

THE EFFECTS OF INCREASES IN THE  
FEDERAL MINIMUM WAGE ON SELECTED  
INDUSTRY LABOR MARKETS FOR THE 1960 - 63 PERIOD

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JAMES FRANCIS MCCARLEY

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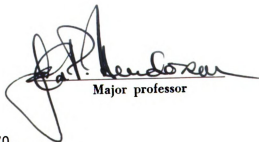


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## ABSTRACT

### THE EFFECTS OF INCREASES IN THE FEDERAL MINIMUM WAGE ON SELECTED INDUSTRY LABOR MARKETS FOR THE 1960-63 PERIOD

By

James Francis McCarley

Since the enactment of the Fair Labor Standards Act of 1938, economic literature has been filled with studies attempting to determine its impact on employment. These studies have been undertaken by both the Department of Labor and independent investigators. In general, they have used a static-tabular approach in their analysis.

The Department of Labor studies have been interpreted as indicating that increases in the statutory minimum have had negligible, and in some cases, even positive effects on employment. These results have been rationalized by some economists on the grounds that wage costs are not an important consideration to firms or that existing unskilled labor suddenly became more efficient.

Critics of the Department of Labor studies maintain that the unorthodox conclusions result from a failure to consider changes in exogenous variables and to employ adequate statistical procedures. The independent investigators,

however, have often been guilty of the same criticisms raised against the Department of Labor.

To overcome these shortcomings we have attempted in this study to specify a model which was subject to inferential tests. To accomplish this task, a model was developed to explain desired and equilibrium levels of industry employment, and from it a reduced form expression was obtained. Because incomplete adjustment to equilibrium levels of industry employment may occur, a stock adjustment model was introduced. A reduced form expression for the observed level of industry employment as a function of the minimum wage, the average wage rate in the industry, a trend variable, and the lagged observed level of industry employment was then obtained.

The final form of the model developed was used to estimate the impact from increasing the federal minimum wage on selected high and low wage SIC three digit manufacturing industries for the 1960-63 period. Both ordinary least squares and ZEF estimation procedures were used, with and without log conversions. In each case the results were theoretically consistent.

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## CHAPTER I

### INTRODUCTION

In 1938 the Congress passed the Fair Labor Standards act to eliminate poverty resulting from substandard wages, to maintain purchasing power at sufficiently high levels to insure full employment, to prevent downward wage spirals during periods of falling aggregate demand, and to foster fair methods of competition.<sup>1</sup> To accomplish these goals, the 1938 Act established a minimum hourly rate for all industries engaged in interstate commerce of 25 cents between October 24, 1938 and October 23, 1939, and provided for an additional increase to 30 cents an hour beginning October 24, 1938.<sup>2</sup> It also provided that by October 24, 1945 the minimum rate for all industries covered by the Act would be increased to 40 cents an hour.<sup>3</sup> The minimum rate, however, could be raised to 40 cents an hour for some industries before the 1945 deadline, after appropriate testimony

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<sup>1</sup>Donald E. Cullen, "Minimum Wage Laws," New York State School of Industrial and Labor Relations, Bulletin 43, (Ithaca, N. Y.: Cornell University, February, 1961), p. 3.

<sup>2</sup>U. S., Statutes at Large, Vol. LII (1938), "Fair Labor Standards Act of 1938," June 25, 1938, ch. 676, sec. 6, p. 1062.

<sup>3</sup>Ibid.

before industry committees, by the Administrator of the Wage and Hour Division of the Department of Labor, providing such increases did not substantially reduce employment or give a competitive advantage to a particular group.<sup>4</sup>

Subsequent amendments to the Fair Labor Standards Act of 1938 increased the minimum hourly rate to 75 cents on January 25, 1950<sup>5</sup> and to \$1.00 an hour on March 1, 1956.<sup>6</sup> In addition, the minimum rate was raised to \$1.15 an hour on September 3, 1961 on those occupations previously covered and to \$1.00 on those employments then included under the Act.<sup>7</sup> In September of 1963 the minimum rate was again increased to \$1.25 an hour on previously covered jobs and to \$1.15 an hour on those occupations initially covered in 1956.<sup>8</sup> The most recent extension of the minimum wage raised the minimum hourly rate to \$1.40 on February 1, 1967 and provided for an additional increase to \$1.60 an hour on February 1, 1968.<sup>9</sup> In addition the 1966 amendments

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<sup>4</sup>Ibid., sec. 8, p. 1064.

<sup>5</sup>U.S., Statutes at Large, Vol. LXIII, pt. 1 (1949), "Fair Labor Standards Amendments of 1949," Oct. 26, 1949, ch. 736, sec. 6, p. 912 and Ibid., sec. 16, p. 920.

<sup>6</sup>U.S., Statutes at Large, Vol. LXIX (1955), "Fair Labor Standards Amendments of 1955," Aug. 12, 1955, ch. 867, sec. 3, p. 711.

<sup>7</sup>U.S., Statutes at Large, Vol. LXXV (1961), "Fair Labor Standards Amendments of 1961," May 5, 1961, Pub. L. 87-30, sec. 5, p. 67.

<sup>8</sup>Ibid.

<sup>9</sup>U.S., Statutes at Large, Vol. LXXX, pt. 1, (1966), "Fair Labor Standards Amendments of 1966," Sept. 23, 1966, Pub. L. 89-601, sec. 301, p. 8381.

provided for a minimum hourly rate of \$1.00 for those industries that were not previously covered by the Act.<sup>10</sup>

### I. Review of Impact Studies

Following the enactment of the Fair Labor Standards Act and its subsequent extensions, a great deal of controversy arose regarding the effect of establishing wage rates at levels different than those determined by the market. As a result of this controversy various industry studies were undertaken to determine the impact of such action after the passage of the original Act and each of its amended forms.

#### Studies Following the FLSA of 1938

H. M. Douty analyzed the impact of establishing a minimum wage in the seamless hosiery industry from September, 1938 to September, 1940.<sup>11</sup> The study was composed of two samples: the first related the effect of the wage change on all the firms in the industry in 1938 and 1940, and the second compared 91 plants which were included in the Bureau of Labor Statistics surveys of 1938 and 1940, and still in business in 1940.<sup>12</sup> The establishment of these minima had a decided upward push on wages in the

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<sup>10</sup>Ibid., sec. 302-303, p. 838-839.

<sup>11</sup>H. M. Douty, "Minimum Wage Regulation in the Seamless Hosiery Industry," Southern Economic Journal, Vol. VIII, (October, 1941), p. 176.

<sup>12</sup>Ibid., p. 179.

industry increasing average hourly earnings from 35.1 cents to 41.3 cents between 1938 and 1940.<sup>13</sup> The effect of establishing a minimum wage in this industry varied greatly among plants with widely separated wage levels.<sup>14</sup> Of the 91 plants included in both the 1938 and 1940 surveys there was a 3.9 per cent reduction in employment.<sup>15</sup> When all plants in the survey were considered the average decline in employment was greatest for those plants which were paying average wages below the 32.5 cent average for the industry in 1938.<sup>16</sup> This largely explains the narrowing of average wage rates between the North and the South<sup>17</sup> and the larger decreases in employment in the South as opposed to those in the North (employment decreased by 5.5 per cent in the South and increased by 5.0 per cent in northern plants).<sup>18</sup> In those southern plants where average hourly rates were 32.5 cents, employment also increased.<sup>19</sup> This study did not indicate that minimum wages resulted in the dislocation of employment because of plant failures, although some

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<sup>13</sup>Ibid.      <sup>14</sup>Ibid., p. 182.

<sup>15</sup>Ibid., p. 183.      <sup>16</sup>Ibid.

<sup>17</sup>Ibid., p. 179.

<sup>18</sup>Ibid., p. 184. Conclusions similar to these were also reached by: J. F. Maloney, "Some Effects of the Federal Fair Labor Standards Act Upon Southern Industry, Southern Economic Journal, Vol. 9, (July, 1942), pp. 15-23.

<sup>19</sup>Douty, loc. cit.



transfer of workers from low to high wage firms did occur.<sup>20</sup> While total employment for the industry did not appear to be significantly affected, Douty does suggest that there may have been some improvements in managerial performance and a substitution of new machinery for old.<sup>21</sup>

A second study utilizing the 1938 and 1940 surveys of 97 plants in the seamless hosiery industry attempts to determine the effect of the minimum wage on employment in the low wage mills between October, 1938 and September, 1939.<sup>22</sup> This investigation divides the sample into three groups: the first paying average hourly earnings of less than 25 cents before October 24, 1938, the second paying average hourly earnings between 25 cents and 27.5 cents and the third paying average hourly earnings between 29 cents and 35.5 cents.<sup>23</sup> After the imposition of the 25 cent minimum the average hourly earnings for all three groups increased with the largest increase occurring in those plants included in Group I.<sup>24</sup> The Group I plants, however, experienced significant decreases in both the number employed and man-hours worked while both of these factors either remained constant or increased for the

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<sup>20</sup>Ibid., p. 186.

<sup>21</sup>Ibid.

<sup>22</sup>A. F. Hinrichs, "Effects of the 25-cent Minimum Wage on Employment in the Seamless Hosiery Industry," Journal of the American Statistical Association, Vol. XXXV (March, 1940), p. 13.

<sup>23</sup>Ibid., p. 15.

<sup>24</sup>Ibid., p. 17.

Group II and Group III plants.<sup>25</sup> In addition to the wage and employment adjustments which occurred in the low wage mills there were changes in production methods from those that were labor intensive to those that were capital intensive.<sup>26</sup>

Studies Following the 1949  
Amendment to the FLSA

The effect of the 1949 amendment to the Fair Labor Standards Act on southern sawmills, fertilizer, men's dress shirts and nightwear, men's seamless hosiery, and wood furniture industries was analyzed by the Bureau of Labor Statistics for the payroll periods just before and after the Act became effective.<sup>27</sup> In each of these five industries a substantial number of employees were earning less than 75 cents an hour in 1949.<sup>28</sup> Examination of payroll records after the new Act became effective indicated adjustments in average earnings higher than necessary to comply with the new minimum, a narrowing of wage rates paid by firms in these industries, and a narrowing of occupational wage differentials in those industries paying their employees by piece rate.<sup>29</sup> The investigation did not, however, indicate any significant nonwage adjustments in the five industries surveyed.<sup>30</sup> The change in capital expenditures

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<sup>25</sup> Ibid., p. 19.

<sup>26</sup> Ibid., p. 22-23.

<sup>27</sup> H. S. Kantor, "Economic Effects of the Minimum Wage," Monthly Labor Review, Vol. LXXVIII (March, 1955), p. 308.

<sup>28</sup> Ibid.

<sup>29</sup> Ibid., p. 309.

<sup>30</sup> Ibid.

and hours worked did not significantly change for the short period considered except in the fertilizer and saw-milling industries where the trend toward mechanization was continued.<sup>31</sup> Additional surveys made for high wage and low wage covered and noncovered industries for the period 1938 to 1951 did, however, indicate increases in average hourly earnings of 121 per cent for the first group, 171 per cent for the second group and 125 per cent for the third group.<sup>32</sup> It was also concluded from these later studies that the effect from establishing wage minima was not only to increase but to substantially maintain the relative improvements in earnings for those workers affected by the minimum wage increase.<sup>33</sup>

#### Studies Following the 1955 Amendment of the FLSA

The effect of raising the minimum wage from 75 cents to \$1.00 an hour was investigated by the Department of Labor for selected low-wage industries. These studies attempted to determine the impact from increasing the minimum wage on nonseasonal and seasonal industries as well as any diffusive effects it might have. They all utilized payroll surveys taken by the Bureau of Labor Statistics prior to and following the effective date of the 1955 amendment.

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<sup>31</sup>Ibid.

<sup>32</sup>Ibid., p. 311.

<sup>33</sup>Ibid.

Nonseasonal industries.---The effect of the 1955 amendment on the cigars, fertilizer, seamless hosiery, and wooden containers industries was examined by the Bureau of Labor Statistics for the August, 1955, February and April 1956, and April, 1957 payroll periods.<sup>34</sup> The initial impact on these industries followed the same pattern as in previous studies,<sup>35</sup> i.e., average hourly earnings in each industry increased, a larger proportion of workers just earned the new minimum compared to those that had previously earned the 75 cent minimum, and a narrowing of the dispersion of average hourly earnings.<sup>36</sup> Most of these changes occurred on the date the new amendment became effective.<sup>37</sup> Analysis of subsequent payroll periods, however, indicated that the wage structure in these industries restored itself with the degree of restoration varying widely in each industry.<sup>38</sup>

Employment for the five industries decreased by slightly over 4 per cent between August 1955 and April 1956, and by more than 8 per cent between April 1956 and April 1957.<sup>39</sup> Total employment between the 1955 and 1957

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<sup>34</sup>N. J. Samuels, "Effects of the \$1 Minimum Wage in Five Industries," Monthly Labor Review, Vol. LXXXI (May, 1958), p. 492.

<sup>35</sup>N. J. Samuels, "Effects of the \$1 Minimum Wage in Seven Industries," Part I, Monthly Labor Review, Vol. LXXX (March, 1957), pp. 323-28 and Part II, Monthly Labor Review, Vol. LXXX (April, 1957), pp. 441-46.

<sup>36</sup>Samuels, loc. cit.

<sup>37</sup>Ibid., p. 493.

<sup>38</sup>Ibid., p. 492.

<sup>39</sup>Ibid., p. 493.

payroll periods declined between 6 per cent and 15 per cent for those industries studied.<sup>40</sup> Some of this decline was attributed to changes in factors such as product demand, rather than from cost increases resulting from the higher minimum.<sup>41</sup> Other studies conducted by the Bureau of Labor Statistics found that only about 15 per cent of the reporting firms attributed the discharge of employees to the \$1.00 an hour minimum, with the most widely used adjustment being increased expenditures for machinery and equipment.<sup>42</sup>

Seasonal industries.--The payroll surveys in the fruit and vegetable canning and freezing, raw cane sugar manufacturing, and tobacco stemming and redrying industries, which are composed of a generally transitory unskilled work force, were taken at peak employment periods in 1955 and 1956.<sup>43</sup> The effect on earnings and employment in these industries again followed patterns similar to those in other low wage industries. There was an immediate increase in the wage level, a greater concentration of workers at the new minimum and a substantial reduction in

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<sup>40</sup>Ibid.      <sup>41</sup>Ibid., p. 494.

<sup>42</sup>N. J. Samuels, "Plant Adjustments to the \$1 Minimum Wage," Monthly Labor Review, Vol. LXXXI (October, 1958), p. 1138.

<sup>43</sup>N. J. Samuels, "Effects of the \$1 Minimum Wage in Three Seasonal Industries," Monthly Labor Review, Vol. LXXX (September, 1957), p. 1087.

occupational and geographic wage differentials.<sup>44</sup> Employment decreased by 12.5 per cent in raw cane sugar manufacturing and by 11.3 per cent in fruit and vegetable canning and freezing, but increased by 1.6 per cent in tobacco stemming and redrying.<sup>45</sup> The total employment for all three industries fell by 3.5 per cent between the 1955 and 1956 period.<sup>46</sup> The variations in employment in the raw cane sugar manufacturing was attributed to obsolescence of existing plants which made them unprofitable and necessitated some of them closing.<sup>47</sup> It was believed, however, that in tobacco stemming and redrying normal fluctuations in employment could have accounted for the 1.6 per cent increase in employment recorded.<sup>48</sup>

. Diffusive effects.--Area studies were conducted using payroll data which divided industries into groups subject to the \$1.00 an hour minimum and those not covered by the 1955 amendment. In these studies the average pay levels between the subject and nonsubject industries widened when the new minimum became effective and narrowed in the following year, but remained wider in April 1957 than in February, 1956.<sup>49</sup> A comparison of the April 1957 payroll period with the June 1959 payroll period indicated no clear

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<sup>44</sup>Ibid.      <sup>45</sup>Ibid., Table I, p. 1088.

<sup>46</sup>Ibid.      <sup>47</sup>Ibid., p. 1090.      <sup>48</sup>Ibid., p. 1091.

<sup>49</sup>L. Badenhop, "Effects of the \$1 Minimum Wage in Seven Areas," Monthly Labor Review, Vol. LXXXI (July, 1958), p. 737.

pattern of change in relative wage differentials between covered and noncovered industries.<sup>50</sup> Comparing the firms which were required to adjust with those that were not it was found that the adjusting firms increased production more, introduced more machinery, increased efficiency at a faster rate, and utilized opportunities to reduce overtime more than did firms not facing an adjustment mandate.<sup>51</sup> In most cases it was believed that wage and nonwage adjustments were made with little or no general effect on employment.<sup>52</sup>

M. R. Colberg, utilizing data for 517 manufacturing plants, compared the impact of the \$1.00 an hour minimum on high and low wage Florida counties.<sup>53</sup> During the four month period between January and April 1956 the new legal minimum had only a slight effect on man-hours worked in high wage counties but a very substantial diminutive effect in the low wage group.<sup>54</sup> From January 1956 to January 1957, however, man-hours worked increased 5.1 per cent for the

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<sup>50</sup>H. Schaffer, "Effects of the \$1 Minimum Wage in Six Areas," Monthly Labor Review, Vol. LXXXIII (May, 1960), p. 474.

<sup>51</sup>P. A. Brinker, "The \$1 Minimum Wage Impact of Fifteen Oklahoma Industries," Monthly Labor Review, Vol. LXXX, (September, 1957), p. 1092.

<sup>52</sup>Ibid., p. 1092 and Badenhop, loc. cit.

<sup>53</sup>M. R. Colberg, "Minimum Wage Effects on Florida's Economic Development," Journal of Law and Economics, Vol. III, (October, 1960), pp. 106-117.

<sup>54</sup>Ibid., p. 113.

high wage counties and decreased by less than 1 per cent in the low wage group.<sup>55</sup> The increase in employment in the high wage counties was largely attributed to increases in product demand during this period.<sup>56</sup> In the low wage group it was felt that the small decline in employment was due to a movement toward a better quality of worker and an attempt to utilize these workers more effectively during seasonal peaks.<sup>57</sup> The decreases in employment were found to be greater the larger the required increase in the wage rates.<sup>58</sup> Colberg attributes part of the decline in population in the low-wage counties to a lack of employment opportunities resulting from the \$1.00 an hour minimum wage.<sup>59</sup>

Studies Following the 1961  
Amendment of the FLSA

The 1961 amendment to the Fair Labor Standards Act initially increased the minimum rate from \$1.00 to \$1.15 an hour and extended its coverage mainly to workers in the retail and construction industries at \$1.00 an hour. At a later date it raised the \$1.15 an hour rate to \$1.25 and the \$1.00 an hour rate to \$1.15. To determine the impact of these changes the Department of Labor conducted a nationwide survey of the retail trade and selected surveys for non-metropolitan areas in the North Central and southern regions and small metropolitan areas of the South.

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<sup>55</sup>Ibid., p. 113.

<sup>56</sup>Ibid., Table 3, p. 112.

<sup>57</sup>Ibid.      <sup>58</sup>Ibid.

<sup>59</sup>Ibid., p. 116.



The surveys conducted between 1960 and 1962 indicated no general upward pressure on the wage structure resulting from increases in the statutory minimum.<sup>60</sup> The wage effects that did occur were generally confined to those increases necessary to bring the wages of covered workers to the new minimum level.<sup>61</sup> The greatest impact occurred in the non-metropolitan areas of the South where 44 per cent of the workers in the wholesale trade and 27 per cent of those in manufacturing received the wage increases necessary to bring their hourly earnings to \$1.15.<sup>62</sup> In some low wage areas employees earning \$1.15 an hour or slightly higher did receive wage increases but there was no consistent pattern to these increases.<sup>63</sup> In addition, workers in noncovered industries did not benefit from the changes that occurred in subject industries, except for food service workers in newly covered retail stores.<sup>64</sup>

Further studies that were conducted indicated that not only did unemployment decrease between 1961 and 1965 but that employment increased at a faster rate and to a greater extent among the less skilled groups in the society

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<sup>60</sup> U.S. Department of Labor, Wage and Hour and Public Contracts Division, Report Submitted to the Congress in Accordance with the Requirements of Section 4 (d) of the Fair Labor Standards Act, (Washington: Government Printing Office, January 1963), p. iii.

<sup>61</sup> Ibid.

<sup>62</sup> Ibid., p. v.

<sup>63</sup> Ibid., p. iii.

<sup>64</sup> Ibid.

during this period.<sup>65</sup> Other studies that included 34 areas in the North Central and southern nonmetropolitan areas, as well as metropolitan localities in the South, indicated that in 25 areas employment increased, in eight areas employment decreased and in one area there was no change.<sup>66</sup> In the 25 areas showing employment advances 15 had increases of 15 per cent or more, and in the eight areas where employment declined the percentage of decline was six per cent or less.<sup>67</sup> These studies maintained that there was no reason to believe employment increases would have been any larger if the minimum wage had not been raised or extended.<sup>68</sup> In general, the conclusions reached were that the 1961 amendment did not result in employment decreases, inflation, a retardation of economic development or smaller profits.<sup>69</sup>

The Department of Labor also conducted a nationwide survey of the impact of the new wage minima on the retail trade for the June 1965 payroll period, which was nine months after the \$1.15 an hour edict became effective for this sector. This study followed previous surveys which

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<sup>65</sup>U.S. Department of Labor, Wage and Hour and Public Contracts Division, Minimum Wage and Maximum Hour Standards Under the Fair Labor Standards Act - An Evaluation and Appraisal, (Washington: Government Printing Office, January 1966), p. 7.

<sup>66</sup>Ibid.

<sup>67</sup>Ibid., pp. 7-8.

<sup>68</sup>Ibid., p. 7.

<sup>69</sup>Ibid., p. 10.

were made in June 1961, which was one month after the \$1.00 an hour minimum became effective, and in June 1962 which was nine months after its effective date.<sup>70</sup> In this case about 70 per cent of the employees were at the new minimum.<sup>71</sup> Employers, however, seemed to anticipate the \$1.25 an hour minimum which was to become effective in three months since more workers were concentrated at this rate than at the \$1.15 an hour level.<sup>72</sup> The greatest proportional effects on employee earnings occurred in the nonmetropolitan areas and in the southern regions.<sup>73</sup> In the noncovered portions of the retail trade the proportion of workers earning less than \$1.25 an hour was reduced by 33 1/3 per cent since the June 1962 study.<sup>74</sup> Employment levels were higher for all regions surveyed after the introduction of each of the minima.<sup>75</sup> Much of the employment decline which occurred in nonmetropolitan areas was thought to be due to a statistical reclassification of these areas resulting from population growth.<sup>76</sup> In the various retail lines employment decreases were greatest in the limited price variety stores where the decline in employment levels more than offset the increases that occurred after the \$1.00 an hour minimum

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<sup>70</sup>U. S. Department of Labor, Wage and Hour and Public Contracts Division, Studies of the Effects of the Minimum Wage and Maximum Hour Standards of the Fair Labor Standards Act in the Retail Trade, (Washington: Government Printing Office, January, 1966), p. 1.

<sup>71</sup>Ibid., p. 3.      <sup>72</sup>Ibid.      <sup>73</sup>Ibid.

<sup>74</sup>Ibid.      <sup>75</sup>Ibid., p. 4.      <sup>76</sup>Ibid.

became effective.<sup>77</sup> In the noncovered segments of the retail trade sector, employment levels increased during each of the periods surveyed.<sup>78</sup>

### Summary and Evaluation of Studies

The investigations of the Department of Labor on the impact of various minimum wages found (1) an increase in average hourly earnings, (2) a narrowing of occupational and geographic wage differentials, (3) some diffusive effects but no consistent pattern to their change, (4) some concentration of workers' average hourly earnings at the new minimum, or its anticipated level, (5) some, but not complete, restoration of prior wage differentials over longer periods and (6) only slight increases or decreases in employment.

The conclusions of these studies regarding the effects on employment from increasing wage rates by imposing minimum wages have been challenged by several authors.<sup>79</sup> In general, it is the contention of this group that labor

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<sup>77</sup>Ibid.      <sup>78</sup>Ibid.

<sup>79</sup>Note for example the following studies: Yale Brozen, "Minimum Wage Rates and Household Workers," The Journal of Law and Economics, Vol. 5, (October, 1962), pp. 103-109; M. R. Colberg, op. cit., pp. 106-116; David E. Kaun, "Minimum Wages, Factor Substitution and the Marginal Producer," Quarterly Journal of Economics, Vol. 79, (August, 1965), pp. 478-486; G. Macesich and C. T. Stewart, "Recent Department of Labor Studies of Minimum Wage Effects," Southern Economic Journal, Vol. 26, (April, 1960), pp. 281-290; J. M. Peterson, "Employment Effects of Minimum Wages, 1938-50," Journal of Political Economy, Vol. 65, (October, 1957), pp. 412-30 and C. P. Sawaya, "The Employment Effects of Minimum Wage Regulation in the Southern Pine Lumber Industry," (unpublished Doctoral dissertation, University of Indiana, June 1958).

markets are substantially competitive in nature and that wage increases through statutory means must therefore result in unemployment in the low wage groups. If unemployment is not apparent they feel that it must either have been offset by changes in exogenous variable or hidden by considering only gross changes in employment. A further contention seems to be that the statistical analysis utilized is sufficiently unsophisticated as to neither preclude nor conclude the results derived from the above studies.

Other authors, however, believe that the results obtained from the Department of Labor studies are substantially valid. They contend that statutory increases do not cause unemployment either because (1) wage costs are not an important consideration to non-profit maximizing firms<sup>80</sup> or (2) that existing factors suddenly become more efficient and therefore do not necessitate changes in the firm's factor mix.<sup>81</sup> It is the belief of these authors

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<sup>80</sup>Richard A. Lester, "Shortcomings of Marginal Analysis for Wage-Employment Problems," American Economic Review, (March, 1946), p. 76. Professor Machlup in an article rebutting Lester and other anti-marginalists contended their failure to see the applicability of marginal analysis was often due to a lack of full insight into the theory itself, to the utilization of inappropriate research techniques, or to an inability to interpret correctly the results of their investigations. He concluded that Lester's study has neither discredited or disproved orthodox theory of the firm. Fritz Machlup, "Marginal Analysis and Empirical Research," American Economic Review, Vol. XXXVI, No. 4, Pt. 1, (September, 1946), pp. 519-54.

<sup>81</sup>Lloyd G. Reynolds, Labor Economics and Labor Relations, (New York: Prentice-Hall, Inc., 1956), p. 661.

that the assumption of a competitive model is invalid and that orthodox theory is generally inept in this instance.

## II. The Present Study

The studies previously discussed in this chapter have typically assumed that industry labor markets were in equilibrium prior to an increase in the minimum wage. Since the demand for labor is a derived demand, this assumption implies that industry product markets were also in equilibrium during these periods. An indication that this contention may not be valid can be seen in Table 1.1 which shows changes in the Federal Reserve Board Index of production activity for selected manufacturers. During the periods of the 1950, 1956 and the 1963 increases of the federal minimum wage, aggregate output was in general expanding for each of these industry classifications. For the 1961 change in the minimum wage, however, production activity was declining for most industry groups.

Under the above conditions the impact of the minimum wage need not lead to a decrease in equilibrium levels of employment. Since manufacturing industries were already expanding employment in an attempt to increase output, it is possible that the equilibrium level of industry employment could increase, decrease, or remain constant after the increase in the federal minimum.

The disemploying effects of an increase in the minimum wage may, however, still be present. They result

TABLE 1.1.--Changes in Industrial Production Indexes, Selected Manufacturers, 1949-64 [1957-59 = 100].

Year	Durable Manufactures					Nondurable Manufactures					
	Primary Metals	Fabricated Metal Products	Machinery	Transportation Equipment	Instruments and Related Products	Clay, Glass and Lumber	Furniture and Misc.	Textile Apparel & Leather Products	Paper & Printing	Chemicals and Petroleum Products	Food Beverages & Tobacco
1949	-14.9	- 7.4	- 7.5	+ .2	- 6.0	- 7.4	- 5.8	- 3.9	- .1	- 1.4	+ .8
1950a	+20.9	+15.6	+13.7	+ 9.3	+ 8.1	+15.4	+12.1	+ 8.5	+ 7.4	+11.3	+ 2.8
1951	+ 8.8	+ 5.8	+10.3	+ 6.5	+ 8.4	+ 4.3	- 3.5	- 1.7	+ 2.7	+ 6.7	+ 1.8
1952	- 9.4	- 2.2	+ 9.1	+10.2	+12.4	- 2.7	+ 2.2	+ 2.1	- 1.7	+ 2.5	+ 1.9
1953	+13.2	+11.3	+ 8.4	+18.6	+ 7.2	- 3.4	+ 7.3	- 1.2	+ 4.9	+ 5.3	+ .9
1954	-21.2	-10.1	-12.8	- 7.9	- 2.4	- 3.1	- 2.9	- 3.3	+ 2.4	- .5	+ 1.6
1955	+27.1	+ 8.1	+ 8.8	+13.2	+ 5.8	+11.1	+11.1	+ 8.6	+ 7.5	+12.1	+ 3.3
1956a	- 2.0	+ .5	+10.6	- 4.6	+ 6.7	+ 1.3	+ 3.1	+ 2.5	- 4.6	+ 4.6	+ 3.5
1957	- 4.2	+ 2.7	- 2.9	+ 9.0	+ 2.6	- 4.5	- 3.4	- 1.1	+ .7	+ 4.2	+ .1
1958	-24.7	- 8.6	-15.4	-16.9	- 5.9	- 3.4	- 4.3	- 1.9	- .8	+ .1	+ 2.7
1959	+12.9	+12.6	+18.3	+14.5	+17.8	+14.4	+15.7	+13.1	+ 8.2	+13.4	+ 4.5
1960	+ .9	+ 2.1	+ 3.7	+ 4.2	+ 6.6	- 2.8	+ 4.3	- .6	+ 3.8	+ 5.0	+ 2.7
1961a	- 2.4	- 1.1	- .4	- 4.6	- .7	- 1.2	+ .8	+ .9	+ 3.4	+ 5.0	+ 3.6
1962	+ 5.7	-10.6	+13.1	+14.7	+ 7.2	+ 4.8	+10.4	+ 6.7	+ 4.3	+12.3	+ 3.1
1963a	+ 8.7	+ 6.3	+ 5.7	+ 8.7	+ 7.2	+ 5.1	+ 4.6	+ 3.4	+ 3.4	+10.6	+ 3.5
1964	+15.8	+ 9.3	+12.2	+ 3.7	+ 6.2	+ 6.7	+ 9.3	+ 6.7	+ 7.4	+10.7	+ 4.0

a Year of increase in the federal minimum wage.

Source: Economic Report of the President, U. S. Government Printing Office, Washington: 1969, p. 269.

because the equilibrium level of output that would have maximized profits of the firms in the industry before the minimum was raised may be greater than the level of employment after the increase in the federal minimum.

These disemploying effects of an increase in the minimum wage on unskilled labor can be illustrated in Figure 1.1 where  $N_1$  is the equilibrium level of employment at the present level of the minimum wage,  $W_m$ , in the industry labor market. The period of expansion in the industry can be represented in this graph by a shift in the demand curve for unskilled labor from  $D_j$  to  $D'_j$ . With this increase in the demand for unskilled labor the new profit maximizing level of employment in the industry would be  $N_2^*$  at the wage rate  $W_m$ . An increase in the minimum wage to  $W'_m$  at this point would cause this industry to employ  $N_2$  units of labor instead of the level  $N_2^*$ . The effect of an increase in the federal minimum would, therefore, cause a decrease in the potential level of employment by the quantity  $N_2^* - N_2$ . The magnitude of this decrease will depend upon the elasticity of the demand function for unskilled labor in the industry.

The above analysis is used in the present study, but it is modified because data is only available for observed levels of total industry employment. The total employment, however, may be interpreted as the sum of the unskilled and skilled levels of industry employment. A model to



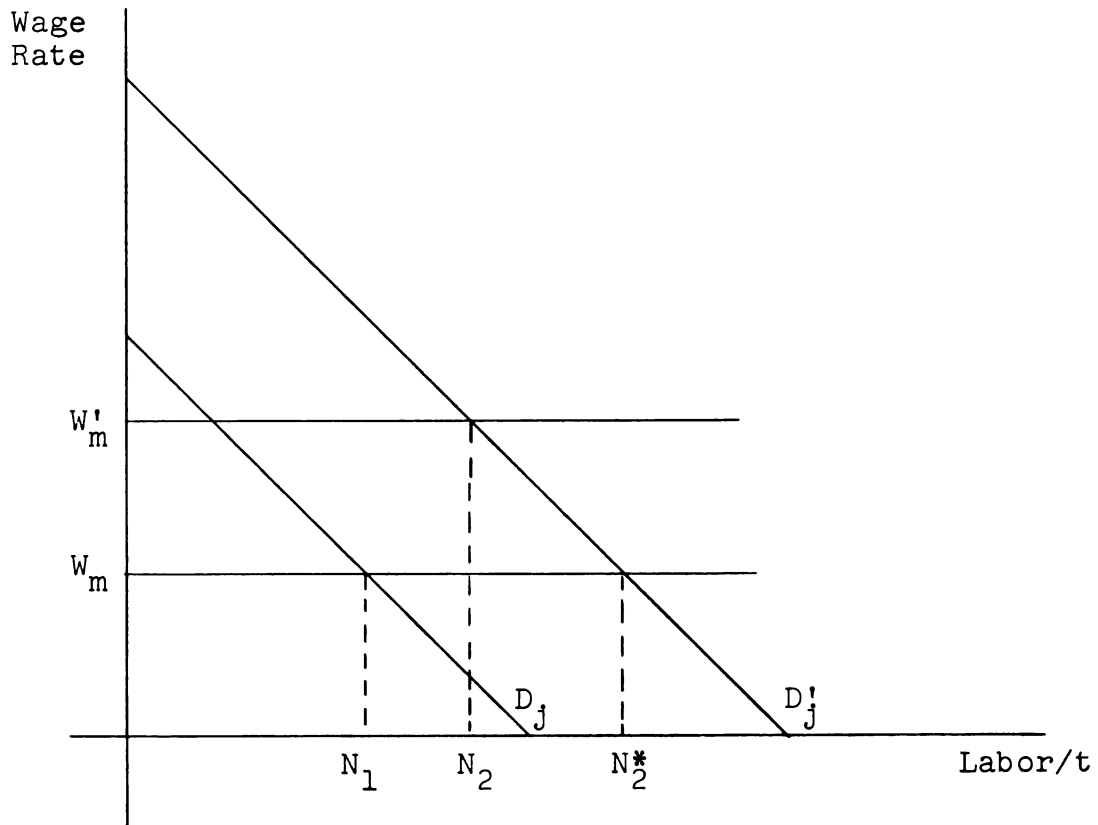


Figure 1.1.--Industry Labor Markets (Unskilled Labor).

establish profit maximizing or equilibrium levels of each type of labor can be developed and a reduced form expression for the total equilibrium level of industry employment obtained. Because of incomplete adjustment, the equilibrium level of industry employment are in general unobservable. An expression for the observed level of industry employment is, therefore, derived.

The final form of the model developed is applied to selected high and low wage SIC three digit manufacturing industries for the 1960-63 period. Both ordinary least squares and Zellner's two stage Aitken's estimation procedures are used. The hypothesis to be tested is that the minimum wage increases may have reduced observed levels of industry employment.

## CHAPTER II

### THE MODEL

#### Introduction

In this chapter we examine the effect of minimum wage increases on an industry's equilibrium level of employment. First, a model is developed to explain the determination of profit maximizing levels of employment for an individual firm. The aggregate equilibrium level of employment is then obtained by summing these quantities over the individual firms. A stock adjustment model is then introduced to allow for the possibility that industry labor markets do not instantaneously adjust to their desired levels of employment. The final form of this model is used in Chapter III to test the central hypothesis of this study.

#### Equilibrium Levels of Employment

For the period of this study the following assumptions will be made: 1) the firms in each industry are rational economic units that have sufficient knowledge of their cost and revenue data to maximize profits; 2) the firms' production functions are multivariable with first derivatives that are positive and monotonically decreasing over the relevant range of input quantities; 3) technology is constant; 4) product demand is a function of and

varies inversely with product prices; 5) the industry supply functions of labor are perfectly elastic; and 6) unskilled labor operates at the minimum wage rate and skilled labor at the skilled rate.

The equilibrium level of employment for the  $i^{\text{th}}$  firm in a typical industry can be determined, given the above assumptions, once the product demand, production and profit functions for the firms have been specified. Assume that these functions for the  $i^{\text{th}}$  firm have the following form:

$$x_{i,t} = X(n_{1,i,t}, n_{2,i,t}) \quad (\text{production function}) \quad (2.1)$$

$$\frac{\partial X}{\partial n_{1,i,t}} > 0, \quad \frac{\partial X}{\partial n_{2,i,t}} > 0, \quad \frac{\partial^2 X}{\partial n_{1,i,t} \partial n_{2,i,t}} > 0$$

$$\text{and} \quad \frac{\partial^2 X}{\partial n_{1,i,t}^2} < 0, \quad \frac{\partial^2 X}{\partial n_{2,i,t}^2} < 0$$

$$(i = 1, 2, 3, \dots, n)$$

$$(t = 1, 2, 3, \dots, m)$$

$$p_{i,t} = P(x_{i,t}, t) \quad (\text{product demand function}) \quad (2.2)$$

$$\frac{\partial P}{\partial x_{i,t}} < 0$$

$$\pi_{i,t} = x_{i,t} p_{i,t} - w_{1,t} n_{1,i,t} - w_{2,t} n_{2,i,t} - f_{i,t} \quad (2.3)$$



where the arguments are defined in the following way:

$x_{1,t}$  = the output of the  $i^{\text{th}}$  firm in a typical industry during period  $t$ .

$p_{1,t}$  = the output price of the  $i^{\text{th}}$  firm in a typical industry during period  $t$ .

$\Pi_{1,t}$  = the profits of the  $i^{\text{th}}$  firm in a typical industry during period  $t$ .

$n_{1,i,t}$  = the input of unskilled labor by the  $i^{\text{th}}$  firm in a typical industry during period  $t$ .

$n_{2,i,t}$  = the input of skilled labor by the  $i^{\text{th}}$  firm in a typical industry during period  $t$ .

$t$  = a trend variable which represents autonomously induced shifts in the firm's product demand functions.

$f_{1,t}$  = the fixed costs of the  $i^{\text{th}}$  firm in a typical industry during period  $t$ .

$w_{1,t}$  = the level of the federal minimum wage during period  $t$ .

$w_{2,t}$  = the skilled wage rate in a typical industry during period  $t$ .

The equilibrium and profit maximizing levels of employment are obtained from the first order conditions for profit maximization. Setting the first partial derivatives of the profit function with respect to each type of labor equal to zero, we obtain the following:

$$\frac{\partial \Pi_{1,t}}{\partial n_{1,i,t}} = x_{1,t} \left\{ \frac{\partial P}{\partial x_{1,t}} \frac{\partial X}{\partial n_{1,i,t}} + \frac{\partial P}{\partial t} \frac{\partial t}{\partial n_{1,i,t}} \right\} \quad (2.4)$$

$$+ P \frac{\partial X}{\partial n_{1,i,t}} - w_{1,t} = 0$$

and

$$\frac{\partial H_{1,t}}{\partial n_{2,i,t}} = x_{i,t} \left\{ \frac{P}{\partial x_{1,t}} \frac{\partial X}{\partial n_{2,i,t}} + \frac{\partial P}{\partial t} \frac{\delta t}{\delta n_{2,i,t}} \right\} \quad (2.5)$$

$$+ P \frac{\partial X}{\partial n_{2,i,t}} - w_{2,t} = 0$$

From these two profit maximizing equations, the production function and the demand equation, we can derive reduce form expressions for the equilibrium quantities of unskilled and skilled labor. Both the quantity of unskilled and skilled labor depend only upon the predetermined variables  $w_{1,t}$ ,  $w_{2,t}$  and  $t$ . These reduced form equations can be written as:

$$n_{1,i,t} = N_1 (w_{1,t}, w_{2,t}, t) \quad (2.6)$$

and

$$n_{2,i,t} = N_2 (w_{1,t}, w_{2,t}, t) \quad (2.7)$$

The equilibrium levels of unskilled and skilled labor can now be derived for a particular industry by summing over the profit maximizing levels of  $n_{1,i,t}$  and  $n_{2,i,t}$  for the various firms in that industry. If there are  $n$  firms in the industry the equilibrium levels of each type of employment can be expressed as follows:

$$n_{1,t}^* = \sum_{i=1}^n n_{1,i,t} = N_1^* (w_{1,t}, w_{2,t}, t) \quad (2.8)$$

$$(i=1,2,3, \dots, n)$$

$$n_{2,t}^* = \sum_{i=1}^n n_{2,i,t} = N_2^* (w_{1,t}, w_{2,t}, t) \quad (2.9)$$

where  $n_{1,t}^*$  and  $n_{2,t}^*$  are equilibrium quantities of unskilled and skilled labor desired by firms in a typical industry during period  $t$ . The other arguments have the same meaning as before.

The exact mathematical form of equations (2.8) and (2.9) are unknown. As is well known, however, a wide class of functions can be closely approximated by a linear function.<sup>82</sup> Therefore, equations (2.8) and (2.9) will be written in the following form.

$$n_{1,t}^* = B_{10} + B_{11}w_{1,t} + B_{12}w_{2,t} + B_{13}t \quad (2.10)$$

and

$$n_{2,t}^* = B_{20} + B_{21}w_{1,t} + B_{22}w_{2,t} + B_{23}t \quad (2.11)$$

From equations (2.1) through (2.9) we would expect that:

$$B_{11} = \frac{\partial n_{1,t}^*}{\partial w_{1,t}} < 0$$

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<sup>82</sup>Arthur S. Goldberger, Econometric Theory, (John Wiley and Sons, Inc., New York), p. 127.



$$B_{12} = \frac{\partial n_{1,t}^*}{\partial w_{2,t}} \geq 0$$

$$B_{13} = \frac{\partial n_{1,t}^*}{\partial t} > 0$$

$$B_{21} = \frac{\partial n_{2,t}^*}{\partial w_{1,t}} \geq 0$$

$$B_{22} = \frac{\partial n_{2,t}^*}{\partial w_{2,t}} < 0$$

$$B_{23} = \frac{\partial n_{2,t}^*}{\partial t} > 0$$

Since:

$$n_t^* = n_{1,t}^* + n_{2,t}^* \quad (2.12)$$

where  $n_t^*$  is the total equilibrium quantity of labor in a typical industry during period  $t$ , and the variables  $n_{1,t}^*$  and  $n_{2,t}^*$  are to be interpreted as they were previously defined in equations (2.9) and (2.10).

Then:

$$n_t^* = B_{10} + B_{11}w_{1,t} + B_{12}w_{2,t} + B_{13}t + B_{20} + B_{21}w_{1,t} + B_{22}w_{2,t} + B_{23}t \quad (2.13)$$

Since skilled wage rates, as they have been defined above, are generally unavailable for manufacturing industries, an alternative formulation is used. This can be accomplished by interpreting the average wage rate in the industry as a weighted average of the unskilled and skilled wage rates during period  $t$ . Mathematically, this relationship is as follows:

$$\bar{w}_t = \frac{n_{1,t}}{n_t} w_{1,t} + \frac{n_{2,t}}{n_t} w_{2,t} \quad (2.14)$$

or

$$w_{2,t} = \frac{n_t}{n_{2,t}} \bar{w}_t - \frac{n_{1,t}}{n_{2,t}} w_{1,t} \quad (2.15)$$

where  $\bar{w}_t$  is the average wage rate in a typical industry during period  $t$ ;  $n_{1,t}$  is the observed number of unskilled workers in a typical industry during period  $t$ ;  $n_{2,t}$  is the observed number of skilled workers in a typical industry during period  $t$ ;  $n_t$  is the sum of  $n_{1,t}$  and  $n_{2,t}$ ; and the other arguments are to be interpreted as before.

Substituting equation (2.15) for  $w_{2,t}$  in equation (2.13) and simplifying yields:

$$\begin{aligned} n_t^* = & (B_{10} + B_{20}) + \{(B_{11} + B_{21}) - (B_{12} + B_{22}) \frac{n_{1,t}}{n_{2,t}}\} w_{1,t} \quad (2.16) \\ & + \{(B_{12} + B_{22}) \frac{n_t}{n_{2,t}}\} \bar{w}_t + (B_{13} + B_{23}) t \end{aligned}$$

or

$$n_t^* = \beta_0 + \beta_1 w_{1,t} + \beta_2 \bar{w}_t + \beta_3 t \quad (2.17)$$

where it follows that:

$$\beta_0 = (B_{10} + B_{20})$$

$$\beta_1 = \{ (B_{11} + B_{21}) - (B_{12} + B_{22}) \frac{n_{1,t}}{n_{2,t}} \} = \frac{\partial n_t^*}{\partial w_{1,t}} \geq 0$$

$$\beta_2 = \{ (B_{12} + B_{22}) \frac{n_t}{n_{2,t}} \} = \frac{\partial n_t^*}{\partial \bar{w}_t} \geq 0$$

$$\beta_3 = (B_{13} + B_{23}) = \frac{\partial n_t^*}{\partial t} > 0$$

Partial Adjustments to Equilibrium  
Levels of Employment in  
Industry Labor Markets

In the previous section we have developed a model to establish equilibrium on desired levels of employment for industry labor markets. Even though these equilibrium levels of employment are desired by the industry their adjustment to them may not be complete during any given period  $t$ . The following stock adjustment model is, therefore, employed to allow for these possible differences between the desired and observed levels of industry employment.

$$(n_t - n_{t-1}) = \lambda (n_t^* - n_{t-1}) \quad 0 \leq \lambda \leq 1 \quad (2.18)$$

or

$$n_t = \lambda n_t^* + (1-\lambda) n_{t-1} + \epsilon_t \quad (2.19)$$

Where  $n_{t-1}$  is the observed level of employment in a typical industry during period  $t-1$ ;  $\lambda$  is the coefficient of adjustment;  $\epsilon_t$  is an error term which meets the standard assumptions of the Classical Linear Regression Model;<sup>83</sup> and the other arguments have their previous interpretations.

The coefficient of adjustment (i.e.,  $\lambda$ ) in equation (2.18) reflects the degree to which the desired and observed change in industry employment have approached each other between the periods  $t$  and  $t-1$ . Theoretically  $\lambda$  can take on any value between zero and one. If the value of  $\lambda$  equals one, then  $n_t = n_t^*$  and the industry labor market has made a complete adjustment to its desired or profit maximizing level of employment in the current period. If, however, the industry has failed to make any adjustment toward the equilibrium level of employment, then  $n_t = n_{t-1}$  and the value of  $\lambda$  equals zero.

Since  $n_t^*$  in equation (2.19) is unobservable we will replace it with equation (2.17) and obtain the following:

$$n_t = \lambda \beta_0 + \lambda \beta_1 w_{1,t} + \lambda \beta_2 \bar{w}_t + \lambda \beta_3 t + (1-\lambda) n_{t-1} + \epsilon_t \quad (2.20)$$

where the arguments have the same meaning as before.

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<sup>83</sup>Goldberger, op. cit., p. 162.

### Summary and the Final Form of the Model

In this chapter, we have developed a model to explain the desired and equilibrium levels of employment in industry labor markets. From this model a reduced form expression was obtained. Due to incomplete adjustments, however, the equilibrium levels of employment are in general unobservable. We, therefore, derived an expression for the observed level of industry employment as a function of the predetermined variables  $w_{1,t}$ ,  $\bar{w}_t$ ,  $t$  and  $n_{t-1}$ . The final form of our model appears as:

$$n_t = \delta_0 + \delta_1 w_{1,t} + \delta_2 \bar{w}_t + \delta_3 t + \delta_4 n_{t-1} + \epsilon_t \quad (2.21)$$

where:

$$\delta_0 = \lambda \beta_0$$

$$\delta_1 = \lambda B_1$$

$$\delta_2 = \lambda \beta_2$$

$$\delta_3 = \lambda \beta_3$$

$$\delta_4 = 1 - \lambda$$

The coefficients in equation (2.22) can theoretically take on negative or non-negative values. This results because the coefficient of adjustment (i.e.,  $\lambda$ ) is assumed to be a non-negative number between zero and one. If  $\lambda$

happens to be positive then our previous discussion regarding the signs of the coefficients in equation (2.17) will still apply.

## CHAPTER III

### THE EMPIRICAL STUDY

#### Introduction

The reduced form expression of the model developed in Chapter II was used to estimate the impact of increases in the federal minimum wage on observed levels of employment in selected high and low wage three-digit manufacturing industries for the years 1960 through 1963. In addition, the selection of industries, the data utilized, and the estimation procedure are discussed in this chapter. The empirical results are then presented. In Chapter IV some concluding remarks are given.

#### The Empirical Model

The reduced form of the model developed in Chapter II appears as follows:

$$n_t = \delta_0 + \delta_1 w_{1,t} + \delta_2 \bar{w}_t + \delta_3 t + \delta_4 n_{t-1} + \epsilon_t \quad (3.1)$$

where  $n_t$  is the observed level of employment in a typical industry during period  $t$ ;  $w_{1,t}$  is the level of the federal minimum wage in period  $t$ ;  $\bar{w}_t$  is the average wage rate in a typical industry during period  $t$ ;  $t$  is a trend variable,  $n_{t-1}$  is the level of employment in a typical industry during

the period  $t-1$  and  $\epsilon_t$  is an error term. As stated previously, it follows that:

$$\delta_0 = \lambda \beta_0$$

$$\delta_1 = \lambda \beta_1 = \frac{\partial n_t}{\partial w_{1,t}} < 0$$

$$\delta_2 = \lambda \beta_2 = \frac{\partial n_t}{\partial \bar{w}_t} < 0$$

$$\delta_3 = \lambda \beta_3 = \frac{\partial n_t}{\partial t} \geq 0$$

$$\delta_4 = 1 - \lambda \quad (0 \leq \lambda \leq 1)$$

The ZEF estimation procedure will provide asymptotically efficient and unbiased estimates of the coefficients in equation (3.1) if certain assumptions regarding the statistical properties of the model are fulfilled. These assumptions include the following: 1) the function is linear in the explanatory variables and  $\epsilon_t$ ; 2)  $\epsilon_t$  is normally and independently distributed with zero mean and constant variance; and 3) the disturbance covariances across industries are not necessarily equal to zero.

#### The Selection of Industries and the Data

The reduced form of the model presented in equation (3.1) was used to estimate the impact of federal minimum wage increases on five low and five high wage three digit manufacturing industries for the years 1960 through 1963. The low wage industries included saw mills and planing



mills, men's and boy's furnishing, knitting mills, women's and misses' outerwear, and household furniture. Since some previous investigatory work has been done in these areas, it was felt that comparisons of the results may be possible. In selecting these industries, however, consideration was also given to the number of observations in each industry group. The high wage industries chosen were iron and steel foundries, construction and like equipment, metal working machinery, communication equipment, and steel rolling and finishing. The criteria for selecting this group were the average wage rate in the industry, as well as, the number of observations.

The employment and wage data for the above industries were obtained from the 1963 Census of Manufacturers, and the 1962, 1961 and 1960 Annual Surveys of Manufacturers. Since the cross sectional and time series data were pooled, the collection of observations on employment and wage rates required a four period matching of this data for a given three digit industry and state. When this was not possible, the industry for that state was dropped as a sample observation.

Only the employment of production workers and their respective wage rates were considered in estimating equation (3.1). Total employment was rejected in favor of production employment because the former category included supervisory personnel and corporate officers. Since this latter group is usually salaried, it was believed that

production employment would be more representative of any effect which changes in the minimum wage may have upon industry employment.

As noted in Chapter I of this study, the federal minimum wage was increased twice in the 1960-63 period. It was raised from \$1.00 an hour to \$1.15 an hour in 1961, and to \$1.25 an hour in 1963.

Because average wage rates for production workers are not published for each state, this variable had to be constructed. This was accomplished by dividing the total wages paid to production workers in each year by the corresponding total man-hours worked by this group. The resulting variable should not be interpreted as average straight time hourly earnings since total wages paid to production workers includes additional payments such as overtime premiums. This variable should also be viewed with some degree of caution in the apparel industry since the man-hours of production workers were not well reported because work in the industry is generally conducted on a piece rate basis.

#### The Estimation Procedure

To estimate the coefficients in equation (3.1) the above cross sectional data was pooled for the 1960-63 period. Consistent estimates of these coefficients were obtained by ordinary least squares under the stated assumptions.

Zellner's two stage Aitken's estimates (abbreviated as ZEF) were also obtained.<sup>84</sup> This procedure produces unbiased estimates that are efficient relative to those obtained by ordinary least squares when the disturbances across equations are highly correlated.<sup>85</sup> If the disturbances are uncorrelated the ZEF estimates are only slightly worse in small samples than those obtained by ordinary least squares.<sup>86</sup>

The ZEF procedure was applied to groupings of industries that were expected on a priori grounds to have highly correlated disturbances. Generally, an attempt was made to combine industries in the same SIC two digit classification.

### The Results

The ZEF results from estimating equation (3.1) appear in Table 3.1.<sup>87</sup> For convenience the average wage rate and

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<sup>84</sup>For a more detailed description of the ZEF procedure and its properties, see: Jan Kmenta and Roy F. Gilbert, "Small Sample Properties of Alternative Estimators of Seemingly Unrelated Alternative Estimators," American Statistical Association Journal, (December, 1968), pp. 1180-1200.

<sup>85</sup>Ibid., p. 1192.

<sup>86</sup>Ibid., p. 1195.

<sup>87</sup>The ordinary least squares results are given in Appendix A. Also shown are the results from using log conversions and applying both estimation procedures. Generally the ZEF results are superior to those obtained by ordinary least squares. When this was not the case, the results are only slightly worse. In addition, the use of log conversions did not usually increase the efficiency of the estimates of the coefficients. A notable exception is industry 232 where the use of ZEF with logs made all the estimates of the coefficients significant at the ten per cent level.

the coefficient of adjustment for each three digit industry are also shown. The industries are listed in the table from the low to the high wage group.

The minimum wage variable in the reduced form of the empirical model is both significant at better than the five per cent level and possesses a negative sign for five of the ten industries investigated. Four of these were previously classified as high wage. The fifth industry in this category has the lowest average wage rate of the ten industries considered. The remaining five industries, however, do not have estimated values for this coefficient that vary significantly from zero at the ten per cent level. These results all correspond with those theoretically expected.

Theoretically consistent results are also obtained for the average wage variable. In all but two industries the coefficient for this variable did not vary significantly from zero at the ten per cent level. In the two remaining industries, it was significant and positive for the high wage industry, and significant and negative for the low wage industry with the same level of confidence.

The trend variable is significant and positive for seven industries at  $\alpha = .10$ . In the remaining industries, it does not vary significantly from zero with a ten per cent confidence level. These results are also within those previously anticipated.

TABLE 3.1--ZEF Results From Estimating Equation (3.1).

SIC	Description	N	$\frac{1}{w}$	$\lambda^2$	$\delta_0$	$\delta_1$	$\delta_2$	$\delta_3$	$\delta_4$	$R^2$
242	Sawmills and Planing Mills	48	\$1.47	.088** (.035)	8945.87** (3741.51)	-10120.19** (3914.88)	450.31** (335.64)	972.92** (318.10)	.912** (.035)	.970
232	Men's and Boys' Furnishings	32	1.52	.017 (.045)	14562.72 (10913.63)	-13126.52 (11186.64)	-2082.22 (1358.82)	1595.98* (893.97)	.983** (.045)	.963
225	Knitting Mills	32	1.64	.008 (.020)	-9559.18 (11537.79)	7628.28 (11748.56)	1327.06 (2290.59)	-687.83 (937.02)	.992** (.020)	.990
233	Womens' and Misses' Outerwear	32	1.67	.009 (.014)	-1358.22 (15530.86)	1328.77 (15269.63)	-984.24 (3033.37)	847.93 (1246.92)	.991** (.014)	.995
251	Household Furniture	48	1.85	-.002 (.027)	4609.81 (3740.91)	-4538.50 (3859.96)	-788.00* (433.91)	791.17** (309.04)	1.002** (.027)	.971
332	Iron and Steel Foundries	28	2.75	.064 (.045)	-365.55 (6067.79)	-12092.52** (4285.02)	4446.67** (1991.67)	1033.70** (383.74)	.936** (.045)	.992
353	Construction and Like Equipment	32	2.75	.017 (.025)	18453.23** (8224.90)	-21563.12** (7882.25)	348.84 (1203.33)	2068.83** (644.46)	.983** .025	.985
354	Metal Working Machinery	32	2.86	-.005 (.055)	28429.87** (12289.79)	-29995.59** (9911.80)	174.63 (2749.20)	2144.38** (845.23)	1.005** (.055)	.979
366	Communication Equipment	28	2.89	-.059 (.105)	11737.36 (26401.83)	-15860.68 (26198.33)	949.64 (3542.31)	1987.20 (2111.28)	1.059** (.105)	.875
331	Steel Rolling and Finishing	48	3.44	.016 (.013)	56030.67** (17096.93)	-58952.70** (15746.23)	343.46 (2735.29)	4113.14** (1282.44)	.984** (.013)	.993

NOTE: Estimated standard errors are in parentheses. One \* indicates significance at the ten per cent level and a double \*\* indicates a five per cent significance level.

<sup>1</sup>This figure represents the average wage rate in the industry for the 1960-63 period.

<sup>2</sup> $\lambda = 1-\delta_4$

The coefficient of adjustment is always significantly different from one at confidence levels of better than one-half of one per cent. It does not, however, differ significantly from zero at the ten per cent level for any of the industries studied except saw mills and planning mills. In each case  $\lambda$  has the appropriate magnitude and sign.

### Some Implications of the Results

Our results indicate that increases in the federal minimum wage reduced observed levels of employment in five of the ten industries investigated for the 1960-63 period. This is surprising only because four of these industries were among these classified as high wage. In the remaining industries increases in the minimum wage did not appear to have any significant effect on observed levels of industry employment at the ten per cent level. This does not imply, as some studies seem to indicate, that increases in the minimum wage do not adversely affect equilibrium levels of unskilled employment in industry labor markets. It simply means that the total effect upon the sum of observed skilled and unskilled employment did not vary significantly from zero. This may occur because the coefficient  $\lambda$  and/or the coefficient  $\beta_1$  in equation (2.17) are equal to zero.

Even though the coefficient of the average wage rate variable did not vary significantly from zero in eight of the ten industries studied, we cannot conclude

that skilled wage rates and equilibrium levels of skilled employment are not inversely related. In addition, these results do not imply that wages are not an important consideration to profit maximizing firms. The only implication is that the effect on observed levels of total industry employment is not significant. The insignificance of the average wage rate coefficient may occur because  $\lambda$  is zero and/or because the coefficient  $\beta_2$  in equation (2.17) is zero.

Since the trend variable was both positive and significant for seven of the industries studied, it appears that increases in product demand have caused most industries to expand employment for the period we are considering. This result seems to imply that these industries were already attempting to expand employment when the minimum wage was increased. It may also explain why some studies that assume equilibrium in industry labor markets conclude that increases in the federal minimum wage caused employment increases. In the three industries where the trend variable was not significant, it may be that the coefficient of adjustment was equal to zero. This possibility is further evidenced by the fact that none of the variables for these industries are significant and that  $\lambda$  did not vary significantly from zero.

The estimates of the coefficient of adjustment for the ten industries investigated during the 1960-63 period seem to indicate that industry labor markets adjust very

slowly to their desired levels of employment. Only industry 242, which had the lowest average wage rate, yielded an  $\lambda$  estimated value of  $\lambda$  which was significantly different from zero and positive.

### Summary

In this chapter the reduced form expression of the model developed in Chapter II was estimated by ordinary least squares and ZEF procedures for ten three digit manufacturing industries during the 1960-63 period. Theoretically consistent results were obtained in each instance. In general these results indicated that: 1) minimum wage increases significantly reduced observed levels of employment in half of the industries studied; 2) changes in average wage rates seldom had a significant effect on observed levels of employment; 3) most industries were expanding employment because product demand was increasing; and 4) industry labor markets adjust very slowly to desired levels of employment regardless of the cause of the disequilibrium.



## CHAPTER IV

### CONCLUDING REMARKS

A substantial part of the investigatory work to determine the impact on employment from increasing the federal minimum wage has been done by the Department of Labor. In general, the results of these studies have been interpreted as indicating that increases in the statutory minimum have had little or no effect on employment. In some cases, the assertion has even been made that employment increased more rapidly in those periods after the minimum wage increase than it had previously.

Some authors feel that the conclusions reached in the Department of Labor studies are generally valid. They believe that increases in the federal minimum wage do not reduce employment either because wage costs are not an important consideration to firms or that the existing unskilled labor employed by the firm suddenly becomes more efficient. These same authors further contend that these results nullify the assumptions of the competitive model and point out the ineptness of orthodox theory in this instance.

Critics of the Department of Labor studies maintain that the unorthodox conclusions result from a failure to

consider changes in exogenous variables and to employ adequate statistical procedures. They contend that since the minimum wage was usually raised during expansionary periods in the economy a comparison of gross changes in employment does not correctly reflect its impact.

Independent investigations, however, have often been guilty of the same criticisms raised against the Department of Labor studies. In many cases these studies have also used static-tabular analysis to determine the effect on employment from increasing the minimum wage in low wage industries. In one instance, only the gross movement of workers into low wage non-covered areas of employment was considered an adequate indication of unemployment resulting from minimum wage increases. These studies were also generally devoid of adequate statistical procedures to effectively determine the impact from increasing the statutory minimum wage.

To overcome these shortcomings, we have attempted in this study to specify a model that could be tested empirically. To accomplish this task a model was developed within the framework of orthodox theory to explain desired and equilibrium levels of industry employment, and from it a reduced form expression was obtained. Because incomplete adjustment to equilibrium levels of employment in industry labor markets may occur, a stock adjustment model was introduced. A reduced form expression for the observed level of industry employment as a

function of the minimum wage rate, the average wage rate in the industry, a trend variable, and the lagged observed level of industry employment was then obtained.

The final form of the model developed was used to estimate the impact from increasing the federal minimum wage on five high and low wage SIC three digit manufacturing industries for the 1960-63 period. Both ordinary least squares and ZEF estimation procedures were used. Each procedure was also applied after the variables had been converted to logs. The ZEF results without log conversions were the most satisfactory. In each case the results were theoretically consistent.

The results from applying the ZEF procedures indicated that: 1) increases in the statutory minimum reduced observed levels of employment in half the industries studied and most often among the high wage group; 2) average wage rates often had no significant effect on observed levels of employment; 3) increases in product demand was causing most industries to expand employment; and 4) the adjustment of industry labor markets to equilibrium levels of employment is slow regardless of the cause for the disequilibrium.

The results do not imply that minimum wage increases have neutral or positive effects on equilibrium levels of employment, or that wage costs are unimportant to profit maximizing firms. In half of the industries, even observed

levels of industry employment appeared to be inversely related to increases in the minimum wage. Failure to note this relationship in the past may be due to the concentration of most studies on selected low wage industries.

It also appears that increases in product demand may have been causing most of the industries studied to expand employment during the 1960-63 period. If this is the case, then the direct relationship between minimum wage increases and observed levels of employment are understandable. This would seem to be a more realistic explanation for the above phenomenon than either the conclusion that firms disregard profits or that unskilled labor magically becomes more efficient.

There seems to be sufficient evidence to accept our hypothesis that increases in the federal minimum wage may have reduced observed levels of industry employment. The sign of the coefficient of the minimum wage variable was negative for eight of the industries studied, but significant at the ten per cent level, for only five of them. Regardless, this result only gives us the relationship between observed employment and minimum wages, not equilibrium levels of employment. It should be remembered, however, that reductions in observed levels of employment may veil even greater decreases in potential levels of employment.

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## APPENDIX A

TABLE A.1.--Ordinary Least Squares Results From Estimating Equation (A.1).

SIC	Description	N	$\bar{w}^1$	$\lambda^2$	$\delta_0$	$\delta_1$	$\delta_2$	$\delta_3$	$\delta_4$	R <sup>2</sup>
242	Sawmills and Planing Mills	48	\$1.47	.096** (.036)	8871.82* (3743.12)	-10081.78** (3915.11)	566.96 (354.73)	954.69* (318.62)	.904 (.036)	.970
232	Men's and Boy's Furnishings	32	1.52	.016 (.041)	14289.20 (10917.78)	-13123.98 (11186.64)	-1909.87 (1370.64)	1591.16 (893.99)	.984 (.041)	.963
225	Knitting Mills	32	1.64	.005 (.019)	-10080.50 (11542.31)	7514.20 (11749.93)	1696.53 (2306.44)	-694.14 (937.03)	.995 (.019)	.990
233	Women's and Misses' Outerwear	32	1.67	.013 (.016)	-2093.16 (15554.20)	1513.89 (15270.88)	-526.21 (3069.36)	805.17 (1247.68)	.987 (.016)	.995
251	Household Furniture	48	1.85	-.006 (.028)	4734.21 (3747.83)	-4534.06 (3859.96)	-886.84* (459.24)	796.55 (309.13)	1.006 (.028)	.971
332	Iron and Steel Foundries	28	2.75	.059 (.041)	-301.01 (6219.38)	-12098.16* (4285.23)	4401.73* (2079.54)	1039.60* (387.16)	.941 (.041)	.992
353	Construction and Like Equipment	32	2.75	.017 (.025)	18446.43* (8227.40)	-21558.76* (7882.31)	352.00 (1205.49)	2068.02 (644.52)	.983 (.025)	.985
354	Metal Working Machinery	32	2.86	-.002 (.056)	27984.46* (12299.18)	-29891.67* (9912.28)	329.29 (2754.13)	2125.43* (845.46)	1.002 (.056)	.979
366	Communication Equipment	28	2.89	-.041 (.110)	10469.70 (26478.42)	-15784.09 (26198.61)	1507.15 (3627.47)	1956.39 (2112.80)	1.041 (.110)	.876
331	Steel Rolling and Finishing	48	3.44	.014 (.013)	55216.74* (17125.76)	-59056.91* (15746.43)	615.05 (2755.21)	4093.57 (1282.81)	.986 (.013)	.993

NOTE: Estimated standard errors are in parentheses. \* indicates significance at the ten per cent level and \*\* indicates a five per cent significance level.

<sup>1</sup>This figure represents the average wage rate in the industry for the 1960-63 period.

$2\lambda = 1-\delta_4$

TABLE A.2.--Ordinary Least Squares Results From Estimating Equation (3.1) (In Ln).

SIC	Description	N	$\bar{w}^1$	$\lambda_2$	$\delta_0$	$\delta_1$	$\delta_2$	$\delta_3$	$\delta_4$	$R^2$
242	Sawmills and Planing Mills	48	\$1.47	0.0203 (0.0286)	0.0501 (0.1143)	-1.0305** (0.4224)	0.0368 (0.0429)	0.1838** (0.0657)	0.9797** (0.0286)	.977
232	Men's and Boy's Furnishings	32	1.52	-0.0031 (0.0367)	0.0153 (0.1639)	-1.5221* (0.8838)	-0.1979 (0.1334)	0.2582* (0.1362)	1.0031** (0.0367)	.973
225	Knitting Mills	32	1.64	0.0056 (0.0236)	0.0300 (0.1136)	0.0228 (0.9662)	0.1936 (0.2243)	-0.0177 (0.1479)	0.9944** (0.0236)	.985
233	Women's and Misses' Outerwear	32	1.67	0.0038 (0.0231)	0.0017 (0.0845)	0.0410 (0.8604)	0.0064 (0.2087)	0.0524 (0.1348)	0.9962** (0.0231)	.990
251	Household Furniture	48	1.85	-0.0024 (0.0221)	-0.0010 (0.0854)	-0.6053 (0.4776)	-0.1355* (0.0722)	0.1727** (0.0737)	1.0024** (0.0221)	.981
332	Iron and Steel Foundries	28	2.75	0.0591 (0.0422)	-0.2699** (0.1264)	-1.4638** (0.6632)	1.1469* (0.5624)	0.2212* (0.1073)	0.9409** (0.0422)	.989
353	Construction and Like Equipment	32	2.75	0.0147 (0.0208)	0.0829 (0.1055)	-1.8805** (0.7372)	-0.0990 (0.2184)	0.3417** (0.1153)	0.9853** (0.0208)	.989
354	Metal Working Machinery	32	2.86	0.0011 (0.0231)	0.0072 (0.0800)	-2.2819** (0.5649)	0.0459 (0.2321)	0.3084** (0.0882)	0.9989** (0.0231)	.992
366	Communication Equipment	28	2.89	0.0253 (0.0650)	0.0887 (0.2325)	-0.7658 (1.4287)	0.1118 (0.3867)	0.1706 (0.2217)	0.9747** (0.0650)	.938
331	Steel Rolling and Finishing	48	3.44	-0.0046 (0.0101)	-0.0704 (0.1043)	-1.0770** (0.4781)	0.1125 (0.2244)	0.1361* (0.0758)	1.0046** (0.0101)	.996

NOTE: Estimated standard errors are in parentheses. \* indicates significance at the ten per cent level and \*\* indicates a five per cent significance level.

<sup>1</sup>This figure represents the average wage rate in the industry for the 1960-63 period.

<sup>2</sup> $\lambda = 1-\delta_4$

TABLE A.3.--ZEF Results From Estimating Equation (3.1) (In Ln).

SIC	Description	N	$\bar{w}^1$	$\lambda_2$	$\delta_0$	$\delta_1$	$\delta_2$	$\delta_3$	$\delta_4$	R <sup>2</sup>
242	Sawmills and Planning Mills	48	\$1.47	0.0138 (0.0277)	0.0250 (0.1109)	-1.0346** (0.4424)	0.0224 (0.0406)	0.1869** (0.0657)	0.9862** (0.0277)	.997
232	Men's and Boy's Furnishings	32	1.52	-0.0013 (0.0365)	0.0266 (0.1631)	-1.5244* (0.8838)	-0.2222* (0.1329)	0.2597* (0.1362)	1.0013** (0.0365)	.973
225	Knitting Mills	32	1.64	0.0092 (0.0225)	-0.0171 (0.1108)	0.0174 (0.9658)	0.2034 (0.2155)	-0.0176 (0.1473)	0.9908** (0.0225)	.985
233	Women's and Misses' Outerwear	32	1.67	-0.0019 (0.0219)	-0.0302 (0.0822)	-0.0555 (0.8598)	0.0446 (0.1925)	0.0474 (0.1344)	1.0019** (0.0219)	.990
251	Household Furniture	48	1.85	-0.0028 (0.0212)	-0.0005 (0.0821)	-0.6057 (0.4776)	-0.1443** (0.0677)	0.1733** (0.0737)	1.0028** (0.0212)	.981
332	Iron and Steel Foundries	28	2.75	0.0561 (0.0421)	-0.2674** (0.1259)	-1.4622** (0.5603)	1.1121** (0.6632)	0.2235** (0.1073)	0.9439** (0.0421)	.989
353	Construction and Like Equipment	32	2.75	0.0148 (0.0207)	0.0875 (0.1051)	-1.8822** (0.7372)	-0.1090 (0.2177)	0.3426** (0.1153)	0.9852** (0.0207)	.989
354	Metal Working Machinery	32	2.86	0.0011 (0.0230)	0.0195 (0.0797)	-2.2820** (0.5649)	0.0182 (0.2310)	0.3100** (0.0882)	0.9989** (0.0230)	.992
366	Communication Equipment	28	2.89	0.0273 (0.0649)	0.0869 (0.2316)	-0.7636 (1.4287)	0.1346 (0.3859)	0.1688 (0.2217)	0.9727** (0.0649)	.936
331	Steel Rolling and Finishing	48	3.44	-0.0070 (0.0099)	-0.0765 (0.1026)	-1.0800** (0.4781)	-1.1043** (0.2206)	0.1372** (0.0757)	1.0070** (0.0099)	.996

NOTE: Estimated standard errors are in parentheses. \* indicates significance at the ten per cent level and \*\* indicates a five per cent significance level.

<sup>1</sup>This figure represents the average wage rate in the industry for the 1960-63 period.

<sup>2</sup> $\lambda = 1-\delta_4$



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