SIZE OF PLOT AND NUMBER OF REPLICATIONS NECESSARY
FOR VARIETAL TRIALS WITH WHITE PEA BEANS.

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SIZE OF PLOT AND NUMBER OF REPLICATIONS NECESSARY FOR VARIETAL TRIALS WITH WHITE PEA BEANS.

Many experiments have been conducted with different field crops to determine what size of plot and how many replications are necessary to give reliable results. nomical administration of funds requires that the plots be small and repeated a minimum number of times. However, efficient plot technique requires that the plots be of sufficient size and repeated often enough to make the results reliable. That information obtained for one crop under certain conditions can not be applied to another is clearly shown by articles on the subject. A complete bibliography (1) of articles dealing with plot technique in general was reported by a committee of the American Society of Agronomy, for the standardization of field experiments. Additional articles (2) were recently reported. A search of this literature fails to reveal any experiments of this nature with white pea beans (Phaseolus vulgaris, L.). Odland and Garber (4) in their work with soybeans, which comes nearest to this particular subject, concluded that under conditions existing where their experiment was conducted, a 16-foot plot one row wide replicated three times was the most satisfactory when both accuracy, and economy of land and of labor were taken into consideration.

The object of the study reported in this paper was to find the proper size of plot and the number of replications of this size necessary for varietal trials with white pea beans when the trials are conducted under conditions similar to those prevailing in this experiment. Due consideration should be given to the amount of land to be used and to the convenience of handling the plots in all field operations.

MATERIAL AND METHODS

The Robust variety of white pea beans is used as a standard of comparison in variety trials at the Michigan Agricultural Experiment Station and is grown extensively throughout the state. Consequently, it was used in this experiment. An area of Conover soil of approximately one and three-quarters' acres, located in one of the regular plant-breeding sections, was chosen for the planting.

The beans were planted in rows twenty-eight inches apart on June 15, 1932, and so spaced that the plants were approximately one to two inches apart. To insure a good stand sixty pounds of beans were planted to the acre instead of the usual rate of forty-five pounds. After the last cultivation, an area 210 feet long and 210 feet wide was chosen from the center of the field for the experiment. A total of 1890 plots was obtained by dividing this area

into twenty-one ten-foot series, each ninety rows wide. A count of the number of plants in each ten-foot plot was made shortly before harvest. At maturity the plants were pulled by hand and allowed to cure. The ten-foot plots were threshed and the beans air dried in a warm room before weighing. Fortunately, the season was favorable for beans and all plots had a good stand without any skips although a fairly wide range in the number of plants per row occurred. The coefficient of correlation between stand and yield was found to be .10± .01. This coefficient was too small to make any correction in the yields for variations in stand.

The statistical constants used to study the problem were obtained by making use of the "analysis of variance" as used by Immer (3). The principles of the method as stated by Immer are:

If the total variability of the observations on all the plots is given in suitable term (sum of squares) it may legitimately be apportioned to various known causes, leaving a remaining portion ascribable to uncontrolled or unknown causes. The latter will then serve as a basis for the calculation of the error of the experiment. The variance (stand Teviation squared) due to any of the known causes or to the uncontrolled or unknown causes may then be found by dividing the sum of squares by the appropriate number of degrees of freedom. term "degrees of freedom" is here used in the sense of "independent comparisons". With N quantities whose mean is fixed, there are in general N-1 independent comparisons or degrees of freedom.

The assumption was made that five varieties or treatments were to be tested by the analysis and that the arrangement of the plots within each replication was at random. Thus, variation between blocks could be removed legitimately from the total variation, by subtracting the sum of squares due to variation between blocks from the total sum of squares of all of the plots. The remainder is due to variation within blocks. The standard error, calculated from this remainder, was used as the error of the experiment. This error will be smaller than the standard error as calculated from the total population only when the variation between blocks is greater than that within blocks.

Example of Analysis of Variance for Entire Plot

An example of the analysis of variance as applied to yields of the 1,890 plots ten feet long and one row wide is given in Table 1.

TABLE 1 - Analysis of Variance of Yield of Beans in Plots Ten Feet Long, One Row Wide, for 21 Series.

| Variation | Degrees of Freedom | of | Variance or Mean Square | Standard Devia- tion | <u>F</u> * | Error in % of Mean Yield |
|---------------------------|--------------------------|------------------------|-------------------------------|----------------------------|------------|-----------------------------------|
| Between blocks | 377 | 6,967,077.18 | 19,678.48 | 140.28 | | |
| Within blocks | 1512 | 4,879,158.20 | 3 ,2 26.96 | 56.80 | 6.10 | 12.67 |
| Total between plots | 1889 | 11,846,2 <u>35.</u> 38 | 6,271.16 | 79.20 | | |

Mean (M) = 448.38

^{*}The ratio of the larger mean square to the smaller.

The formula used to obtain the total sum of squares was $S(X^2) - S(X)M_X$ where X represents a variate and M_X the mean of the population. This formula is briefed to $S(X^2) - C$, where C (correction) represents $S(X)M_X$. The formula for the sum of squares between blocks was $\frac{S(B^2)}{5} - C$ in which B was the sum of the five plots in a block and C was the correction term $S(X)M_X$. The sum of B^2 was divided by 5 to place the value on a single plot basis.

The total sum of squares 11,846,235.38 was obtained by subtracting C or 379,966,062.62 from S(X²) or 391,812,298.00. The sum of squares between the 378 blocks of the five adjacent plots added sidewise of the series was obtained in a similar way except the summation was divided by five, to place values on the basis of a single plot, before subtracting the correction factors. Thus $\underline{S(B^2)}$ - C became 1,934,665,699.00 - 379,966,062.62 which gave 6,967,077.18. The sum of squares due to variation within blocks was the difference between the total sum of squares and that portion due to variation between blocks. Since the total of 1890 plots was considered, there were 1889 (N-1) degrees of freedom attributable to the total sum of squares. There were 378 blocks (of five plots each) and consequently 377 (N-1) degrees of freedom due to blocks. The difference of 1512 was the number of degrees of freedom due to variation between the five plots within each of the 378 blocks.

sum of squares was divided by its respective number of degrees of freedom to give the mean square. The standard error is the square root of the mean square, and, calculated from the remainder was 56.80 grams or 12.67 per cent of the mean yield of 448.38 grams.

The significance of the difference between the variance between blocks and that within blocks was determined by the F-test developed by Snedecor (5). The values of F in these tables are given for two different levels of significance, the five per cent and the one per cent points for selected numbers of degrees of freedom. The former is expected to be exceeded in random sampling from a homogeneous population five times in one hundred trials, the latter only once. The five per cent point is taken as a convenient minimum level of significance. In Table 1 the observed value of F exceeds the one per cent point and it can be concluded that removal of the variation between blocks was worthwhile.

Example of Analysis of Variance for Center Rows of Plot

The standard error of a three-row plot was determined, with the outer rows discarded to eliminate possible competition from other plots. The center values were used from each three-row plot. The analysis of variance is given in Table 2.

TABLE 2 - Analysis of Variance of Yield of Beans in Three-Row Plots, Ten Feet Long, of which only the Center Row was Used, for Twenty-one Series.

| Variation | Degrees of Freedom | Sum of Squares | Variance or Mean Square | Standard Devia- tion | F | Error in % of Mean Yield |
|---------------------------|--------------------------|----------------------|-------------------------------|----------------------------|------|-----------------------------------|
| Between blocks | 125 | 2,277,829.65 | 18,222.64 | 134.99 | | |
| Within blocks | 504 | 1,887,275.60 | 3,744.59 | 6 1 .19 | 4.87 | 13.40 |
| Total between plots | 629 | 4,165,105.25 | 6,707.64 | 81.90 | | |

$$M = 456.73$$

The number of degrees of freedom attributable to total variation was 629, since there were 630 three-row plots in the entire field of 1890 single rows, ten feet long. Each block of five then required fifteen rows. The 630 three-row plots subdivided into blocks of five contributed 125 degrees of freedom, leaving 504 degrees of freedom attributable to variation between plots within blocks.

The observed value of F exceeded the one per cent point and it was concluded that the removal of the variation between blocks was worthwhile.

The standard error of a single plot was 61.19 grams or 13.40 per cent of the mean yield (456.73 grams) of all the central rows in the three-row plots in the field.

RESULTS AND DISCUSSION

The weights in grams of cleaned beans for the 1890 plots

are given in Table 3. Analyses were made of the fields of plots 1, 2, 3, 4, 5, 6, 9 and 18 rows wide and 10, 20, 30, 40 and 50 feet in length, respectively. A total of 108 different sizes and shapes of plots were studied, 72 of these involved the entire plot while the other 36 used only the center rows, the outside or border rows being discarded to eliminate competition between plots. Certain of these combinations could not be based on the entire area of 21 series and 90 rows wide. This was partially due to the original assumption involving five varieties or treatments. The four-row and five-row plots could not use all 90 rows because 90 is not divisible by four, nor could the 18 plots five rows wide be subdivided into blocks containing five plots each. Also, 21 series could not be subdivided into plots 20, 40, or 50 feet in length, but could be divided into plots 30 feet in length. To overcome these limitations, values were calculated for areas 21 and 20 series in length containing 90, 80, and 75 rows in width. In order to compare the values obtained with plots four and five rows wide with those 1, 2, 3, 6, 9, and 18 rows wide, the standard errors were also calculated for the areas 21 and 20 series long, and 80 and 75 rows wide, respectively. The same conditions hold true for the plots having the border rows discarded. Similar calculations were made whenever necessary.

TABLE 3 - Yield of Beans in grams from 1890 Single-Row Plots, Each Ten Feet Long.

TABLE 3 - (continued)

STANDARD ERRORS IN PER CENT OF MEAN

The standard deviations, as obtained by analysis of variance for the plots of varying lengths and widths, were expressed in percentage of their respective means and are shown in Tables 4 and 5. Before these tables and all subsequent tables can be discussed, it is necessary to determine whether the values obtained for the plots four and five rows wide can be compared with those obtained for other widths. The values are considered to be comparative in this and in all future tables because similar calculations from 1, 2, and 3-row plots for the areas of 80 and 75 rows in width, respectively, are in close agreement with those calculated for the entire area of 90 rows. These values are also shown in Tables 4 and 5.

Considering the entire plots, the standard error in per cent of the mean, decreased as the length of the plot was increased. The greatest reductions come when the plots were increased from 10 to 20 feet in length. Considerable further reduction in percentage error occurs when the plots were increased to 30 feet in length, but when the length of the plot was increased to 40 or 50 feet, further reduction does not compensate for the increase in land used.

Increasing the width of the plot from one row to two rows shows a great reduction in percentage error, compara-

4 - Standard Errors of Single Plots in Percentage of the Mean, of Yields of Plots Varying in Size and Shape, when Entire Plot was Used. TABLE

| Across the Tield. 21 | Length of Plot in Feet. | Number of Series of Plots Used End to End | | | Width of | of Plots | in Rows | S | | | |
|---|-------------------------------|--|------------------|------|----------|---------------------|---------|--------|---------|-----|-------|
| 21 | | Across the Field. | 1 | 2 | 3 | 4 | 5 | 9 | 6 | | 18 |
| 20 \$\frac{\kappa_{12} \cdot \beta_{1}}{\kappa_{12} \cdot \cd | | 21 | 79 | 9.98 | | | | 15.30 | 1 | | 9.83 |
| 20 \$\frac{\kappa_{12} \cdot 27}{\kappa_{12} \cdot 27} \frac{\pi_{2} \partial_{2} \p | 10 | | M12.88 | | 2 | | 15 8.15 | | | | l II' |
| Signature Sign | | 20 | 812.11 812.27 | مإه | 30 9. (1 | | | 의 | 9) | | • |
| 10 & 9.38 \(\times 7.82 \) \(\times 9.36 \) \(\times 9.36 \) \(\times 9.36 \) \(\times 7.82 \) \(\times 7.89 \) \(\times 7.89 \) \(\times 7.89 \) \(\times 7.99 \) \(\times 7.90 \) \(\times 7.90 \) \(\times 7.28 \) \(\times 6.51 \) \(\times 6.51 \) \(\times 6.63 \) \(\times 7.20 \) \(\times 6.72 \) \(| | | श्र 2. 25 | | 9.7 | | 8 6. | | | | |
| 10 & 9.43 & 2.82 & 3.08.73 & 3.7.45 & 3.7.45 & 3.8.73 & 3.8.7.94 & 3.8.14 & 2.07 & 8 & 7.94 & 3.8.14 & 2.07 & 8 & 7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & 3.8.7.94 & 3.8.7.95 & | | | 1.1 | 7 | 30 8.03 | | | 1 | Μ | ريا | 8.99 |
| 1 | 00 | 10 | - 1 | eg | | ∞ | | | | | |
| (2) (2) (2) (3) (4) (4) (5) (5) (5) (5) (5) (5) (5) (5) (5) (5 | | | ਲ 9.36 | | ' ' | | 7 | | | | |
| 7 \$\&\frac{1}{12}\$ \\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\\ | | | 8 7.89 | ó | ³₀ 6.90 | | | 60 | 10 7. | ک | 8.48 |
| 5 \$\frac{\chi}{8} \chi \cdot 28 \\ \chi \chi \chi \chi \chi \chi \chi \chi | 30 | 2 | 9 | 6.72 | 1 | 7 | | | | | |
| 5 & 7.28 45 6.51 30 6.63 | _ | | - F | | 25 (.04 | | 15 6.22 | | | | |
| 5 & 7.23 & 6.24 20 7.47 35.72 35.66 30.6.25 35.72 4 & 6.78 & 5.81 20.720 35.553 35.5 | - | | $ \cdot $ | 1 1 | ف | | | d 7.75 | 10 6.18 | | 7.92 |
| 4 \$ 6.72 \$ 5.66 \$ 6.25 \$ 5.81 \$ 5.12 7 4 \$ 6.74 \$ 5.81 \$ 5.81 \$ 5.81 \$ 5.55 | <u></u> | Ŋ | H | - 1 | | ٦ | | | | | |
| 8 6.72 以 5.66 3。6.25 スプログライン 1.59 ル 6.12 スプログライン 1.50 スプログライン 1.56.44 スプログライン 1.55.53 | | | 긱 | | 25 6.65 | | 5 | | | | |
| 4 & 6.78 40 5.81 20 7.20 125 6.44 125 5.53 125 6.44 125 5.53 125 6.44 125 6.45 125 6.44 125 6.45 125 6.45 125 6.45 125 6.45 125 6.45 125 6.45 125 6.45 125 6.45 125 6.45 125 6.45 125 | | | ó | 1 1 | 30 | | | | 5 | را | 7.91 |
| 6.74 125 6.44 | S S | 寸 | اف | ΓJ | | $ egthinspace{2mm}$ | | | | | |
| | | | | | 25 6.44 | | 15.53 | | | | |

*Small figures in each compartment of Tables 4 and 5 indicate the number of plots side to side across the field.

5 - Standard Errors of Single Plots, in Percentage of the Mean, of Yields of Plots Varying in Size and Shape, when Border Rows were Removed. TABLE

| | 6 | | | 7 | 10.06 | | | 76.6 | | | 8.59 | | | 7.61 | | | 6.64 | | | 24.9 | | |
|------------------------|------------|------------|------------|--------|----------|--|---------|----------|---------|----------|-------|---------|-----------|---------|---------|------|---------|---------|---------|---------|---------|---------|
| | 9 | | | # | 11.12 /0 | | | 11.01 | | | 9.47 | | | 8.36 % | | | 8.11/0 | | | 7.77 10 | | |
| | 5 | | | 3 | 2/ | | 9.93 | SY | | 9.81 | 15 | | 8.00 | 15 | | 6.63 | 15 | | 6.33 | 3/ | | 6.01 |
| | 4 | | | 2 | | 20.12.09 | 51 | | ≈ 12.09 | אכ | | 20 9.84 | 57 | | 20 8.72 | 31 | | 20 7.73 | 51 | | 20 7.72 | 51 |
| | 2 | | **** | ы | 13.40 J | | 2513.42 | 3, 13.42 | | 194.6122 | 30.26 | | 25 10. 34 | 30 8.67 | | | 30.8.06 | _ | 25 8.01 | 30 7.36 | - | 25 7.44 |
| Width of Plot | in Rows. | Number | of Rows | Used. | | | | | | | | | | | | | | | | | | |
| Number of Series of | Plots Used | End to End | Across the | Field. | | 21 | | | ე ე | | | 01 | | | _ | | | ι. | | | # | |
| Length of Plot | in Feet. | | | | | 10 | | | | | | 200 | | | 30 | | | ₽ | | | 50 | |

*Small figures in each compartment of Tables 4 and 5 indicate the number of plots side to side across the field.

ble to increasing the length from 10 to 20 feet. Increasing the width of plots to three rows reduces the percentage error slightly while further increases in width change the results but very little.

When the border rows are discarded and only the center rows used in the calculations, the results are similar to those obtained when the entire plot is used. A great reduction in percentage error was noted when the length of plot was increased from 10 to 20 feet and a still further reduction when increased to 30 feet, but very little difference was noted beyond that. Increasing the width of the plot from three rows to four rows resulted in a slight decrease in the percentage error but when the width was increased to five rows a considerable reduction was observed. the width of the plot to six and nine rows showed very little change one way or the other. The standard error was in all but one case greater than when all the rows of the same sized plots were used in the calculations. Increasing the size of plot to allow for the discarding of the border rows resulted in some increase in the percentage of the standard error in all but three cases where a corresponding number of rows were harvested. This means that there was greater variability between the plots of a block when only the center rows were harvested than when the entire plot of the same number of harvested rows was used.

NUMBER OF REPLICATIONS

The theoretical number of replications required by any given size of plot to reduce to four per cent the standard error of the mean of that number of replications may be obtained by dividing the standard error percentages, as given in Table 4 and Table 5, by four and squaring these quotients. The values thus obtained are given in Tables 6 and 7. The value of four per cent was chosen because normally this is the approximate size of the standard error of nine replications of bean plots at the Michigan Station.

The theoretical number of replications for the entire plot tended to decrease as the length of the plot was increased. The greatest difference in number was observed when the plot was increased from 10 to 20 feet, the number of replications for the 20-foot plots being nearly one—half of that for the 10-foot plot in the two narrowest width plots. There was a further noticeable decrease when the length of the plot is increased from 20 to 30 feet but scarcely any when increased beyond that.

When the width of the plot was increased from one row to two rows there was also a decided decrease in the number of replications required but further increase in the number of rows per plot made very little change in the theo-

TABLE 6 - Theoretical Number of Replications Needed to Reduce the Standard Error of the Mean to Four Per Cent, when Entire Plot was Used.

TABLE 7 - Theoretical Number of Replications Needed to Reduce the Standard Error of the Mean to Four Per Cent when the Border Rows were Removed.

| | | | | | | | | | _ | 1 | 7- | | | | | | | | _ | ! |
|--------------------------------------|--------------------------|--------|------|-----|------|------|-----|------|-----|-----|------|-----|-----|-----|------|-----|-----|-----|----------|-----|
| 6 | ı |) | 6.3 | | , | 6.2 | | | 4.6 | | | 3.6 | - | | 2.8 | | | 2.6 | | |
| . 9 | - | + | 7.7 | | | 7.6 | | | 5.6 | | | 4.4 | | | T. † | | | 3.8 | | |
| ī | ı | 2 | | | 6.2 | | | 6.0 | | | 0.4 | | | 2.7 | | | 2.5 | | | 2.3 |
| ≠ | c | ۲ | | 9.1 | | | 9.1 | | | 6.0 | | | 4.8 | | | 3.7 | | | 3.7 | |
| 2 | r | 7 | 11.2 | | 11.3 | 11.2 | | 11.3 | 9.9 | | 1.79 | 4.7 | | 4.8 | 4.0 | | 4.0 | 3.4 | | 3.4 |
| Width of Plot in Rows. | Number of Rows | Usea. | | | | | | | | | | | | | | | | | | |
| Number of Series of Plots Used | End to End Across the | Field. | | 21 | | | ଧ | | | 10 | | | _ | | | נט | | | 寸 | |
| Length of Plot in Feet. | | | | 10 | | | | | | ଧ | | | 30 | | | 요 | | | 2 | |

retical number of replications necessary to reduce the standard error of the mean to four per cent.

Practically the same thing is true when the border rows were discarded and only the center rows used in the calculations. The greatest decrease was noted in the three-row plots when the length of the plot was increased from ten to twenty feet. Increasing the length of the plots to thirty feet tended to reduce the number still more but when increased to forty and fifty feet the reduction in theoretical number of replications was small.

When the two center rows of the four-row plots or the three center rows of the five-row plots were used instead of the center one of the three-row plots, the theoretical number of replications was appreciably decreased, but a further increase in number of rows to six and nine did not reduce the number greatly.

COMPARATIVE LAND EFFICIENCY

Tables 6 and 7 indicate that as the plot was increased in length and, to some extent, in width the number of replications needed to put the standard error to the common basis of four per cent of its mean was reduced. However, the larger plots required more land. To determine, then, which size and shape of plot used the land most effectively the comparative land efficiency values for the 105 sizes

and shapes of plots were obtained. It was assumed that a plot ten feet long and one row wide had an efficiency of 100.0 per cent. According to Table 6, a plot ten feet long and two rows wide required 6.2 replications to reduce the standard error of the mean to four per cent. Since the plot was two rows wide, 6.2 was multiplied by two and this product divided into 10.0 (the theoretical number of replications needed for a plot ten feet long and one row wide), then multiplied by 100 to give per cent. Thus

$$\frac{10.0 \times 100}{6.2 \times 2} = 80.6$$
 per cent.

Values for land efficiency calculated in this way are given in Tables 8 and 9.

It will be noted in Table 8 that when the length of the plot was increased from ten to twenty feet the percentage of land efficiency was materially decreased but when the length of plot was increased from twenty to thirty feet the land efficiency percentage value did not decrease.

When the length of plot was increased still further there was a considerable decrease in the percentage of land efficiency. An increase in the number of rows per plot from one to two caused a decrease in percentage of land efficiency. Still further decreases in land efficiency were noted when the number of rows per plot was increased to three, to five and so on to eighteen.

TABLE 8 - Percentage Efficiency in Use of Land of Plots Varying in Size and Shape, when Entire Plot was Used.

| | | į | | . 1 | | . I | 1 | | 20 | - | | | | | 1 1 | | | 18 1 | | 1 |
|--------------------------------------|------------|--------|-------|-------|-------|-------|-------|-------|------|---------------|------|------|------|------|------|------|------|------|--------|------|
| | | 18 | 9.2 | | | 8.6 | | | 5.0 | | | 4.1 | | | 3.2 | | | 2.6 | | |
| | | 6 | 20.8 | | | 19.5 | | | 12.9 | | | 11.9 | | | 10.6 | | | 8.7 | | |
| WS | | 9 | 25.2 | | | 23.2 | | | 14.8 | | | 13.4 | | | 10.2 | | | 8.5 | | |
| s in Ro | | 5 | | | 49.9 | | | 37.7 | | | 27.1 | | | 58.9 | | | 22.9 | | | 19.6 |
| Width of Plots in Rows | | † | | 40.4 | | | 36.0 | | | 24.7 | | | 22.2 | | | 16.8 | | | 14.5 | |
| Width | | 3 | 56.4 | | 58.4 | 51.8 | | 52.8 | 37.8 | | 39.5 | 37.4 | | 37.2 | 27.8 | | 28.5 | 25.0 | | 24.1 |
| | | 2 | 80.5 | 83.2 | | 4.47 | 75.4 | | 59.7 | 6 1 .5 | | 61.4 | 61.9 | | 43.3 | 48.3 | | 45.8 | 44.6 | |
| | | ٦ | 100.0 | 100.0 | 100.0 | 100.0 | 100.0 | 100.0 | 83.4 | 84.6 | 85.7 | 85.9 | 88.1 | 88.7 | 69.1 | 71.9 | | 8.49 | 65.5 | 66.2 |
| Number of Series of Plots Used | End to End | Field. | | 21 | | | 20 | | | 10 | | | _ | | | 5 | | | ℷ | |
| Length of Plot in Feet. | | | | | 10 |) | | | | 20 | | | 30 | | | 울 | | | 20 | |

TABLE 9 - Percentage Efficiency in Use of Land of Plots Varying in Size and Shape when the Border Rows were Removed.

| C | 7 | _ | 17.6 | | | 16.4 | | | 11.0 | | | 10.3 | | | 9.2 | | | 7.8 | | |
|------------------------|------------|-------------------|------|------|------|------|----------|------|------|------|------|------|------|------|------|------|------|------|------|------|
| 7 |) | † | 21.6 | | | 20.2 | | | 13.6 | | | 12.7 | | | 9.3 | | | 8.1 | | |
| L | | 2 | | | 33.6 | | | 31.2 | | | 23.5 | | | 25.1 | | | 18.7 | | | 16.6 |
| - | ŀ | 2 | | 28.6 | | | 25.7 | | | 19.4 | | | 18.3 | | | 15.2 | | | 12.6 | |
| r | 2 | F | 29.8 | | 30.7 | 27.3 | | 27.6 | 23.2 | | 23.4 | 23.7 | | 23.9 | 18.8 | | 19.5 | 18.0 | | 18.1 |
| Width of Plot | Number | of Rows Used. | | | | | | | | | | | | | | | | | | |
| Number of Series of | End to End | Across the Field. | | 21 | | | <u>୧</u> | | | 10 | | | _ | | | 77 | | | # | |
| Length of Plot | in Feet. | | | 10 | | | | | | ଧ | | | 30 | | | 오 | | | 50 | 1 |

In Table 9 are given the comparative land efficiency values for the plots when the border rows were discarded and only the center rows were used. As in Table 8, the shorter plots were the more efficient ones. Unlike the results from the entire plots, it is seen that the land efficiency values for the three center rows of the five-row plots 10-20-and 30-feet long were slightly higher than the efficiency values for the center rows of plots three and four rows wide of similar length.

A comparison of Tables 8 and 9 shows that harvesting the entire plot is more efficient than using only part of the plot and that the differences in efficiency are much greater in the narrower widths of plot.

SUMMARY

Summarizing the results which have been obtained it can be said that for this field of beans:

The standard errors in per cent of their respective means were greater in nearly all cases when only the center rows of a plot were used than when the entire plot was harvested. This was true whether the comparison was between equal numbers of rows harvested or equal numbers of rows per plot.

The standard error in percentage decreased as the area

of the plot increased, except as noted in the preceding paragraph. The decreases were consistent with increases in length, but were not entirely consistent with increases in width.

Rather large decreases in magnitude of per cent standard error occurred when the plot length was increased from ten to twenty and on to thirty feet. Comparatively small decreases were found when the plot length was increased beyond thirty feet.

Rather large decreases in magnitude of per cent standard errors were observed as the plot width was increased to two rows but variable results were obtained as the plot width was increased beyond two rows.

On the whole, increasing the size of the plot by increasing its length proved much more effective in reducing the standard error in per cent of its mean than making a similar increase in plot size by increases in width.

The number of replications needed to reduce the standard errors to the same comparable basis of four per cent followed the same trends as did the standard errors. This is not to be wondered at as the number of replications was dependent on the magnitude of the standard error.

An increase in the theoretical number of replications decreases the standard error more rapidly than an equivalent increase in the size of plot. This is especially true when

the plot is increased in width.

The comparative land efficiency values bring out, more strongly than the standard errors alone, the greater desirability of harvesting entire plots in comparison to harvesting only the center rows of the plot.

When the entire plot was harvested, plots one row wide and thirty feet long were slightly more efficient than plots the same width and twenty feet long and much more efficient than plots two rows wide and only ten feet long.

Although the comparative land efficiency values indicate that the ten-foot, one-row plots were the most efficient in the use of land, yet there are several other factors not considered under standard error and land efficiency values which must be considered. These are ease of planting, of harvesting, of threshing and subsequent laboratory determinations.

It is much easier, less time consuming, and subject to less mistakes to plant four replications (see Table 6) of plots thirty feet long, one row wide than to plant ten replications of plots ten feet long, one row wide. Harvesting, threshing and other procedure would all be in favor of the larger of these two sizes of plots with their respective numbers of replications.

CONCLUSIONS

The data obtained from 1890 ten-foot plots of white pea

beans studied by the variance method indicates that plots thirty feet long and one row wide replicated four times were more efficient in the use of land for the reducing of the standard error in per cent of its mean than all of the other lo7 shapes and sizes studied, except the original ultimate units ten feet long and one row wide.

Field operations, such as planting, harvesting and threshing, and subsequent laboratory determinations indicate that the use of the plots thirty feet long, one row wide, which require but four replications is more desirable than that of the more land-efficient but smaller plots, ten feet long, one row wide, which require ten replications.

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