks			
LIFODDBHC FIFODDBHC	7 5	4 6	8 8	5 3	1 1	2 4	6 7	3 2
Spearman's * significa	Rank Co nt at Ø	rrela .05	tion Coe	fficie	nt r s	= 0.7	/86×	

Firms

Table 22. For the comparison SL vs. DDB:LIFOHC, Adjusted R^2 , firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

		Firms								
Combination Techniques	Armco	USX	Lukens	LTV	GM	FM	АМ	Chrysler	_of	
				Adjust	ed R ²					
SLLIFOHC DDBLIFOHC	0.69 0.67	8. 88 8.87	0.32 0.31	0.67 8.73	0.98 0.98	0.94 0.95	8.72 8.72	0.88 0.87		
				R	lanks					
SLLIFOHC DDBLIFOHC	5 7	3.5 4	6 8	1 5	2 1	2 2	4 6	3.5		
Spearman's * significa	Rank Co nt at 0	rrela .01	tion Coe	fficie	ent r _s	= 0.8	75×			

Table 23. For the comparison SL vs. DDB:FIFOHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination				Fir	ms				
of Technigues	Armco	USX	Lukens	LTV	GM	FM	АМ	Chrysler	
		Adjusted R ²							
SLFIFOHC	0.80	0.71	0.63	0.82	0.92	8.82	0.73	8.87	
DDBFIFOHC	0.79	0.72	8.66	0.85	0.91	0.81	8.72	0.87	
				F	lanks				
SLFIFOHC	5	7	3.5	1	3.5	1	3.5	2	
DDBFIFOHC	5	6.5	8	3	1	4	6.5	2	
Spearman's 1 * significa	Rank Co nt at 6	orrela .01	tion Coe	fficie	ent r _s	= 0.9)50×		

Table 24. For the comparison RC vs. HC:LIFOSL, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination				Firm	15					
of Techniques	Armco	USX	Lukens	LTV	GM	FN	λМ	Chrysler		
		Adjusted R ²								
RC	8.75	8.76	8.36	8.66	8.93	8.88	0.92	0.89		
HCLIFOSL	0.69	8.88	0.32	8.67	0.98	8.94	0.72	0.87		
				Ran	ks					
RC	6	5	8	7	1	4	2	3		
HCLIFOSL	6	3	8	7	1	2	5	4		
Spearman's : *significan	Rank Co t at 0.	errela 85	tion Coe	ffici	ent r _s	= 0.7	86*			

Table 25. For the comparison RC vs. HC:FIFOSL, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination								
of Techniques	Armco	USX	Lukens	LTV	GM	FM	АМ	Chrysler
		Adjusted R ²						
RC	8.75	8.76	0.36	8.66	0.93	0.88	8.92	8.89
HCFIFOSL	0.80	8.71	0.63	8.82	0.92	8.82	0.73	8.87
				Ran	n k s			
RC	6	5	8	7	1	4	2	3
HCFIFOSL	5	7	8	3.5	1	3.5	6	2
Spearman's	Rank Co	orrela	tion Coe	fficie	ent r	= 0.5	90	

Firms

Table 26. For the comparison RC vs. HCO:LIFODDB, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination		Firms						
of Techniques	Armco	USX	Lukens	LTV	GM	FM	лм	Chrysler
		Adjusted R ²						
RC	0.75	8.76	8.36	8.66	0.93	8.88	0.92	8.89
HCL I FODDB	0.67	8.87	0.31	8.73	0.91	8.81	0.73	0.87
				Rar	nks			
RC	6	5	8	7	1	4	2	3
HCLIFODDB	6	2.5	7	5.5	1	4	5.5	2
Spearman's 1 * significa	Rank Co nt at 8	rrela .05	tion Coe	fficie	ent r _s	= 0.7	'35×	

Table 27. For the comparison RC vs. HC 'FIFODDB, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination of								
Techniques	Armco	USX	Lukens	LTV	GM	FM	λМ	Chrysler
	Adjusted R ²							
RC	0.75	0.76	0.36	0.66	0.93	0.88	0.92	8.89
hcfifoddb	0.79	0.74	0.66	0.85	0.91	0.81	0.73	6.87
					Ranks			
RC	6	5	8	7	1	4	2	3
HCFIFODDB	5	6	8	3	1	4	7	2

.

Firms

Comparison Spearman's Rank Correlation Coefficient r LIFO vs. FIFO: SLHC 8.785* LIFO vs. FIFO: DDBHC 8.786* SL vs. DDB : LIFOHC 8.862* SL vs. DDB : FIFOHC 8.958* RC vs. HCLIFOSL 8.786× RC vs. HCFIFOSL 0.598 RC vs. HCLIFODDB 8.735* RC vs. HCFIFODDB 8.476 * significant at the 0.05 level ** significant at the 8.61 level

Table 28. Adjjusted R², Spearman's rank correlation coefficient for the combined sample of the 'Iron and Steel' and 'Automobile' industries.

TABLE 38. For the comparison LIFO vs. FIFO:SLHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry

Combination		Firms					
of Techniques	Armco	USX	Lukens	LTV			
		Adju	sted R ²				
LIFOSLHC FIFOSLHC	0.69 0.80	8.88 8.71	0.32 0.63	0.67 0.82			
			Ranks				
LIFOSLHC 2	2 4	1 1	4	3	FIFOSLHC		
Spearman's Ra Critical valu	nk Correl e of r a	ation Co t 0.05 l	efficient r evel of sig	= 0.60 Nificanc	e = 1.00		

Table 31. For the comparisonLIFO vs. FIFO:DDBHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.

Conhination		Fi	rns -		
of Techniques	Arnco	USX	Lukens	LTV	
		Adj us	ted R ²		
LIFODDBHC	8.67	8.87	8.31	0.73	
FIFODDBHC	0.79	8.74	0.66	8.85	
		Ra	nks		
LIFODDBHC 2 3	3 4	1	4	2	F I FODDBHC
Spearman's Ra Critical valu	ank Corre ue of r	lation C at 0.05	oefficient level of si	r = -0. għificar	40 nce = 1.00

TABLE 32. For the comparison SL vs. DDB:LIFOHC,Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination			Firms		
of Techniques	Armco	USX	Lukens	LTV	
		Adj	usted R ²		
SLL I FOHC DDBL I FOHC	0.69 0.67	0.88 0.87	0.32 0.31	0.67 0.73	
		Ra	nks		
SLLIFOHC 3 l	2 4	1 2	4	3	DDBL I FOHC
Spearman's Ra Critical valu	nk Correl e of r a	ation Co t 0.05 l	efficient z evel of sig	= 0.80 Nificanc	e = 1.88

TABLE 33. For the comparison SL vs. DDB:FIFOHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.

.

Combination _		Fir	RS					
of Techniques	Armco	USX	Lukens	LTV				
		Adj u	sted R ²					
SLFIFOHC	0.80	0.71	0.63	8.82				
DDBFIFOHC	0.79	0./4	0.65	8.85				
			Ranks					
SLFIFOHC 2	2 4	3 1	4	1	DDBF I FOHC			
Spearman's Ra * significant	nk Correl at 0.05	ation Co	efficient r	s = 1.00 S	*			

Table 34.	For the comparison RC vs. HCLIFOSL, Adjusted R ² ,
	firm rankings, and Spearman's rank correlation
	coefficient for the Iron and Steel industry.

Combination		Fir	25		
of Techniques	Armco	USX	Lukens	LTV	
		Ajus	ted R ²		
RC HCLIFOSL	0.75 0.69	0.76 0.88	8.36 8.32	0.66 0.67	
		Ra	nks		
RC 2 1	2 4	1 3	4	3	HCLIFOSL
Spearman's R *significant	ank Correl at 0.05	ation Co	efficient r	s = 1.00	*

Table 35. For the comparison RC vs. HCFIFOSL, adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination		F	i ras		
of Techniques	Armco	USX	Lukens	LTV	
		Adju	sted R ²		
rc HCFIFOSL	0.75 0.80	0.76 0.71	0.36 0.63	8.66 8.82	
			Ranks		
RC 2 3	2 4	1 1	4	3	HCF1FOSL
Spearman's R Critical valu	ank Correl ue of r_ a	ation Co t 0.05 1	efficient r evel of sig	= 0.60 Nificanc	e = 1.00

-			Fir			
Combi of Techn	nation	Armco	USX	Lukens	LTV	
			Adjus	ted R ²		
RC HCLIF	ODDB	0.75 0.67	0.76 0.87	0.36 0.31	8. 66 8. 73	
				Ranks		
RC 3	1	2 4	1 2	4	3	HCL I FODDB
Spear Criti	rman's Ran Ical value	nk Correla e of r a	ation Coe t 0.05 le	fficient r vel of sig	= 0.80 Nificanc	r = 1.00

Table 36. For the comparison RC vs. HCLIFODDB, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.

Table 37. For the comparison RC vs. HCFIFODDB, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combinat	ion		Firms			
of Techniques		Armco	USX	Lukens	LTV	
			Adj us	ted R ²		
RC		0.75	0.76	0.36	8.66	
HCLIFOSL		0.79	8.74	0.66	0.85	
			Ra	nks		
RC		2	1	4	3	HCLIFOSL
2	3	4	1			
Sp earma n Critical	's Ra valu	nk Correl le of r_a	ation Co t 0.05 lo	efficient r evel of sig	= 0.60 Nificanc	e = 1.00

Comparison	Spearman's Rank Correlation Coefficient r
LIFO vs. FIFO: SLHC	8.68
LIFO vs. FIFO: DDBHC	-0.40
SL vs.DDB: LIFOHC	0.80
SL vs. DDB: FIFOHC	1.00*
RC vs. HCLIFOSL	1.00*
RC vs. HCFIFOSL	0.60
RC vs. HCLIFODDB	0.80
RC vs. HCFIFODDB	0.60
* significant at 0.05	

Table 38. Adjusted R², Spearman's rank correlation coefficient for the Automobile industry.

Combination						
of Tecnnifques	GM	FM	ANC	Chrysler		
	Adjusted R ²					
LIFOSLHC LIFODDBHC FIFOSLHC FIFODDBHC RC	0.98 0.98 0.92 0.91 0.93	6.94 6.95 6.82 6.81 6.88	0.72 0.72 0.73 0.73 0.92	0.87 0.88 0.87 0.87 0.89		
		Ran	ks			
LIFOSLHC LIFODDBHC FIFOSLHC FIFODDBHC RC	1 1 1 1	2 2 3 4 4	4 4 3 2	3 3 2 2 3		
Kendall's Coef Value of S=83* Critical value * significant	ficient of (of S at 1.) at 0.01	Concordance W 80 level of s	= 0.664 ignificance	= 80.5		

TABLE 39.	Adjusted R ² , firm rankings, and Kendall's
	coefficient of concordance for the Automobile industry.

Combination	Firms							
of Techniques	GM	FM	AMC	Chrysler				
		Adust	ed R ²					
SLLIFOHC DDBLIFOHC	0.98 0.98	0.94 0.95	8. 72 8. 72	0.88 0.88				
	Ranks							
SLLIFOHC DDBLIFOHC	1	2 2	4	3 3				
Spearman's Ram * significant	nk Correlati at 0.05	on Coefficien	nt r _s = 1.00					
Table 43. For fir co	r the compar rm rankings, efficient fo	ison SL vs. I and Spearman or the Automob	DB: FIFOHC, is rank cor oile industr	Adjusted R ² , relation Y.				
Combination		Firms						
of Techniques	GM	FM	АМС	Chrysler				
	Adjusted R ²							
SLF I FOHC DDBF I FOHC	0.92 0.91	8.8 2 8. 81	0.73 0.73	0.87 0.87				
	Ranks							
Slf i Fohc DDBF i Fohc	1	3 3	4	2 2				
Spearman's Ram * significant	nk Correlati at 0.05	on Coefficien	nt r = 1.00 5	*				

Table 42. For the comparison SL vs DDB (LIFOHC, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Automobile industry.

Table 43.	For the comparison RC vs. HCLIFOSL, Adjusted R ² ,
	firm rankings, and Spearman's rank correlation
	coefficient for the Automobile industry.

Combination	Firms					
of Techniques	GM	FM	ANC	Chrysler		
Adjusted R ²				<u> </u>		
RC HCLIFOSL	8.93 8.98	0.88 0.94	0.92 0.72	0.89 0.87		
		Rani	(5			
RC HCLIFOSL	1 1	4 2	2 4	3		

Spearman's Rank Correlation Coefficient r = 0.20 Critical value of r at 0.05 level of significance = 1.00

Table 45. For the comparison RC vs. HC: FIFOSL, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Automobile industry.

Combination	Firms				
of Techniques	GM	FM	АМС	Chrysler	
Adjusted R ²					
rc HCFIFOSL	0.93 0.92	0.88 0.82	0.92 0.73	0.89 0.87	
		Ran	ks		
RC HCFIFOSL	1 1	4 3	2 4	3 2	
Spearman's Ra Critical valu	nk Correlati e of r_ at Ø	on Coefficien .05 level of	nt r = 0.40 significar) 1ce = 1.00	

Table 46.	For the comparison RC vs. HC: LIFODDB, Adjusted R ² ,
	firm rankings, and Spearman's rank correlation
	coefficient for the Automobile industry.

Combination	Firms				
of Techniques	GM	FM	лис	Chrysler	
RC HCLIFODDB	0.93 0.98	8. 88 8.95	0.92 0.72	8.89 8.88	
		Rar	nks		
RC	1	3	2	4	
HCLIFODDB	1	2	4	3	
Spearman's Rar Critical value	nk Correlati e of r_ at 0	on Coefficier .05 level of	nt r = 0.40 significanc	ce = 1.80	

TABLE 47. For the comparison RC vs. HC: FIFODDB, Adjusted R², firm rankings, and Spearman's rank correlation coefficient for the Automobile industry.

Combination	Firms				
of Techniques	GM	FM	умс	Chrysler	
RC HCF I FODDB	0.93 0.98	0.88 0.95	0.92 0.72	0.89 0.88	
		Ran	ks		
RC HCF I FODDB	1 1	3 3	2 4	4 2	
Spearman's Rar Critical value	nk Correlations of r_at 0	on Coefficier .05 level of	nt r = 0.20 Significan	ace = 1.00	

Comparison	Spearman's Rank Correlation Coefficient r
LIFO vs. FIFO: SLHC	8.80
LIFO vs. FIFO: DDBHC	0.80
SL vs.DDB: LIFOHC	1.00*
SL vs. DDB: FIFOHC	1.00*
RC vs. HCLIFOSL	0.20
RC vs. HCFIFOSL	8.40
RC vs. HCLIFODDB	8.46
RC vs. HCFIFODDB	0.20
* Significant at 0.05	

Table	48.	Adjusted R ²	, Spei	arman's	rank c	orrelation
		coefficient	for t	the auto	mobile	industry

Table 49. Test Criterion, firm rankings, and Kendall's coefficient of concordance for the 'Iron and Steel' and 'Automobile' industries.

Combination of Techniaues Armco USX Lukens LTV GM FM AM Chrysler Test Criterion 0.442 0.251 0.997 0.993 0.828 0.679 LIFOSLHC 0.592 0.892 LIFODDBHC 0.596 0.838 0.435 0.326 0.984 0.956 0.834 0.653 FIFOSLHC 0.936 0.771 0.840 1.058 1.060 1.128 0.717 0.534 FIFODDBHC 0.936 0.777 0.840 1.220 1.064 1.298 0.715 0.526 RC 0.514 0.784 0.428 0.443 1.027 0.888 0.907 0.825 Ranks LIFOSLHC 3 6 7 8 1 2 5 4 LIFODDBHC 6 3 7 8 5 1 2 4 FIFOSLHC 6 5 3 2 7 8 4 1 FIFODDBHC 6 5 2 3 7 8 4 1 7 6 5 8 1 3 2 4 RC Kendall's Coefficient of Concordance W = 0.507 Value of S=532 Value of = 17.73 * Critical value of at 0.02 level of significance = 16.62 * significant at $\alpha = 0.02$

Firms

Table 58. For the comparison LIFO vs. FIFO: SLHC, Test Criterion, firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination of Techniques	Armco	USX	Lukens	LTV	GM	FM	AM Ch	rysler		
		Test Criterion								
LIFOSLHC	8.592	8.829	8.442	8.251	0.997	0.993	0.828	8.679		
FIFOSLHC	0.936	0.771	0.840	1.058	1.060	1.298	0.717	0.534		
				Rank	S					
LIFOSLHC	6	3	7	8	1	2	4	5		
FIFOSLHC	4	6	5	3	2	1	7	8		
Spearman's I Critical vai	Rank Com lue of m	rrelat r_ at (ion Coe: 8.05 le	fficien vel of	nt r = signif	9.262 icance	= 0.6	43		

Table 51. For the comparison LIFO vs. FIFO: DDBHC, Test Criterion, firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile'industries.

	Firms										
Combination of Techniques	Armco	USX	Lukens	LTV	GM	FM	AM Ch	rysler			
		Test Criterion									
L I FODDBHC F I FODDBHC	0.596 0.936	0. 838 0.777	0.435 0.840	0.326 1.220	0.984 1.064	0 .956 1.298	6.834 0.715	0.653 0.526			
				Rank	5						
LIFODDBHC FIFODDBHC	6 4	3 6	7 5	8 2	1 3	2 1	4 7	5 8			
Spearman's I Critical va	Rank Com lue of m	rrelat r_at	ion Coef 0.05 lev	ficien el of s	t r = signif	0.095 icance	; 2 = 0.6	43			

Firms

Table 52. For the comparison SL vs. DDB: LIFOHC, Test Criterion, firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

	• • • • • • • • • • • • • • • • • • •									
Combination of Techniques	Armco	USX	Lukens	LTV	GM	FM	ам с	Chrysler		
			T	est Cr	iterio	n				
SLLIFOHC DDBLIFOHC	0.592 0.596	0.829 0.838	0.442 0.435	0.251 0.326	0.997 0.984	0.993 0.956	0.828 0.834	8 0.679 1 0.653		
				Rani	ks					
SLLIFOHC DDBLIFOHC	6 6	3 3	7 7	8 8	1	2 2	4	5 5		
Spearman's 1 * significan	Rank Co: nt at 0	rrelat .01	ion Coe	fficie	nt r s	= 1.00	8 *			

Table 53. For the comparison SL vs. DDB: FIFOHC, Test Criterion, firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

	Firms									
Combination of Techniques	Armco	USX	Lukens	LTV	GM	FM	λМ	Chrysler		
			T	est Cr.	iterio	n				
Slfifohc DDBFifohc	0.936 0.936	0.771 0.777	0.840 0.840	1.058 1.220	1.060 1.064	1.128 1.298	8. 71 8. 71	7 0.534 5 0.526		
				Ra	nks					
SLFIFOHC DDBFIFOHC	4	6 6	5 5	3 3	2 2	1 1	7 7	8 8		
Spearman's 1 * significa	Rank Com nt at 0	rrelat .01	ion Coe:	fficie	nt r _s	= 1.00	9 *			

Rieme

Table 54. For the comparison RC vs. HC: LIFOSL, Test Criterion, firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Combination of Techniques	λrmco	USX	Lukens	LTV	GM	FM	AM (Chrysler
			T	est Cr	iterio	n		
RC HCLIFOSL	0.514 0.392	0. 784 0.829	0.428 0.442	8.443 8.251	1.027 0.997	0. 888 0.993	0.90 0.82	7 0.825 8 0.679
				Ra	nks			
RC HCLIFOSL	6 7	5 3	8 6	7 8	1 1	3 2	2 4	4 5
Spearman's F * significar 	Rank Co: nt at 8	rrelat .01	ion Coe	fficie	nt r _s	= 0.81	0 ×	
Table 55. Fo fi co ir	or the (irm ran) oeffici(ndustri(compar kings, ent fo es.	ison RC and Sp r 'Iron	vs. H earman and S	C:FIFO 's ran teel'	SL, Te k corr and 'A	st Cr elatio utomol	iterion, on bile'
				Fir	ns			
Combination of								
Techniques	Armco	USX	Lukens	LTV	GM	FM	AM (Chrysler
			T	est Cr	iterio	n		
RC HCFIFOSL	0.514 0.999	0.784 0.819	0.428 0.864	0.443	1.027	8.88 1.128	0. 90 0.71	7 0.825 7 0.534

Firms

Combination of Techniques	Armco	USX	Lukens	LTV	GM	FM	ум с	Chrysler
			T	est Cr:	iterio	n		
RC	0.514	0.784	0.428	0.443	1.027	8.88 8	0.907	0.825
HCFIFOSL	0.999	0.819	0.864	1.069	1.060	1.128	0.717	0.534
				Rani	ks			
RC	6	5	8	7	1	3	2	4
HCFIFOSL	4	6	5	2	3	1	7	8
Spearman's	Rank Co	rrelat	ion Coe:	fficien	nt r	= -0.0/	48	643
Critical va	lue of	r_at	0.05 le	vel of	signi:	ficanc	e = 0.	

	Firms									
Combination of Techniques	Armco	USX	Lukens	LTV	GM	ги	ам с	hrysler		
		Test Criterion								
RC HCLIFODDB	0.514 0.596	8.784 6.838	8.428 8.435	8.44 3 8.326	1.027 0.984	8.888 8.956	0.907 0.834	8. 825 8. 653		
				Ran	ks					
RC HCLIFODDB	6 6	5 3	8 7	7 8	1	3 2	2 4	4 5		
Spearman's : * Significa	Rank Co: nt at 0	rrelat .01	ion Coe	fficie	nt r _s	= 0.8	452*			

Table

Table 57. For the comparison RC vs. HC: FIFODDB, Test Criterion, firm rankings, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

	Firms										
Combination of Techniques	Armco	USX	Lukens	LTV	GM	FM	АМ	Chrysler			
		Test Criterion									
RC HCFIFODDB	0.514 0.936	0. 784 0. 777	0.428 0.840	0.44 3 1.220	1.027 1.064	0.888 1.298	0.9 0.7	07 0.825 15 0.526			
				Rani	ks						
RC HCF I FODDB	6 4	5 6	8 5	7 2	1 3	3 1	2 7	4 8			
Spearman's Critical va	Rank Co lue of	rrelat r_at	ion Coe 0.05 le	fficie vel of	nt r signi:	= -0.0 ficanc	48 e =	0.643			

Table 58. Market power test Criterion, and Spearman's rank correlation coefficient for 'Iron and Steel' and 'Automobile' industries.

Comparison	Spearman's Rank Correlation Coefficient r
LIFO vs. FIFO: SLHC LIFO vs. FIFO: DDBHC SL vs.DDB: LIFOHC SL vs. DDB: FIFOHC RC vs. HCLIFOSL RC vs. HCFIFOSL RC vs. HCLIFODDB RC vs. HCFIFODDB	0.262 0.095 1.00** 1.00** 0.810* -0.048 0.845* -0.048
* significant at .05 ** significant at .01	

Combination	Firms								
of Techniques	Armco	USX	Lukens	LTV					
		Test C	riterion						
LIFOSLHC LIFODDBHC FIFOSLHC FIFODDBHC RC	0.592 0.596 0.936 0.936 0.514	0.829 0.838 0.771 0.777 0.784	0.442 0.435 0.840 0.840 0.428	0.251 0.326 1.058 1.220 0.443					
		RANKS							
LIFOSLHC LIFODDBHC FIFOSLHC FIFODDBHC RC	2 2 2 2 2 2	1 1 4 4 1	3 3 3 3 4	4 4 1 3					
Kendall's Coef Value of S=21 Critical value	ficient of C of S at 0.0	concordance 5 level of	W = 0.168 significance	= 62.6					

Table 59. Test criterion, firm rankings, and Kendall's coefficient of concordance for the iron and steel industry

Table 68. For the comparison LIFO VS.FIFO: SLHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination		Fir	ms	
of Techniques	Armco	USX	Lukens	LTV
		Test C	riterion	
LIFOSLHC	0.592	0.829	0.442	8.251
FIFOSLHC	0.936	0.771	8.848	1.058
		Ran	ks	
LIFOSLHC	2	1	3	4
FIFOSLHC	2	4	3	1
Spearman's Ran Critical value	k Correlatic of r_at0. s	on Coefficie 85 level d	ent r = -0.8 of significan	0 10e = 1.00

Table 61. For the comparison LIFO VS.FIFO:DDBHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination		Fir	ns	
of Techniques	Armco	USX	Lukens	LTV
		Test C	Criterion	
LIFODDBHC FIFODDBHC	0.596 0.936	8. 838 8.777	0.435 0.840	0.326 1.220
		Ran	nks	
LIFODDBHC FIFODDBHC	2 2	1 4	3 3	4 1
Spearman's Ran Critical value	c Correlation of r at 0.	on Coefficie 05 level c	ent r = -0.8 of significan	0 Ace = 1.00

86

Table 62. For the comparison SL VS. DDB:LIFOHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination		Fir	`RS	
of Techniques	Armco	USX	Lukens	LTV
		Test C	riterion	
SLLIFOHC	0.592	8.829	0.442	8.251
DDBLIFOHC	0.596	0.838	8.435	0.326
		Ra	inks	
SLL I FOHC	2	1	3	4
DDBLIFOHC	2	1	3	4
Spearman's Rank * significant a	k Correlatic at .05	on Coefficie	ent r = 1.00 s)*

Table 63. For the comparison SL VS. DDB:FIFOHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination of Techniques		Fir	ns		
	Armco	USX	Lukens	LTV	
		Test C	riterion		
SLFIFOHC	0.936	8.771	8.848	1.058	
DDBFIFOHC	0.936	0.777	0.840	1.220	
		Ran	iks		
SLFIFOHC	2	4	3	1	
DDBFIFOHC	2	4	3	1	
Spearman's Ran * significant a	k Correlatic at 0.05	on Coefficie	ent r = 1.00 S	*	

Combination		Fir	RS	
of Techniques	Armco	USX	Lukens	LTV
		Test C	riterion	
RC	0.514	8.784	0.428	8.443
HCLIFOSL	0.592	0.829	0.442	0.251
		Ran	nks	
RC	2	1	4	3
HCLIFOSL	2	1	3	4
Spearman's Ran Critical value	k Correlatic of r_ at 0.	on Coefficie 85 level d	nt r = 0.80 of significar	nce = 1.00

Table 64. For the comparison RC VS. HCLIFOSL, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Table 65. For the comparison RC vs. HCFIFOSL, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination		Fir	'As	
of Techniques	Armco	USX	Lukens	LTV
		Test C	riterion	
RC	0.514	0.784	0.428	0.443
HCFIFUSL	0.336	U.//I Par	0.040	1.050
		A BI	IK 3	
RC	2	1	4	3
HCFIFOSL	2	4	3	1
Spearman's Ran Critical value	k Correlatic of r_ at 0.	on Coefficie 05 level d	ent r = -0.4 of significan	10 nce = 1.00

Combination		Fir	ns	
of Techniques	Armco	USX	Lukens	LTV
		Test C	riterion	
RC	8.514 8.595	8. 784	8.4 28 8.435	8.443 8.326
	0.550	Rank	5	01320
RC	2	1	4	3
HCLIFODDB	2	1	3	4
Spearman's Ran Critical value	k Correlatic of r at 0. s	n Coefficie 85 level d	ent r = 0.80 of significar	nce = 1.00

Table 66. For the comparison RC vs. HCLIFODDB, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Table 67. For the comparison RC vs. HCFIFODDB, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Iron and Steel industry.

Combination		Fir	RS	
Combination of Techniques	Armco	USX	Lukens	LTV
		Test C	riterion	
RC	0.514	0.784	0.428	8.443
HCFIFODDB	0.936	0.777	0.840	1.220
		Ran	iks	
RC	2	1	4	3
HCF I FODDB	2	4	3	1
Spearman's Ran Critical value	k Correlatic of r _s at 0.	on Coefficie 05 level c	ent r = -0.4 of significan	10 1ce = 1.00

Table 68.	Market power to	est criterion,	firm rankings	for	the
	Iron and Steel	industry.			UNE

Comparison	Spearman's Rank Correlation Coefficient r _s
LIFO vs. FIFO: SLHC	-0.80
LIFO vs. FIFO: DDBHC	-0.80
SL vs.DDB: LIFOHC	1.00*
SL vs. DDB: FIFOHC	1.00*
RC vs. HCLIFOSL	0.80
RC vs. HCFIFOSL	-0.40
RC vs. HCLIFODDB	0.80
RC vs. HCFIFODDB	-8.40
* significant at 0.05	

Combination	 	Pirms			
of Techniques	GM	FM	АМС	Chrysler	
		T es t Cr	iterion		
LIFOSLHC LIFODDBHC FIFOSLHC FIFODDBHC RC	0.997 0.984 1.060 1.064 1.827	0.993 0.956 1.128 1.298 0.888	0.828 0.834 0.717 0.715 0.907	0.679 0.653 0.534 0.526 0.825	
		Ran	ks		
LIFOSLHC LIFODDBHC FIFOSLHC FIFODDBHC RC	1 1 2 2 1	2 2 1 1 3	3 3 3 3 2	4 4 4 4	
Kendall's Coe Value of S=10 Critical valu Critical valu * significant	efficient o) * le of S at le of S at : at .01	f Concordan 0.05 level 0.01 level	ce W = 0.80 of significa of significa	nce = 62.6 nce = 80.5	

Table 69.	Market power Test criterion,	firm rankings
	and Kendall's coefficient of	concordancefor the
	Automobile industry.	

Table 68. For the comparison LIFO vs.FIFO:SLHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

Combination	Firms			
of Techniques	GM	FM	АМС	Chrysler
		Test Cr	iterion	
LIFOSLHC FIFOSLHC	0.997 1.060	8.993 1.298	0.828 0.717	0.679 0.534
		RANK	<u>(5</u>	
LIFOSLHC FIFOSLHC	1 2	2 1	3 3	4
Table 71. For	the compari terion, firm	son LIFO ve rankings a	of signific 	BHC, Test n's rank mobile industry.
Combination	Firms			
of Techniques	GM	FM	лис	Chrysler
		Test C	riterion	
LIFODDBHC FIFODDBHC	0.984 1.064	8. 956 1.298	0.834 0.715	0. 653 0.526
		Rar	nks	
LIFODDBHC FIFODDBHC	1 2	2 1	3 3	4
Spearman's Ran Critical value	nk Correlati e of r _e at 6	ion Coeffici 1.85 level d	ient r = 0 of signific	.80 ance = 1.00

Table 72. For the comparison SL vs. DDB: LIFOHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

Combination of Techniques	Firms						
	GM	FM	АМС	Chrysler	•		
	Test Criterion						
SLLIFOHC	8.997	0.993	0.828	8.679			
DDBLIFOHC	8.984	0.956	0.834	8.653			
		Ran	nks				
SLLIFOHC	1	2	3	4			
DDBLIFOHC	1	2	3	4			
Spearman's Ran * Significant	k Correlati at 0.05	on Coeffici	ent r = 1	.00*			

Table 73. For the comparison SL vs. DDB:FIFOHC, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

Combination of Techniques	Firms				
	GM	FM	AMC	Chrysler	
		Test Cr	iterion		
SLFIFOHC	1.060	1.128	0.717	0.534	
DDBFIFOHC	1.064	1.298	0.715	0.526	
		Rank	5		
SLFIFOHC	2	1	3	4	
DDBFIFOHC	2	1	3	4	
Spearman's Rar * significant	nk Correlati at 0.05	on Coeffici	ent r = 1 s	.00 *	

Combination	Firms						
of Techniques	GM	FM	AMC	Chrysler			
	Test Criterion						
RC HCLIFOSL	1.027 0.997	0. 888 0.993	0.907 0.828	0.825 0.627			
		Rank	.5				
RC HCLIFOSL	1	3 2	2 3	4			
Spearman's Ran Critical value	k Correlati of r_at 0	on Coeffici .85 level c	ent r = 0 of signific	.80 ance = 1.00			
Table 75. For fir coe	the compar m rankings efficient fo	ison RC vs. and Spearma or the Autom Fi	HCFIFOSL, n's rank c obile indu rms	Test criterion, orrelation stry.			
of Techniques	GM	FM	АМС	Chrysler			
		Test Cr	iterion	4.4			
RC HCFIFOSL	1.027 1.060	0.888 1.128	0.907 0.717	0.825 0.534			
		Ran	iks				
RC HCLIFOSL	1	3 2	2 3	4			
Spearman's Rar Critical value	nk Correlati e of r_at 0 s	ion Coeffici 1.85 level c	ent r = 8 of signific	.40 ance = 1.00			

Table 74. For the comparison RC vs. HCLIFOSL, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.
	Firms		
sler	GM	u es G	of Techniques
825	.027	1.0	RC
53			ACLIF ODDB
	1	1 DB 1	RC HCLIFODDB
1	1 1 Correlation (DB 1 n's Rank Cor	RC HCLIFODDB Spearman's Critical y

Table 76. For the comparison RC vs. HCLIFODDB, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

Table 77. For the comparison RC vs. HCFIFODDB, Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

Combination of Techniques	Firms			
	GM	FM	AMC	Chrysler
		Test C	riterion	
RC	1.027	0.888	0.907	0.825
HCF I FODDB	1.064	1.298	0.715	0.526
		Ran	ks	
RC	1	3	2	4
HCF I FODDB	2	1	3	4
Spearman's Ran Critical value	k Correlat of r at (ion Coeffic 8.05 level (ient r = 0 of signific:	.40 ance = 1.00

Table 78. Market Power Test criterion, firm rankings and Spearman's rank correlation coefficient for the Automobile industry.

Comparison	Spearman's Rank Correlation Coefficient r
LIFO vs. FIFO: SLHC	0.80
LIFO vs. FIFO: DDBHC	8.88
SL vs.DDB: LIFOHC	1.00*
SL vs. DDB: FIFOHC	1.00*
RC vs. HCLIFOSL	0.80
RC vs. HCFIFOSL	0.40
RC vs. HCLIFODDB	0.80
RC vs. HCFIFODDB	0.40
* significant at .01	

VII CONCLUSIONS

As is analyzed in the previous section, accounting methods do not seem to have a significant effect on the relevant parameters of estimation, when analyzed jointly for the complete sample of firms (results from the analysis of variance and the Kendall's coefficient of concordance tests). From the more disaggregate pairwise analysis with Spearman's rank correlation coefficient, however, it is very clear that changes in certain accounting techniques have a very significant effect on the ranking of firms both by market power as well as by the goodness of fit of the model. In the context of the model developed in the study, the techniques with the most effects are the changes in inventory measurement techniques i.e. from LIFO to FIFO. These results are contrary to the results obtained by Nair [1979]] in that his study did not show significant effects due to accounting techniques.

The results of this study emphasize to the importance of paying attention to the uderlying accounting data when intercomparisons of a test criterion or goodness of fit are made. Since in no case did the same accounting technique give the best rank for all firms, it points to the fact that intercomparisons of results when different techniques are used by the firms can be very misleading.

97

This study also points out an important fact that was overlooked in the earlier study by Hagerman and Senbet[1975] in that it is not sufficient to look at aggreagated crossectional results when the research question is at the disaggregate firm level. So their research finding that on averagee most firms in an industry use the same accounting methods is not relevant when examining firms in industries where different accounting methods are used. In that case intercomparisons of estimated data without regard to the underlyling techniques can lead to misleading analyses. However, the question as to which techniques need to be paid attention to depends on the model being tested and the particular firms involved in the intercomparison. Since the model that was developed in this study dealt with the dependent as well as the independent variables involving measurement of inventories, it is not surprising that in most cases the change of LIFO to FIFO and vice versa had the most profound effect.

The effects of techniques on the goodness of fit of the estimated model has implications when making intermodel comparisons as Nair did, or when testing the validity of a particular economic model. As this study shows, the explanatory power of the model does depend on the particular accounting techniques underlying the data.

In this study, an attempt has been made to examine two questions. Referring to the diagramme on page 4, the two

98

questions of interest are, (1) whether the measurement model B that is based on accounting measurements captures the same phenomenon as the unobservable model λ , and (2) whether the accounting model is sensitive to alternative accounting methods underlying the dependent and the independent variables.

The normative answer to the first question is that if the economic phenomenon is to be measured by accounting numbers, it is important to model that phenomenon of interest in terms of the definitions that accountants use. For example inventories should be modeled and not assumed nonexistant if accounting measurement of the Sales Revenue and the Cost of Goods sold is involved.

The empirical answer to the second question is that at the micro level of an individual firm where the differences due to accounting techniques are not aggregated out as in the case of industry level analysis, it is very important to be aware of the accounting techniqes that might impact on the dependent and the explanatory variables.

Finally this study has attempted to analyze the effects of accounting techniques from the perspective of impact on decision rules in the case where accounting data is actually used. In the light of the results from this

99

study it would be worthwhile to study and develope other models that are more suited for measurement with accounting variables.

VIII APPENDIX

The decision problem is to maximize eqn. (7) in the text, i.e.,

$$\begin{aligned} & \text{Max} \qquad P(S_{t+1}, \overline{Q})\overline{Q} - WL - C^{F}(S_{t+1}) - C^{R}(Sr_{t} + R_{t}) \\ & (L, I, R_{t}, S_{t+1}, Q)^{T} \\ & - \mu_{\gamma} P_{\theta}^{K}(R_{E} + I) - \mu_{\Delta} b_{\theta}^{R}(Sr_{t} + \overline{R}_{t}) \\ & + (1/1 + r^{*}) P(S_{t+1}, \overline{Q})S_{t+1} \\ & \text{s.t.} S_{t} + Q(R_{E} + I, L, Sr_{t} + R_{t}) = \overline{Q} + S_{t+1} \end{aligned}$$

Where, $Q(K_{E} + I, L, Sr_{t} + R_{t}) = Quantity of finished goods$ manufactured. $Forming the Lagrangian <math>L_{\star}$, Max $L^{\star} =$

First order conditions are,

 \overline{Q} : $\delta P \overline{Q} / \delta \overline{Q} = \mu = 8$ i.e. Marginal Revenue = μ = Marginal costtherefore, μ = MC

.

$$\begin{split} \mathbf{S}_{t+1} &: \quad (\delta P/\delta \mathbf{S}_{t+1})(\overline{\mathbf{Q}} + (1/1+r*)\mathbf{S}t+1) - \mathbf{C}^{\mathbf{F}} - \mu = \mathbf{B} \\ \text{or,} \quad (\delta P/\delta \mathbf{S}_{t+1})(\overline{\mathbf{Q}} + (1/1+r*)\mathbf{S}t+1) = \mathbf{C}^{\mathbf{F}} + \mu \qquad (\mathbf{A}.2) \\ \text{Here} \quad \overline{\mathbf{Q}} \delta P/\delta \mathbf{S}_{t+1} \text{ is the marginal revenue due to holding an} \\ \text{extra unit of finished goods inventory.} \end{split}$$

I:
$$\mu_{\gamma} P_{0}^{K} / MP_{K} - \mu \delta Q / \delta I = 0$$

or,
$$\mu_{\gamma} P_{\theta}^{K} / MP_{K} = \mu$$
 (A.3)
where MP_K = $\delta Q / \delta I$ is the marginal product of capital other

$$\vec{R}_{t}: - C^{R} - \mu_{\Delta} b_{B}^{R} + \mu \delta Q / \delta R = 0$$
or,
$$(C^{R} + \mu_{\Delta} b_{B}^{R}) / M P_{R} = \mu$$
(A.4)
where $M P_{R} = \delta Q / \delta R$ is the marginal product of raw
materials.

L:
$$-W + \mu \delta Q/\delta L = 0$$

or, $W/MP_L = \mu = MC$ (A.5)

Where MP_L = $\delta Q / \delta L$ is the marginal product of labor.

Therefore, from the first order conditions (λ .1), (λ .3), (λ .4), and (λ .5),

$$\mu = MC = \mu_{\gamma} P_{\theta}^{K} / MP_{K} = (C_{R} + \mu_{\Delta} b_{\theta}^{R}) / MP_{R} = W / MP_{L}$$
 (A.6)

It is important to note here that in the model incorporating market power, the first order condition with respect to S_{t+1} , the finished goods ending inventory, is not used. This is similar to the treatment of advertising expense in a model proposed by Savyer $[1983]^{A1}$ in which the advertising is not a part of the final model of market power. Another way of thinking about this is that the marginal cost of holding finished goods inventory is a selling cost rather than a cost of production, and therefore is not a part of the factors of production, AC in eqn. (3).

Multiplying equations (A.3), (A.4), and (A.5) with $K = K_E + I$, L, and $R = R_t + Sr_t$ (raw materials available), respectively, and adding gives,

 $(KMP_{K} + RMP_{R} + LMP_{L})MC = \mu_{\gamma}P_{0}^{K}K + \mu_{\Delta}b_{0}^{R}R + C^{R}R + WL$

or,
$$MC = \frac{\mu_{\gamma} P_{\theta}^{R} K + \mu_{\Delta} b_{\theta}^{R} R + C^{R} R + WL}{Q} \frac{Q}{(KMP_{K} + RMP_{R} + LMP_{L})}$$

or^{2A}, MC =
$$\frac{\mu_{\gamma} P_{\theta}^{K} K + \mu_{\Delta} b_{\theta}^{R} R + C^{R} R + WL}{Q} \frac{1}{FC}$$

Substituting for MC in eqn. (2), we get,

$$P = \frac{\mu_{\gamma} P_{0}^{R} R + \mu_{\Delta} b_{0}^{R} R + C^{R} R + WL}{Q} = \frac{1}{FC}$$

$$P = \frac{1}{\epsilon_{QP}}$$

The above equation on simplification gives,

$$PQ = \frac{PQ}{\epsilon_{QP}} + \frac{\mu_{\gamma} P_{\theta}^{R} K + \mu_{\Delta} b_{\theta}^{R} R + C^{R} R + WL}{FC}$$

Multiplying both sides of the equation by FC/PQ and adding 1 gives,

$$1 + FC = 1 + \frac{FC}{\epsilon_{QP}} + \frac{\mu_{\gamma} P_{\theta}^{R} K + \mu_{\Delta} b_{\theta}^{R} R + C^{R} R + WL}{PQ}$$

or, $1 = 1 + FC[\frac{1}{\epsilon_{QP}} - 1] + \frac{\mu_{\gamma} P_{\theta}^{R} K + \mu_{\Delta} b_{\theta}^{R} R + C^{R} R + WL}{PQ}$
Multiplying throughout by $1 = PQ/PQ$ and subtracting from both sides (WL + $C^{R}R + bR_{t}$), the cost of goods
manufactured,

$$\frac{PQ - WL - C^{R}R - bR_{t}}{PQ} = 1 + FC[\frac{1}{\epsilon_{QP}} - 1] + \frac{\mu \gamma P_{\theta}^{K}K}{PQ} + \frac{\mu_{\Delta} b_{\theta}^{R}R}{PQ} - \frac{bR_{t}}{PQ}$$
(A.7)

Here R_t is the quantity, and b is the price of raw materials actually used in production. It is important to note that if the price of raw materials R actually used b $\approx b_0^R$, the actual situation will approximate a FIFO inventory measurement technique, and likewise b $\approx b_1^R$ will approximate LIFO. Examining the FIFO case^{A3} by sustituting b for b_8^R , eqn. (A.7) is equivalent to,

$$\frac{PQ - AC \cdot Q}{PQ} = 1 + FC[\frac{1}{\epsilon_{QP}} - 1] + \frac{\mu \gamma P_{g}^{A} K}{PQ} + \frac{\mu_{a} bR}{PQ} - \frac{bR_{t}}{PQ} \quad (A.8)$$

In order to extend this case of a single seller to a
differentiated product, oligopoly setting,

 $[\epsilon_i \alpha + S_i(1 - \alpha)]/\epsilon_i$ is substituted for $1/\epsilon_{pg}$ ^{A4} in the above eqn., which gives,

$$\frac{PQ - WL - C^{R}R - bR_{t}}{PQ} = (1 - FC + \frac{\alpha}{\epsilon_{i}}) + \frac{(1 - \alpha)FCS_{i}}{\epsilon_{i}} + \frac{K}{\kappa}$$

$$\frac{\mu_{\gamma}P_{\theta}K}{PQ} + \frac{\mu_{a}bR}{PQ} - \frac{bR_{t}}{PQ} (A.9)$$

Here I make a simplifying assumption that the total raw materials available are proportional to total materials used in production, i.e. $R = hR_t$ where h is a constant. This implies that the raw materials ending inventory $Sr_{t+1} = (h - 1)R_t$, or that the raw materials ending inventory is kept at some constant multiple of the optimal quantity of raw materials available.

While this may not be the exact scenario for every firm, it is a plausible story for a feasible inventory policy. Substituting hR_t for R in eqn. (A.8) and combining the last two terms on the right hand side,

$$\frac{PQ - WL - C^{R}R - bR_{t}}{PQ} = (1 - FC + \frac{\alpha}{\epsilon_{i}}) + \frac{(1 - \alpha)FC}{\epsilon_{i}}S_{i}$$
$$+ \frac{\mu\gamma P_{\theta}^{K}R}{PQ} + (h\mu_{\Delta} - 1)\frac{bR_{t}}{PQ} \quad (A.10)$$

The left hand side in the above equation is equal to gross profit if all of the goods manufactured are sold. In order to measure accounting gross profit which in terms of the quantity of goods actually sold instead of quantity of goods manufactured, both sides of the equation are mutiplied by, $(\overline{9}/9)/(\overline{9}/9)$.

Now the left hand side of eqn. $(\lambda.18)$ becomes,

and the right hand side becomes,

$$(1 - FC + \underline{\alpha}) + (\underline{1 - \alpha})FC Si + \frac{\mu \gamma P_{g}^{K} K}{\underline{0}} + \frac{\overline{0}}{\underline{0}} + \frac{bR_{t}}{\underline{0}} = 1]$$

To evaluate \overline{Q}/Q , $P_{g}^{K} \overline{Q}/Q$, and $bR_{t} \overline{Q}/Q$.

1.
$$\bar{Q}/Q: Q - \bar{Q} = S_{t+1} - S_t$$

Multiplying both sides by AC

 $AC(Q - \overline{Q}) = AC(S_{t+1} - S_t)$

Change in finished goods inventories (AI in dollars) when both inventories are measured in current prices.

Therefore, $\Delta I = \lambda C \overline{Q}(Q/Q - 1)$

On simplification this gives,

 $\frac{\overline{Q}}{\overline{Q}} = \frac{1}{(\Delta I / CGS + 1)}$

2. PKK 0/0:

Substituting for \overline{Q}/Q from the last equation,

$$P_{\theta}^{K} \overline{Q} / Q = \frac{1}{(\Delta I / CGS + 1)} P_{\theta}^{K} K$$

Amount of capital other than rav materials (i.e. Plant and Equipment at current prices at the beginning of the period) to be allocated to the finished goods actually sold.

3. $b\overline{R}_t \ \overline{Q}/Q$: now $Sr_t + \overline{R}_t = R_t + Sr_{t+1}$, and $R_t \ \overline{Q}/Q = \overline{R} = quentity of raw materials actually used$ in the goods sold. $Therefore substituting for <math>R_t$ and $R_t \ \overline{Q}/Q$ $b\overline{R} = b(Sr_t + \overline{R}_t - Sr_{t+1}) \ \overline{Q}/Q$ $= (b(Sr_t - Sr_{t+1}) + b\overline{R}_t] \ \overline{Q}/Q$ = (Change in value + Purchases of raw] Quantity sold of raw materials materials Quantity manufactured inventories valued at b = b_e Substituting for \overline{Q}/Q , P_{g}^{K} , \overline{Q}/Q , and $bR_{t}^{}$, \overline{Q}/Q . in eqn (A.11), <u>Accounting Gross Profit</u> = $(1 - FC + \frac{\alpha FC}{\epsilon_{i}}) + (\frac{1 - \alpha}{\epsilon_{i}})^{FC.Si}$ $+ \frac{\mu}{\tau t} \frac{1}{\frac{\Delta I}{CGS} + 1} \frac{1}{Sales}$ Change in the values ray

$$\begin{array}{c} \text{change in the values} & \text{fav} \\ \text{of raw material} & + \text{material} \\ + (h\mu_{\Delta} - 1) \begin{bmatrix} 1 \\ \frac{1}{\Delta I} \end{bmatrix} \begin{bmatrix} \text{inventories} & \text{purchases} \end{bmatrix} (A.12) \\ \hline \text{Sales} \\ \hline \text{CGS} + 1 \end{bmatrix}$$

IX FOOTNOTES

- In 1972 and 1973, for example, Senator Phillip Hart introduced an Industrial Reorganization Act under which one of the indicators of monopoly power was when a company's after tax return on stockholders' equity exceeded 15 percent for five consecutive years (Goldschmidt, Appendix B).
- ² Hagerman and Senbet [1976], Nair[1979].
- ³ See Weiss [1974] for a review of 46 papers and Benston [1973,1985] for a reveiw of about 20 more.
- Concentration can be defined as the percentage of total industry sales contributed by the largest firms, ranked in the order of market share.
- 5 Professor Lerner proposed an index of monopoly power as:

price-marginal cost price

here the degree of monoply power depends on the divergence between price and marginal cost

6
The FOC's give,
aT/aQ = P'(Q)Q + P(Q) - C'(Q) = 8
i.e. -P'(Q)Q/P = {P(Q) - C'(Q)}/P
or, -(dQ/Q)/(dP/P) = 1/{GP} = {P(Q) - C'(Q)}/P

 α can be thought of as a kind of elasticity of the rival's output with respect to firm's own. So $\alpha = 1$ implies that the rival firms imitate the firm's actions in the same proportion, while $\alpha > 1$ implies a greater degree of collusion (Martin[1985])

- 8 Inclusion of the last period's inventory s, at last period's price p. (gt-1,s,) will not change the model as none of the variables feature in this period's decision problem.
- 9

7

The "one period" model given by eqn.(6) gives the same results as the dynamic version. This can be

seen by setting up the dynamic version of the PDV as, x x 00

$$-P_{\theta}^{N_{1}}K_{1} = b_{\theta}^{N_{2}}K_{2} + \frac{\int (1/1+r^{*})^{t} P_{t}(S_{t+1}, \overline{Q}_{t})\overline{Q}_{t} - W_{t}L_{t}}{t=1}$$

$$- C^{F}(S_{t+1}) - C^{F}(K_{2,t+1}) - P_{t}^{K_{1}}(K_{1,t+1} - (1 - \delta)K_{1,t})$$

$$- b_{t}^{K_{2}}(K_{2,t+1} - (1 - \delta)K_{2,t})$$

here K₁ and K₂ are the two kinds of capital (i.e. long lived assets and raw materials respectively). Collecting K₁, and K₂, terms inside the integration and using the definitions of $_{\delta}$ and $_{\delta}$, we get,

$$PDV = \int (1/1+r^{*})^{t} P_{t}(S_{t+1}, \overline{g}_{t}) \overline{g}_{t} - W_{t}L_{t} - C^{F}(S_{t+1}) \\ - C^{R}(K_{2,t+1}) - \mu_{\delta,t}P_{t-1}K_{1,t} - \mu_{\delta,t}P_{t-1}K_{2,t}$$

For optimising purposes this is the same as equation (7). This result follows from Arrow's[1964] Nyopia principle which states that,

"---the most striking feature of the optimal policy is it's independence of future movements of the profit function-----. This myopic property of the optimal capital policy implies a considerable economy of information needs in a firm's decision making process." (Arrow [1964],pp.27-8.)

Also when the model is developed to incorporate imperfect markets, the test for market power essentially is the same as the test that is developed here. For a complete treatment see Martin [1985].

- 10 It is interesting to note here that if one was to assume constant returns to scale, i.e. FC = 1, the testfor market power would be of the form $a_{ij} + a_{ij}S_{ij} > 0$.
- 11 The correct definition and measurement of market share is an unresolved issue in Industrial economics literature, and is not in the scope of this study.
- 12 Errors in the context of the model in this study refer to the differences that are introduced in variables such as value of inventories, CGS, and plant and equipment when alternative accounting techniques are used.

- 13 Note that Replacement cost restatement implicitly includes the LIFO method of inventory measurement.
- 14 See Judge pp. 515.
- 15 The error in this variable is only due to techniques such as LIFO vs. FIFO, and HC vs. RC. This error is equal to 'zero' if the percentage difference introduced in the CGS is the same as the percentage difference introduced in the ending inventory.
- ¹⁶ There is support in the literature for the existence of constant returns to scale(see Scherer, chapter 4 for a good review).
- 17 The quarterly data filed with the SEC is peferable to the quarterly compustat data as the latter can introduce biases when trying to estimate the ages of the long term assets (see Thies and Revsine[1977].
- ¹⁸ Source-Weiss [1974] (pp. 187).
- 19
 Adjusted R² = (1-(1-R²))(N-1)/(N-K)
 where R² = Unadjusted coefficient of the goodness of
 fit of a model
 N = Number of observations
 K = Number of parameters estimated
- 28 This reversal of the null from the hypotheses formulated earlier results in a loss of power of the test statistic. However there is precedence for such a procedure in the finance literature where the only way to test for market efficiency is to have a null of market efficiency (no information effect) and then accept it if there is no information effect.

FOOTNOTES APPENDIX

A.1 In the model proposed by Sawyer[1983], profits of a firm are,

T = $P(\overline{A}, Q) - C(Q) - P^{\overline{A}} A$, maximizing the F.O.C. is, aT/aQ = MR = MC which is equivalent to, $P(1 - 1/\epsilon)$ = MC. Assuming constant returns to scale and substituting AC for MC in the above equation gives, $(P - AC)/PQ = 1/\epsilon$ which is eugivalent to $(PQ - C(Q))/PQ = 1/\epsilon$. This is the Price cost margin or the Lener index. This measure of market power does not incorporate the advertising cost.

- A.2 <u>KNP_K + LMP_L + RMP_R = <u>Q</u> Total number of goods produced assuming that each unit produced was the marginal unit. Multiplying both sides of the above equation with 1 = MC/MC gives, Q.MC/Total cost = Q.MC/Q.AC =1/FC.</u>
- A.3 For the LIFO case exactly the same results follow except that a new implicit cost of rental for capital consisting of raw materials is defined as,

and now $\hat{B}_1 \approx b$ is the cost of materials used in production.

A.4 Same transformation as from eqn. (4) to eqn. (5).

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